



1 **Hydrological Drought forecasting under changing environment**
2 **in Luanhe River basin**

3 Min Li^{1,2}, Mingfeng Zhang², Runxiang Cao³, Yidi Sun², Xiyuan Deng^{4,5},

4 ¹ State Key Laboratory of Hydraulic Engineering Simulation and Safety , Tianjin
5 University, Tianjin, China;

6 ² College of Hydraulic Science and Engineering, Yangzhou University, JiangSu,
7 China;

8 ³College of Water Resources, North China University of Water Resources and Electric
9 Power, Zhengzhou 450046, China;

10 ⁴Nanjing Hydraulic Research Institute, Nanjing, 210029, China; ⁵State Key
11 Laboratory of Hydrology-Water Resources and Hydraulic Engineering , Nanjing,
12 210029, China

13 Correspondence to: Min Li (limintju@126.com); Xiyuan Deng (xydeng@nhri.cn)

14

15



16 **Abstract:** Hydrological drought forecasting can mitigate the socio-economic and
17 ecological impacts of drought. It is an important disaster reduction strategy to forecast
18 the occurrence of hydrological drought according to the forecasting system. In this
19 paper, the conditional distribution model with human activity factor as exogenous
20 variable was constructed to forecast the hydrological drought based on
21 meteorological drought, and then compared with the traditional normal distribution
22 model and conditional distribution model. The results show that the runoff series of
23 Luanhe River Basin from 1961 to 2010 was non-stationary. For the traditional
24 conditional probability models, the transition probabilities of drought were affected
25 by SPI time scales and forecasting periods. In order to analyze the impact of human
26 activities on hydrological drought, we constructed the human activity factor based on
27 the method of restoration. Subsequently, the conditional distribution models involving
28 human index were constructed and the influence of human activities on drought
29 transition probability was analyzed. With the increase of human index (*HI*) value,
30 hydrological droughts tend to transition to more severe droughts. Finally, a scoring
31 mechanism was applied to evaluate the performance of three drought forecasting
32 models. According to the scores of the three drought forecasting models, the
33 conditional distribution model involving of human activity factor can further improve
34 the forecasting accuracy of drought in Luanhe River Basin.

35 **Keyword:** Changing environment; Drought forecasting; Human activity factor;
36 Luanhe River basin

37

38 **1 Introduction**

39 Typically, meteorological drought is regarded as the beginning of a drought event;
40 after the occurrence of meteorological drought, other drought phenomena occur, such
41 as hydrological drought (Miriam et al., 2018; Fuentes et al., 2022; Wang et al., 2021).
42 However, there is a delay period from meteorological drought to hydrological drought
43 (Ding et al., 2021; Xu et al., 2019). Therefore, the occurrence of hydrological drought
44 can be forecasted according to meteorological drought monitoring. Accurate
45 hydrological forecast information is beneficial to reduce the losses caused by
46 hydrological drought. (Behzad and Hamid, 2019; Melanie et al., 2018).

47 Statistical technology is an effective prediction method that has been widely used
48 in drought forecasting in recent years (Bonaccorso et al., 2015). Focusing on
49 statistical techniques, several mathematical statistical models have been applied to
50 forecast drought, such as neural network models (Mehdi et al., 2016; Maryam et al.,
51 2017; Ahnadi et al., 2011), time series modelling (Mohammad et al., 2020; Natsagdorj



52 et al., 2021; Stojković et al., 2020) and hybrid models (Alquraish et al., 2021; Abbasi
53 et al., 2021; Bagher et al., 2013). Some scholars focus on the transition probability of
54 the drought class, which is mainly based on a certain drought index, such as the
55 standardized precipitation index (SPI), the Palmer drought severity index (PDSI) or
56 the standardized runoff index (SRI) (McKee, 1993; Palmer, 1965; Shukla, 2008).
57 Mallya et al. (2013) assessed a drought probability based on a hidden Markov model
58 (HMM) and then analysed the drought characteristics of Indiana. Moreira et al. (2013)
59 calculated the SPI time series in the Alentejo area from 1932 to 1999, and then
60 loglinear models were fitted to assess drought class transition probabilities. Based on
61 the Multivariate Standardized Precipitation Index (MSPI), Aghelpour and Varshavian
62 (2021) proposed the hybrid model to forecast the hydrological drought in Iran, which
63 significantly improved the forecasting accuracy. Majid et al. (2019) used
64 Archimedean copulas to model the relationship between the SPI and standardized
65 hydrological drought index (SHDI), and the results indicated that hydrological
66 drought class forecasting in the coming month is promising with less than 10% error.

67 Considering the impact of the changing environment, Bonaccorso et al. (2015)
68 calculated SPI values under distinct time scales and analysed the conditional
69 probabilities from the current SPI values to the future SPI classes. Ren et al. (2017)
70 found that a model using large-scale climatic indices as covariates can improve the
71 accuracy of meteorological drought forecasting in the Luanhe River Basin. Although
72 some progress has been made in the study of drought forecasting, there are few
73 studies considering the impact of changing environments.

74 To date, some studies have found nonstationary characteristics in the
75 hydrological series of the Luanhe River Basin under changes in the environment
76 (Wang et al., 2018; Li et al., 2015; Wang et al., 2016). The nonstationarity of
77 hydrological series may lead to the nonstationarity of the relationship between
78 hydrological series (for example, precipitation and runoff series), and traditional
79 drought prediction methods are no longer applicable (Wang et al., 2022; Dixit et
80 al., 2022; Muhammad et al., 2020; Zhao et al., 2018; Charles, 2017; Carmelo and Jü



81 rgen, 2018).

82 The research contents of this paper are as follows: (1) The SPI series and SRI
83 series are calculated according to the monthly rainfall and runoff data of the Luanhe
84 River Basin from 1961 to 2010. (2) A multivariate normal distribution model (Model
85 1), conditional distribution model (Model 2) and conditional distribution model with
86 the human index (*HI*) as an exogenous variable (Model 3) were constructed to
87 calculate the transition probabilities from current SPI classes or values to future SRI
88 classes. (3) A scoring mechanism was applied to the evaluation of the three
89 probability models.

90 In addition to the introduction, this paper also contains the following sections.
91 Section 2 introduces the study area and data. Section 3 briefly describes the methods
92 used in the research. Section 4 introduces the model construction and calculation
93 results and analyses the results. Section 5 presents the prospects.

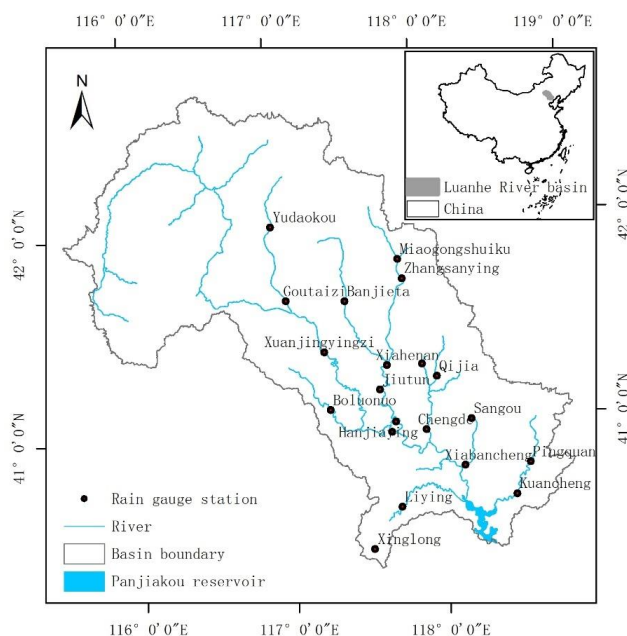
94 **2 Study area and data**

95 The Luanhe River Basin, located in the subtropical monsoon region, covers an
96 area of approximately 33700 square kilometres. Its geographical location is shown in
97 Figure 1. Due to the influence of geographical location and topography, the annual
98 average north-south temperature difference in the basin is 11.5 °C, and the annual
99 rainfall distribution is uneven. Less rain in spring and winter makes the area prone to
100 meteorological drought and hydrological drought, while there is relatively more
101 rainfall in summer. The average rainfall in summer is approximately 200-560 mm,
102 resulting in highly variable annual runoff of the basin. The concentrated rainfall in
103 summer has also become one of the remarkable features of the climate in this area. In
104 recent years, the precipitation and inflow of the Luanhe River Basin have gradually
105 decreased, the water level of the Panjiakou Reservoir in the lower reaches of the basin
106 has decreased, the runoff has also decreased, and the frequency of meteorological
107 drought and hydrological drought has significantly increased. Especially after entering
108 the 21st century, the river basin has exhibited the phenomenon of continuous drought
109 and even extreme drought. With the change in the global climate and the impact of



110 human activities on the basin environment, drought disasters in the Luanhe River
111 Basin occur frequently, causing significant social and economic losses.

112 In this paper, the monthly rainfall data from 26 stations in the Luanhe River
113 Basin from 1961 to 2010 are provided by the Hebei Provincial Hydrology and Water
114 Resources Investigation Bureau. The average monthly rainfall data of the area are
115 obtained by spatial interpolation. The runoff data from 1961 to 2010 come from the
116 inflow runoff series of the Panjiakou Reservoir. The SPI and SRI can be calculated for
117 1-month, 3-month, 6-month, and 12-month time scales to characterize meteorological
118 drought and hydrological drought based on these data.



119
120 Figure 1 The geographical location of the Luanhe River Basin

121 3 Methods

122 3.1 Nonstationarity test method

123 In the case of environmental changes, nonstationarity may occur in hydrological
124 series. The Pettitt test, as one of the important methods to test whether there is
125 nonstationarity in time series, can identify whether there are change points in the



126 sample series (Malede et al., 2022). Assuming that the sample sequence is
 127 $x = (x_1, x_2, \dots, x_n)$, the formula is as follows:

$$128 \quad U_{t,n} = U_{t-1,n} + \sum_{i=1}^n \text{sgn}(x_t - x_i) \quad (t = 2, 3, \dots, n)t_0 \quad (1)$$

$$129 \quad \text{sgn}(x_t - x_i) = \begin{cases} 1 & x_t - x_i > 0 \\ 0 & x_t - x_i = 0 \\ -1 & x_t - x_i < 0 \end{cases} \quad (2)$$

130 where $U_{t,n}$ is the test statistic, which indicates the cumulative number of the
 131 values at time t greater than or less than the values at time i . In addition, if $K_{t_0,n}$
 132 satisfies:

$$133 \quad K_{t_0,n} = \max |U_{t,n}| \quad (t=1,2,\dots,n) \quad (3)$$

134 Then, t_0 is considered to be the change point, and the cumulative probability of
 135 possible change is determined by $K_{t_0,n}$:

$$136 \quad P_{t_0,n} = 2 \exp\left(-\frac{6K_{t_0,n}^2}{n^3 + n^2}\right) \quad (4)$$

137 Given the significance level $\alpha=0.05$, if $P_{t_0,n} > 0.95$, it means that the point is a
 138 significant change point (Li et al., 2022; Koudahe et al., 2018). Furthermore,
 139 combined with the Mann-Kendall test, the trend characteristics of the sample series
 140 can be obtained (Linchao et al., 2018).

141 The sliding T test is a basic method commonly used in statistics. According to
 142 the mean and variance of the two sample sequences before and after the change points
 143 in the runoff time series, the two sample sequences are tested (Li et al., 2020):

$$144 \quad t = \frac{\bar{x}_1 - \bar{x}_2}{S \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \quad (5)$$

$$145 \quad S = \sqrt{\frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}} \quad (6)$$

$$146 \quad S_1^2 = \frac{1}{n_1 - 1} \sum_{t=1}^{n_1} (x_t - \bar{x}_1)^2 \quad (7)$$

$$147 \quad S_2^2 = \frac{1}{n_2 - 1} \sum_{t=1}^{n_1+n_2} (x_t - \bar{x}_2)^2 \quad (8)$$



148 Here, assume that the change point is x_t , n_1 and n_2 represent the sample size
149 before and after the change point, S_1^2 and S_2^2 represent the variance of the samples
150 before and after the change point, respectively. If the statistic t satisfies $t > t_\alpha$ as the
151 significance level is $\alpha = 0.05$, the point can be considered the change point.

152 The Spearman correlation test can be applied to test the trend of time series, and
153 the specific description refers to the article of Bishara and Hittner (2012).

154 3.2 Human activity index (*HI*)

155 The double cumulative curve method can test the nonstationarity of the bivariate
156 correlation between rainfall series and runoff series, and the point where the
157 underlying surface is significantly altered by human activities can be determined
158 according to the position of the slope change of the curve. The linear regression
159 relationship of the cumulative rainfall and runoff series can be calculated according to
160 the following formula:

$$161 \quad \sum x = k \sum y + b \quad (9)$$

162 Here, x is the runoff series; y is the rainfall series; k is the correlation coefficient of the
163 regression equation; and b is the intercept of the regression equation.

164 Human activities are the main reason for the nonstationarity of the runoff series
165 in the watershed, so the human activity index (*HI*) can be constructed to quantify the
166 impact of human activities on runoff. Based on the linear regression relationship
167 established between the accumulated precipitation and the accumulated runoff before
168 the change point, the theoretical runoff sequence during the human activity period can
169 be calculated from the measured precipitation sequence. SRI' represents the
170 standardized runoff index value without human activity interference, and SRI
171 represents the normalized runoff index value calculated based on the measured runoff
172 sequence under the disturbance of human activities. The *HI* is obtained by subtracting
173 the theoretical SRI' and the actual SRI , and the calculation formula is as follows:

$$174 \quad HI = SRI' - SRI \quad (10)$$



175 When $HI > 0$, it can be assumed that human activities exacerbate hydrological
 176 drought, $HI < 0$ has the opposite effect, and $HI = 0$, the watershed is considered
 177 undisturbed by human activities.

178 3.3 Multivariate normal distribution model

179 The SPI is one of the important indicators for evaluating meteorological drought
 180 in the basin, and the SRI is an important indicator for evaluating hydrological drought
 181 in the basin. According to the rainfall data and runoff data in the basin, the SPI and
 182 SRI at different time scales can be calculated. Table 1 provides the drought class
 183 classification and corresponding SPI values and SRI values (Kolachian and Saghafian,
 184 2021).

185 Table 1 Drought class classification and corresponding SPI values and SRI values

| SPI/SRI values | Class |
|--------------------|----------|
| > -0.99 | Normal |
| -1.00 to -1.49 | Moderate |
| -1.50 to -1.99 | Severe |
| ≤ -2.00 | Extreme |

186 As a traditional drought class forecasting model, the multivariate normal
 187 distribution model (Model 1) can forecast the future SRI class according to the current
 188 SPI class. Assuming that both the current SPI and SRI sequence satisfy a
 189 multivariable normal distribution, the joint probability density can be expressed as
 190 follows (Chang et al., 2022):
 191 follows (Chang et al., 2022):

$$192 \quad f_{Z_{v,\lambda}^{(k)} W_{v,\lambda+M}^{(k)}}(t, s) = \frac{1}{2\pi|\Sigma|} \cdot \exp\left(-\frac{1}{2} X^T \Sigma^{-1} X\right) \quad (11)$$

193 Here, Σ is the covariance matrix, and $X = [t, s]^T$. The form of the covariance
 194 matrix is as follows:

$$195 \quad \Sigma = \begin{bmatrix} 1 & \text{cov}[Z_{v,\lambda}^{(k)}, W_{v,\lambda+M}^{(k)}] \\ \text{cov}[Z_{v,\lambda}^{(k)}, W_{v,\lambda+M}^{(k)}] & 1 \end{bmatrix} \quad (12)$$

196 Furthermore, according to the joint probability density function of the SPI value
 197 $Z_{v,\lambda}^{(k)}$ at v year and month λ and the future M months SRI value $W_{v,\lambda+M}^{(k)}$, the analytical



198 formula of the transition probability of the future SRI drought class can be obtained
 199 (Zhang et al., 2017):

$$200 \quad P\left[W_{v,\lambda+M}^{(k)} \in C_M\right] = \frac{\iint_{C_N C_M} f_{Z_{v,\lambda}^{(k)} W_{v,\lambda+M}^{(k)}}(t, s) \cdot dt \cdot ds}{\int_{C_N} f_{Z_{v,\lambda}^{(k)}}(t) \cdot dt} \quad (13)$$

201 where C_M represents the drought class and $f_{Z_{v,\lambda}^{(k)}}(t)$ represents the marginal
 202 density function of $Z_{v,\lambda}^{(k)}$ in the current λ month.

203 3.4 The conditional distribution model

204 The conditional distribution model (Model 2) proposed by Bonaccorso et al.
 205 (2015) is described as follows: when one group of sample data X obeys a normal
 206 distribution and satisfies $X \sim N(\mu_1, \Sigma_1)$, while another group of sample data Y also
 207 obeys a normal distribution, namely, $Y \sim N(\mu_2, \Sigma_2)$, then the total sequence can be
 208 written as follows:

$$209 \quad B = \begin{bmatrix} X \\ Y \end{bmatrix} \begin{matrix} r \\ p-r \end{matrix} \sim N_p \left(\begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} \right) \quad (14)$$

210 When sequence Y obeys a normal distribution, the distribution of sequence X
 211 under the Y condition still satisfies a normal distribution, namely, the distribution of
 212 $(X | Y)$ is as follows (Gong et al. 2021):

$$213 \quad (X | Y) \sim N(\mu_3, \Sigma_3) \quad (15)$$

214 where μ_3 represents the expected value under the conditional distribution, and
 215 Σ_3 is the conditional covariance matrix:

$$216 \quad \mu_3 = \mu_1 + \Sigma_{12} \Sigma_{22}^{-1} (y - \mu_2) \quad (16)$$

$$217 \quad \Sigma_3 = \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21} \quad (17)$$

218 Assuming that the current SPI and the SRI sequence that transitioned from
 219 meteorological drought satisfy a binary normal distribution, then the probability of the
 220 transition to the future SRI drought class under the current SPI value can be deduced
 221 as follows (Ren et al., (2017)):



$$P[W_{v,\lambda+M} \in C_M | Z_{v,\lambda} = z_0] = \int_{C_{Mi}}^{C_{Ms}} \frac{1}{\sqrt{2\pi}\sigma_Z} e^{-\frac{1}{2}\left(\frac{x-\rho z_0}{1-\rho^2}\right)^2} dx \quad (18)$$

where $Z_{v,\lambda}$ represents the SPI value of the current month λ , $W_{v,\lambda+M}$ represents the SRI value of the $\lambda + M$ month, C_{Ms} and C_{Mi} are the upper and lower limits of the drought class C_M , and the correlation coefficient between the current SPI value and the future SRI value is ρ . Furthermore, the current SPI and future SRI can be expressed as the standard normal cumulative distribution function Φ :

$$P[W_{v,\lambda+M} \in C_M | Z_{v,\lambda} = z_0] = \Phi\left[\frac{C_{Ms} - \rho z_0}{1 - \rho^2}\right] - \Phi\left[\frac{C_{Mi} - \rho z_0}{1 - \rho^2}\right] \quad (19)$$

The calculation of the correlation coefficient ρ is as follows:

$$\rho = \frac{\text{cov}[Z_{v,\lambda}^{(k)}, W_{v,\lambda+M}^{(k)}]}{\sqrt{\text{var}(Z_{v,\lambda}^{(k)}) \text{var}(W_{v,\lambda+M}^{(k)})}} \quad (20)$$

K represents the time scale of the drought index. Assuming that the cumulative rainfall Y and runoff X satisfy a normal distribution, then after the standardization process, the SPI value $Z_{v,\lambda}^{(k)}$ corresponding to cumulative rainfall Y and SRI value $W_{v,\lambda+M}$ corresponding to runoff X obey the standard normal distribution (Wu, 2019), namely:

$$\text{var}(Z_{v,\lambda}^{(k)}) = \text{var}(W_{v,\lambda+M}^{(k)}) = 1 \quad (21)$$

$\text{cov}[Z_{v,\lambda}^{(k)}, W_{v,\lambda+M}^{(k)}]$ represents the covariance between the current SPI and the SRI value with a forecast period of M months. The calculation is as follows:

$$\text{cov}[Z_{v,\lambda}^{(k)}, W_{v,\lambda+M}^{(k)}] = \frac{1}{\sqrt{\sum_{i=0}^{k-1} \sigma_{\lambda+M-i}^2 \sum_{j=0}^{k-1} \sigma_{\lambda-j}^2}} \cdot \sum_{i=0}^{k-1} \sum_{j=0}^{k-1} \text{cov}[X_{v,\lambda+M-j}, Y_{v,\lambda-i}] \quad (22)$$

3.5 The conditional distribution model involving HI as an exogenous variable

According to the above conditional probability model, when considering HI as an exogenous variable, the model (Model 3) can be extended as follows:



$$P[W_{v,\lambda+M} \in C_M | Z_{v,\lambda} = z_0, H_{v,\lambda} = h_0] = \int_{C_M} \frac{1}{\sqrt{2\pi}\sigma_z} e^{-\frac{1}{2}\left(\frac{x-\mu_z}{\sigma_z}\right)^2} dx \quad (23)$$

$$\mu_z = E[W_{v,\lambda+M} | Z_{v,\lambda}, H_{v,\lambda}] = \Sigma'_{12} (\Sigma'_{22})^{-1} \begin{bmatrix} z_0 \\ h_0 \end{bmatrix} \quad (24)$$

$$\sigma_z^2 = \text{var}[W_{v,\lambda+M} | Z_{v,\lambda}, H_{v,\lambda}] = 1 - \Sigma'_{12} (\Sigma'_{22})^{-1} \Sigma'_{21} \quad (25)$$

Where:

$$\Sigma'_{12} = [\text{cov}(W_{v,\lambda+M}, Z_{v,\lambda}) \quad \text{cov}(W_{v,\lambda+M}, H_{v,\lambda})] \quad (26)$$

$$\Sigma'_{22} = \begin{bmatrix} 1 & \text{cov}(Z_{v,\lambda}, H_{v,\lambda}) \\ \text{cov}(H_{v,\lambda}, Z_{v,\lambda}) & 1 \end{bmatrix} \quad (27)$$

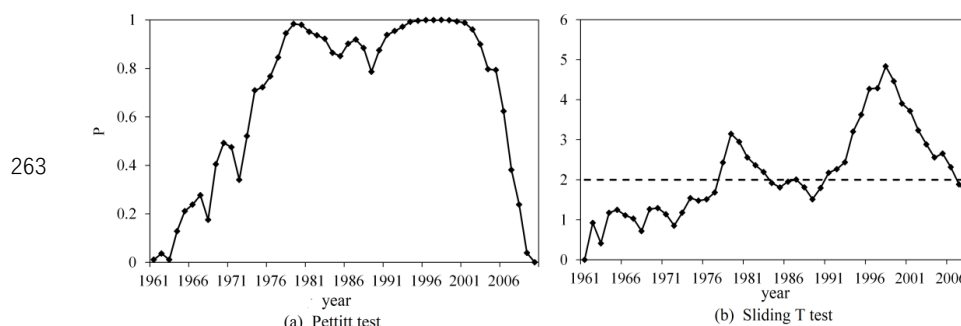
$$\Sigma'_{21} = (\Sigma'_{12})^T \quad (28)$$

250 4 Results and discussion

251 4.1 Nonstationarity analysis

252 In this paper, the area average monthly rainfall data of the Luanhe River Basin
 253 from 1961 to 2010 are obtained by spatial interpolation. The runoff data come from
 254 the inflow runoff series of the Panjiakou Reservoir. Given the significance level
 255 $\alpha = 0.05$, the nonstationarity test results are shown in Figure (2).

256 Figure 2 (a) shows that the years of possible runoff change were 1979, 1996,
 257 1997, 1998, and 1999. The P values in 1979 and 1998 were infinitely close to 1,
 258 which were considered to be extremely significant runoff change points. Among all
 259 the possible points satisfying $t > t_\alpha$, there are two maximum points (Figure 2 (b)),
 260 namely, 1979 and 1998, which are considered to be possible runoff change points.
 261 The final change point needs to be judged based on the actual situation of the
 262 watershed.



263

264

Figure (2) The change points of the runoff series

265

266

267

268

269

The results of the Spearman correlation test (Table 2) indicate that the runoff series showed an upwards trend before 1979, but the trend was not significant. However, there was a significant downwards trend in the series after 1979. In general, the runoff series showed a significant downwards trend.

Table 2. Spearman correlation test results of runoff series trend

| Runoff series | statistic t | Critical value t_{α} |
|-------------------|---------------|-----------------------------|
| The whole series | -3.471 | ± 2.009 |
| Serie before 1979 | 0.691 | ± 2.009 |
| Serie after 1979 | -2.292 | ± 2.009 |

270

271

272

273

274

275

276

277

278

279

In addition, according to historical records, there were no extreme rainstorm events recorded during 1979. It can be inferred that the cause of the sudden change in annual runoff in 1979 was not the formation of heavy rainstorms in the previous period or the same period. Since the start of 1979, the underlying surface conditions of the basin have undergone large changes due to human activities, so it is determined that 1979 is the change point of the runoff sequence in the basin. Therefore, 1979 was finally determined as the change point of the runoff sequence of the Luanhe basin from 1961 to 2010. This conclusion is consistent with Li et al. (2015) and Wang et al. (2015).

280

4.2 Transition probabilities from current SPI values to future SRI classes

281

282

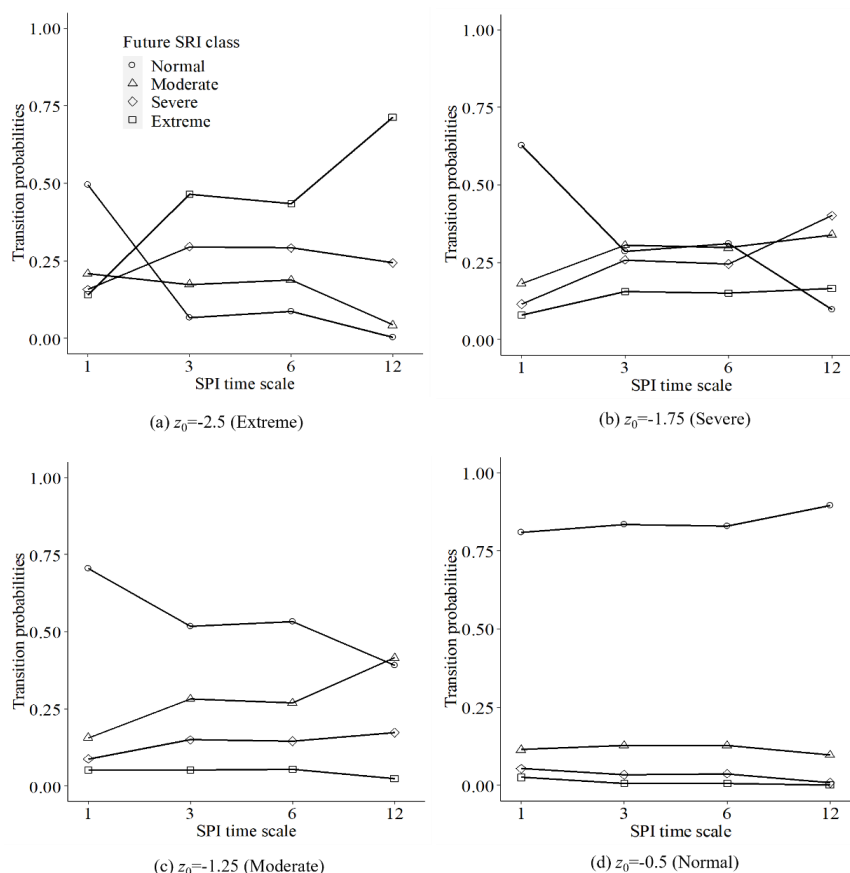
283

According to the normality test results of rainfall and runoff series, it is reasonable to apply the conditional distribution model. To analyse the influence of different time scales of SPI on the transition probabilities, using the forecast period as



284 one month and the time scales of SPI on 1-month, 3-month, 6-month and 12-month as
 285 examples, the probabilities of converting SPI values to SRI classes were calculated
 286 (Figure (3)).

287 As shown in Figure (3), when meteorological drought is categorised as extreme
 288 drought, the probability of maintaining the SRI class in the extreme drought state is
 289 greater as the time scale of the SPI increases. As the SPI is a 12-month time scale, the
 290 drought transition probability is close to 1. However, while the time scale is small, the
 291 response of the future SRI value to rainfall is faster, so the probability of tending to
 292 the normal state is greater. In the future, the response of the SRI value to rainfall is
 293 relatively fast, so it is more likely to tend to a normal state.

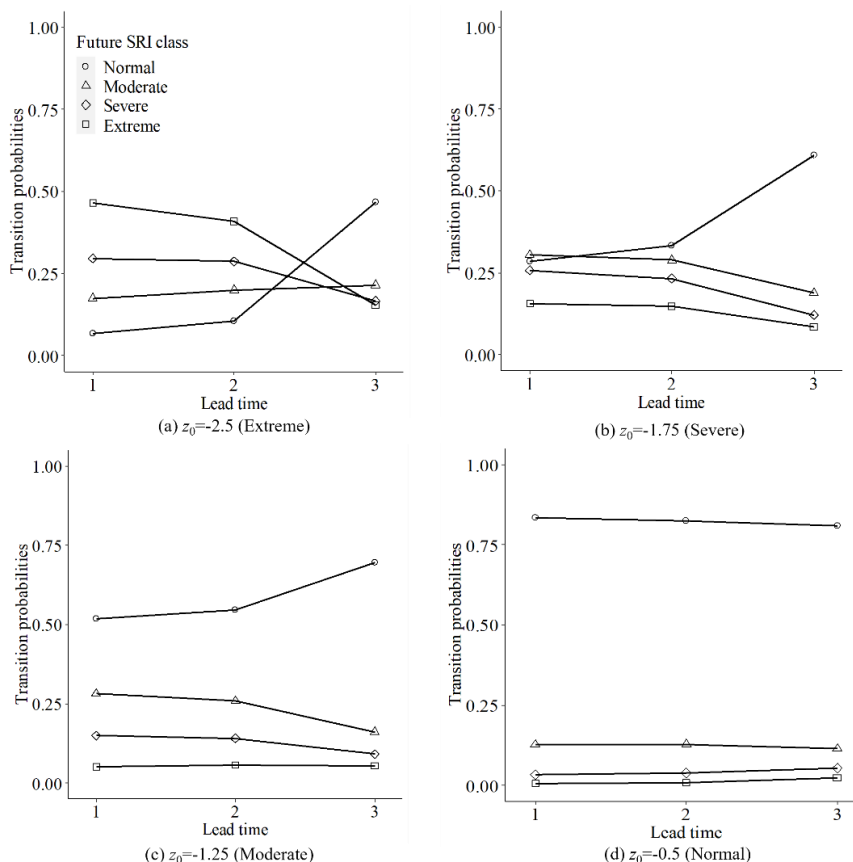


294
 295

Figure (3). Influence of the SPI time scale on transition probabilities (z_0 : initial value of SPI)



296 In addition, the transition probabilities of drought are distinct for different
297 forecast periods. As seen in Figure 4(a), while the current period $z_0 = -2.5$, i.e., the
298 meteorological drought is extreme drought and the forecast period is 1 and 2 months,
299 the probability of its future SRI class being extreme drought is the highest. Moreover,
300 the probability of its future SRI drought class returning to normal status becomes
301 higher as the forecast period becomes longer. When the current period $z_0 = -1.75$
302 (Figure 4 (b)), namely, the meteorological drought is severe drought and the forecast
303 period is 1 month, its future SRI class tends to be normal or moderate drought. While
304 the forecast period becomes longer, its drought degree gradually decreases and tends
305 to be normal. When the current period $z_0 = -1.25$ (Figure 4 (c)), namely, the
306 meteorological drought is a moderate drought, the future SRI class tends to be a
307 normal or moderate drought state as the forecast period is 1 month. In addition, its
308 drought degree gradually decreases and tends to be normal, while the forecast period
309 becomes longer. It is worth noting that the current $z_0 = 0$ (Figure 4 (d)), and the
310 probability that the future SRI class is normal as the forecast period is 1, 2 and 3
311 months is greater than 0.8.



312

313

Figure (4) Influence of forecast period on transition probabilities (z_0 : initial value of SPI)

314

315

316

317

318

319

320

From the above analysis, when the forecast period is short ($M=1$ or 2), the hydrological drought class obtained from the transition of meteorological drought tends to be the same as that of meteorological drought. With the extension of the forecast period ($M=2$ or 3), the overall SRI class obtained from the transition tends to be slightly lower than the SPI drought class or to the normal state, i.e., the hydrological drought class obtained from the transition tends to be slightly lighter than the meteorological drought on the whole or to be maintained in the normal state.

321

4.3 Transition probabilities with involving HI as the covariate

322

323

According to the above nonstationarity test results, 1979 was the change point, and the linear regression relationship of the cumulative rainfall and runoff series

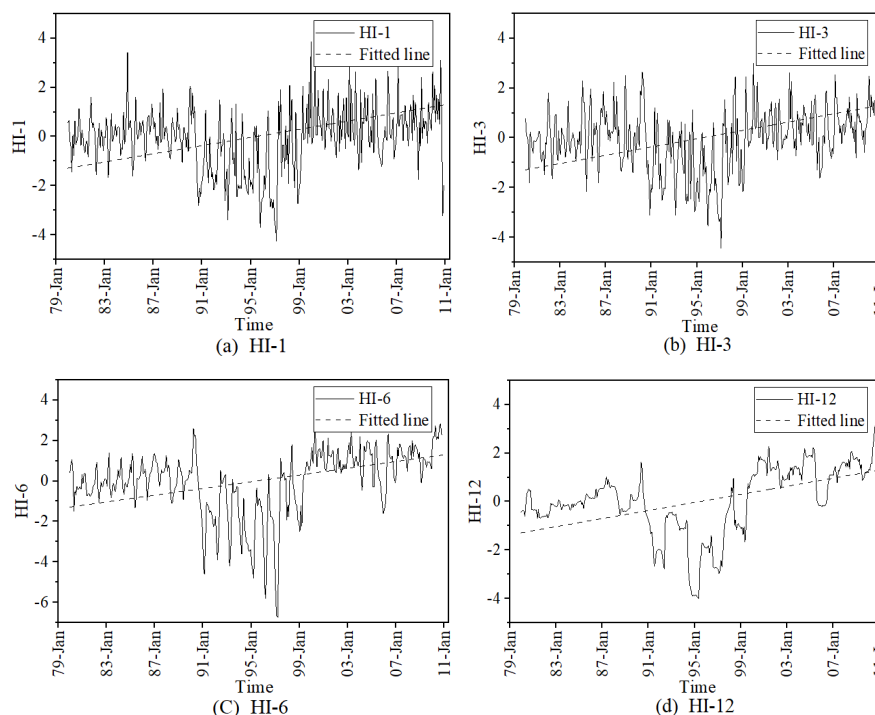


324 before and after the change point were established. The calculation results are shown
 325 in Table 3:

326 Table 3 Linear regression relationship between cumulative precipitation (x / mm) and cumulative
 327 runoff (y / 10^6 m^3)

| Period | Linear regression equation | Correlation coefficient |
|-----------|----------------------------|-------------------------|
| 1961~1979 | $x = 0.0276y + 2.7566$ | 0.99 |
| 1980~2010 | $x = 0.0307y - 30.652$ | 0.98 |

328 The HI results for different time scales are shown in Figure 5.



329 Figure 5 Different average periods of *HI* (*HI-1*: *HI* with 1-month time scale; *HI-3*: *HI* with 3-
 330 month time scale; *HI-6*: *HI* with 6-month time scale; *HI-12*: *HI* with 12-month time scale)
 331

332 As shown in Figure 5, the *HI* at all monthly scales generally ranges upwards,
 333 which means that human activities have intensified the occurrence of hydrological
 334 drought.

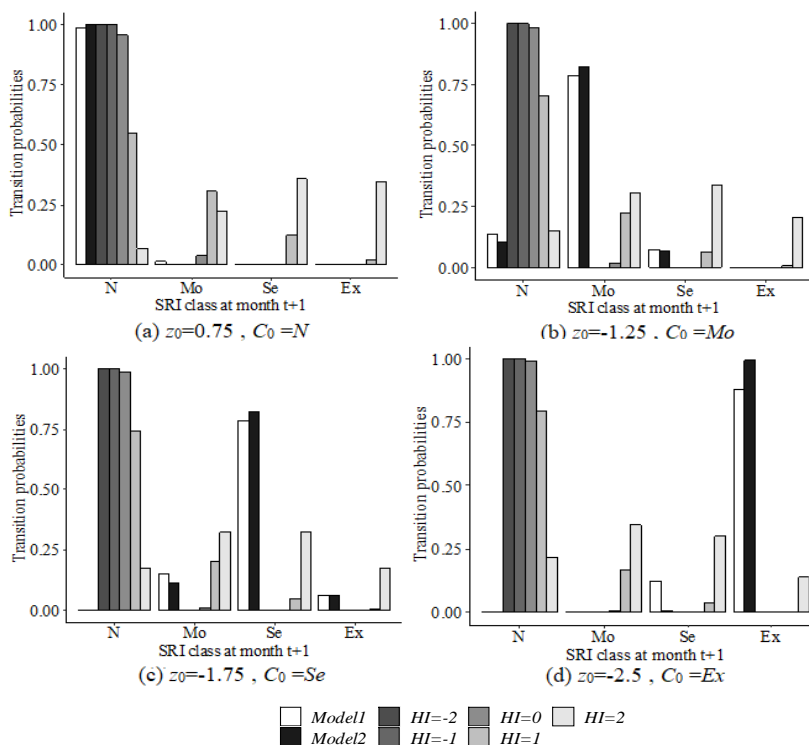
335 The *HI* of different monthly scales were standardized, taking the 12-month time
 336 scale as an example, and the results were calculated as shown in Table 3.

337 Table 3. *HI-12* Monthly Mean and Standard Deviation



| | J | F | M | A | M | J | J | A | S | O | N | D |
|------|-------|-------|-------|-------|-------|------|------|------|------|------|------|------|
| Mean | -0.04 | -0.03 | -0.03 | -0.03 | -0.03 | 0.00 | 0.06 | 0.06 | 0.10 | 0.10 | 0.09 | 0.06 |
| Sd | 1.36 | 1.37 | 1.38 | 1.41 | 1.41 | 1.51 | 1.40 | 1.40 | 1.45 | 1.44 | 1.44 | 1.43 |

338 Furthermore, the drought transition probabilities involving *HI* can be calculated
 339 from Eq. (23). Using the forecast period of one month from December and the SPI
 340 time scale of 12 months as an example, the drought transition probabilities from
 341 current SPI values to future SRI classes can be calculated (Figure 6). To analyse the
 342 effect of human activities on the drought transition probability more clearly, the
 343 calculation results of the three models are compared here separately. The horizontal
 344 coordinate indicates the drought classes corresponding to the SRI for the coming
 345 month, and the vertical coordinate is the drought transition probability.



346 Figure 6 Drought transition probability under the influence of human activities (C_0 denotes the
 347 initial drought class of SPI in the multivariate normal model; z_0 represents the initial value of SPI
 348 in the conditional distribution model; Model 1: The normal distribution model; Model 2: The
 349 conditional distribution model; Model 3: The conditional distribution model involving *HI*)
 350



351 In Figure 6 (a), when the initial $z_0=0.75$ and $c_0=N$, the results shown in Model 1
352 and Model 2 are similar, and the probability transitions of SPI values to SRI classes in
353 the future month in the normal class are close to 1. However, the results of Model 3
354 indicate that the probabilities of maintaining SRI in the normal class in the future
355 decrease as HI increases. When $HI=2$, the probability of transition to severe drought
356 or extreme drought is higher.

357 From the initial $z_0=-1.25$ and $c_0=Mo$ (Figure 6 (b)), it can be seen from the
358 results of Model 3 that the transition probabilities of SPI values to a normal SRI class
359 in the coming month are higher when HI is less than 1. As the HI increases, the
360 transition probabilities of the SPI values to a moderate drought or even a more severe
361 drought in the future increase. In addition, the probabilities of maintaining drought at
362 moderate drought are the highest when human activities are not considered, and
363 Model 2 shows a higher probability.

364 While the initial meteorological drought class is a severe drought (Figure 6 (c)),
365 the probabilities of the future SRI drought class being in the normal class become
366 larger as the HI decreases. When the effect of human activities is not considered, the
367 probability that the current SPI value transitions to the SRI class under severe drought
368 in the future month is the highest, and the probability of being in the normal class is
369 the lowest. For Model 1, the probability of the SRI classes transitioning to a moderate
370 drought is higher than the result of Model 2.

371 It is noteworthy that when the initial $z_0=-2.5$ and $c_0=Es$ (Figure 6 (d)), the
372 probabilities of transition of the SPI values to future SRI classes at the normal class
373 are close to 1 as $HI<0$. However, hydrological drought is more likely to be moderate
374 drought or severe drought as HI are greater than 0, and the transition probabilities
375 exceed 0.25. For Model 1 and Model 2, the probabilities of transition of current SPI
376 values or classes to the future month SRI classes also in extreme drought are both
377 higher than 0.75. Model 1 shows a higher probability than Model 2 when the SRI
378 class transitions to severe drought.



379 4.4 Model evaluation and analysis

380 To quantitatively evaluate the prediction accuracy of Model 1, Model 2 and
 381 Model 3, the study period is divided into a correction period (1961-2003) and a
 382 verification period (2004-2010), and then the drought transition probability from the
 383 SPI value or class to the SRI class in the future M -month is calculated. The monthly
 384 drought transition probability is summed to evaluate the model (Chen et al., 2013):

$$385 \quad \text{Score} = \frac{1}{12n} \sum_{t=1}^{12} \sum_{s=1}^n p_{s,t} \quad (29)$$

386 where $p_{s,t}$ characterizes the transition probability in month t of year s , and n
 387 is the length of the validation period. The calculation results are shown in Table 5.

388 With the same time scale of SPI, the model scores of Model 1 and Model 2
 389 lowers as the forecast period M lengthens, while the model scores of Model 3 are not
 390 significantly affected by the forecast period M . Model 1 had the highest rating of 0.36
 391 at an SPI of 1-month time scale and a forecast period of one month; Model 2 reached
 392 the highest model rating of 0.74 at a 12-month time scale and a forecast period of one
 393 month; and model-3 performed well at an SPI of 1-month time scale and a 12-month
 394 time scale. Overall, model-3 has the highest rating, and Model 1 has the lowest rating
 395 for the same SPI time scale and the same forecast period, which also indicates that the
 396 forecast accuracy of the conditional distribution model considering the HI is higher
 397 for short-term forecasts with a forecast period of 3 months or less, and involving the
 398 HI can further improve the forecast accuracy of the model.

399 Table 5. Model Evaluation (Model 1: Multivariate normal distribution model; Model 2:
 400 Conditional distribution model; Model 3: Conditional distribution model with HI)

| Model type | Lead time M | SPI time scale | | | |
|------------|---------------|----------------|------|------|------|
| | | 1 | 3 | 6 | 12 |
| Model 1 | 1 | 0.36 | 0.36 | 0.28 | 0.22 |
| | 2 | 0.11 | 0.35 | 0.27 | 0.22 |
| | 3 | 0.02 | 0.34 | 0.26 | 0.22 |
| Model 2 | 1 | 0.69 | 0.52 | / | 0.74 |
| | 2 | 0.69 | 0.47 | / | 0.67 |
| | 3 | 0.69 | 0.44 | 0.39 | 0.60 |



| | | | | | |
|---------|---|------|------|------|------|
| | 1 | 0.72 | 0.64 | 0.59 | 0.71 |
| Model 3 | 2 | 0.71 | 0.64 | 0.59 | 0.71 |
| | 3 | 0.72 | 0.64 | 0.60 | 0.71 |

401 5 Conclusions

402 Many studies have pointed out that human activities have a significant impact on
403 watershed runoff in the Luanhe River Basin. In this paper, three probability models
404 were constructed to calculate the transition probabilities from current SPI classes or
405 values to future SRI classes; then, a scoring mechanism was applied to evaluate the
406 performance of the models.

407 Under the condition of considering the *HI*, the calculation results of the drought
408 transition probability show that when the value of *HI* is less than 0, human activity
409 slows the occurrence of hydrological drought and the probability of maintaining
410 hydrological drought at the normal class peaks. With the increase in the *HI* value, it is
411 easier for hydrological droughts to transition to more severe droughts. The calculation
412 results of Model 1 and Model 2 show that the future hydrological drought classes are
413 likely to be the same as the meteorological drought classes in the current period.

414 Finally, a scoring mechanism was applied to the evaluation of the models, and
415 the forecast results of the three models were evaluated. The results demonstrate that
416 when the SPI time scale is the same, the scores of Model 1 and Model 2 lower as the
417 forecast period lengthens. In most cases, Model 2 performs better than Model 1, and
418 the performance of Model 3 is the most stable of the three models and has the highest
419 score. In addition, the performance of Model 3 is not affected by the forecasting
420 period. The conditional probability model considering *HI* is more suitable for the
421 Luanhe River basin, where human activities have a high influence.

422 Although this study has made some progress in the forecasting of hydrological
423 drought in the changing environment, only one exogenous variable was calculated to
424 quantify the impact of human activities, and the climate factors can be further
425 considered in future studies. In addition, *HI* can be analysed specifically, such as land
426 use and social economy.



427 **Limitation:** Under changing environmental conditions, the driving factors of drought
428 can be analysed from the physical mechanism, such as considering the influence of
429 large-scale climate indices or hydro-meteorological variables, to further improve the
430 forecasting accuracy of hydrological drought.

431 **Ethical Approval:** This work meets the ethical and moral requirements.

432 **Consent to Participate:** M L. MF Z. RX C. YD S and XY D all agreed to participate
433 in the research for the article.

434 **Consent to Publish:** M L. MF Z. RX C. YD S and XY D all agreed to publish this
435 article.

436 **Authors Contributions:**

437 M L(First Author and Corresponding Author): Conceptualization, Methodology,
438 Software, Investigation, Formal Analysis, Writing-Original Draft;

439 MF Z: Data Curation, Writing-Original Draft;

440 RX C: Visualization, Investigation;

441 YD S: Resources, Supervision;

442 XY D: Visualization, Writing-Review & Editing.

443 **Competing interest:** M L. MF Z. RX C. YD S and XY D all declare that there is no
444 conflict of interest.

445 **Data availability statement:** We are grateful to the Hydrology and Water Resource
446 Survey Bureau of Hebei Province for providing runoff data. The data and materials of
447 the research are available.

448 **References**

449 [1] Abbasi Abbas,Khalili Keivan,Behmanesh Javad,Shirzad Akbar. Estimation of ARIMA model
450 parameters for drought prediction using the genetic algorithm[J]. Arabian Journal of
451 Geosciences,2021,14(10): 841-841.

452 [2] Ahnadi mahmoud. Climatic drought forecasting using artificial neural network in Hamedan region[J].
453 New York Science Journal,2011,4(8):15-19.



-
- 454 [3] Alquraish Mohammed, Ali. Abuhasel Khaled, S. Alqahtani Abdulrahman, Khadr Mosaad. SPI-Based
455 Hybrid Hidden Markov–GA, ARIMA–GA, and ARIMA–GA–ANN Models for Meteorological Drought
456 Forecasting[J]. Sustainability, 2021, 13(22): 12576-12576.
- 457 [4] Bagher Shirmohammadi, Hamidreza Moradi, Vahid Moosavi, Majid Taie Semiro, Ali Zeinali.
458 Forecasting of meteorological drought using Wavelet-ANFIS hybrid model for different time steps (case
459 study: southeastern part of east Azerbaijan province, Iran)[J]. Natural Hazards, 2013, 69(1): 389-402.
- 460 [5] Behzad Ahmadi, Hamid Moradkhani. Revisiting hydrological drought propagation and recovery
461 considering water quantity and quality[J]. Hydrological Processes, 2019, 33(10): 1492-1505.
- 462 [6] Bishara Anthony J, Hittner James B. Testing the significance of a correlation with nonnormal data:
463 comparison of Pearson, Spearman, transformation, and resampling approaches.[J]. Psychological
464 methods, 2012, 17(3): 399-417.
- 465 [7] Bonaccorso B, Cancelliere A, Rossi G. Probabilistic forecasting of drought class transitions in Sicily
466 (Italy) using Standardized Precipitation Index and North Atlantic Oscillation Index[J]. Journal of
467 Hydrology, 2015, 526: 136-150.
- 468 [8] Carmelo Cammalleri, Jürgen V. Vogt. Non-stationarity in MODIS fAPAR time-series and its impact on
469 operational drought detection[J]. International Journal of Remote Sensing, 2018, 40(4): 1428-1444.
- 470 [9] Chang Guobin, Zhang Shubi, Liu Zhiping. Understanding the adjustment from an information theoretic
471 perspective[J/OL]. Geomatics and Information Science of Wuhan University: 1-17[2022-03-07]. (in
472 Chinese)
- 473 [10] Charles Onyutha. On Rigorous Drought Assessment Using Daily Time Scale: Non-Stationary
474 Frequency Analyses, Revisited Concepts, and a New Method to Yield Non-Parametric Indices[J].
475 Hydrology, 2017, 4(4): 48-48.
- 476 [11] Chen Jie, Brissette François P, Chaumont Diane, Braun Marco. Performance and uncertainty evaluation
477 of empirical downscaling methods in quantifying the climate change impacts on hydrology over two
478 North American river basins[J]. Journal of Hydrology, 2013, 479: 200-214.
- 479 [12] Ding Yibo, Xu Jiatao, Wang Xiaowen, Cai Huanjie, Zhou Zhaoqiang, Sun Yanan, Shi Haiyun. Propagation
480 of meteorological to hydrological drought for different climate regions in China[J]. Journal of
481 Environmental Management, 2021, 283: 111980-111980.
- 482 [13] Dixit, Soumyashree, Jayakumar, K. V.. A Non-stationary and Probabilistic Approach for Drought
483 Characterization Using Trivariate and Pairwise Copula Construction (PCC) Model[J]. Water Resources



-
- 484 Management,2022(prepublish): 1-20.
- 485 [14] Fuentes Ignacio,Padarian José,Vervoort R. Willem. Spatial and Temporal Global Patterns of Drought
486 Propagation[J]. *Frontiers in Environmental Science*,2022:788248-788248.
- 487 [15] Gang Zhao,Huilin Gao,Shih-Chieh Kao,Nathalie Voisin,Bibi S. Naz. A modeling framework for
488 evaluating the drought resilience of a surface water supply system under non-stationarity[J]. *Journal of*
489 *Hydrology*,2018,563: 22-32.
- 490 [16] Gong Haonan, Xie Botao, Wang Junrong, et al. Long-term prediction of extreme response of deepwater
491 floating platform based on environmental contour method[J]. *The Ocean Engineering*,2021,39(05):28-
492 38. (in Chinese)
- 493 [17] Kolachian Roya,Saghafian Bahram. Hydrological drought class early warning using support vector
494 machines and rough sets[J]. *Environmental Earth Sciences*,2021,80(11):390-390.
- 495 [18] Koudahe K,Koffi D,Kayode JA,Awokola SO,Adebola AA. Impact of Climate Variability on Crop Yields
496 in Southern Togo[J]. *Environment Pollution and Climate Change*,2018,2(1): 1-9.
- 497 [19] Li Jianzhu, Wang Y X, Li S F, Hu R. Hu. A Nonstationary Standardized Precipitation Index
498 incorporating climate indices as covariates[J]. *Journal of Geophysical Research:*
499 *Atmospheres*,2015,120(23):082-095.
- 500 [20] Li Linchao, Yao Ning, Li Yi, Liu De Li, Wang Bin, Ayantobo Olusola O. Future projections of extreme
501 temperature events in different sub-regions of China[J]. *Atmospheric Research*,2018,217:150-164.
- 502 [21] Li Xin,Fang Guohua,Wen Xin,Xu Ming,Zhang Yu. Characteristics analysis of drought at multiple
503 spatiotemporal scale and assessment of CIMP6 performance over the Huaihe River Basin[J]. *Journal of*
504 *Hydrology: Regional Studies*,2022,41:101-103.
- 505 [22] Li Xinxin, Ma Xixia, Li Xiaodong, Zhang Wenjiang. Method Consideration of Variation Diagnosis and
506 Design Value Calculation of Flood Sequence in Yiluo River Basin, China[J]. *Water*,2020,12(10): 2722-
507 2722.
- 508 [23] Maledé Demelash Ademe, Agumassie Tena Alamirew, Kosgei Job Rotich, Linh Nguyen Thi
509 Thuy, Andualem Tesfa Gebrie. Analysis of rainfall and streamflow trend and variability over Birr River
510 watershed, Abbay basin, Ethiopia[J]. *Environmental Challenges*,2022,7:100528-100528.
- 511 [24] Mallya G, Tripathi S, Kirshner S, Rao S. Govindaraju. Probabilistic Assessment of Drought
512 Characteristics Using Hidden Markov Model[J]. *Journal of Hydrologic Engineering*,2013,18(7): 834-
513 845.



- 514 [25] Maryam Mokhtarzad, Farzad Eskandari, Nima Jamshidi Vanjani, Alireza Arabasadi. Drought forecasting
515 by ANN, ANFIS, and SVM and comparison of the models[J]. Environmental Earth
516 Sciences, 2017, 76(21): 729-729.
- 517 [26] McKee, T.B., Doesken, N.J., Kleist, J. The relationship of drought frequency and duration to time scales.
518 In: Proc. 8th Conference on Applied Climatology, Anaheim, California, 1993: 179–184.
- 519 [27] Mehdi Rezaeianzadeh, Alfred Stein, Jonathan Peter Cox. Drought Forecasting using Markov Chain
520 Model and Artificial Neural Networks[J]. Water Resources Management, 2016, 30(7): 2245-2259.
- 521 [28] Melanie Oertel, Francisco Javier Meza, Jorge Gironás, Christopher A. Scott, Facundo Rojas, Nicolás
522 Pineda-Pablos. Drought Propagation in Semi-Arid River Basins in Latin America: Lessons from Mexico
523 to the Southern Cone[J]. Water, 2018, 10(11): 1564-1564.
- 524 [29] Miriam Fendeková, Tobias Gauster, Livia Labudová, Dana Vrabliková, Zuzana Danáčová, Marián
525 Fendek, Pavla Pekárová. Analysing 21st century meteorological and hydrological drought events in
526 Slovakia[J]. Journal of Hydrology and Hydromechanics, 2018, 66(4): 393-403.
- 527 [30] Mohammad Mehdi Moghimi, Abdol Rassoul Zarei, Mohammad Reza Mahmoudi. Seasonal drought
528 forecasting in arid regions, using different time series models and RDI index[J]. Journal of Water and
529 Climate Change, 2020, 11(3): 633-654.
- 530 [31] Moreira Elsa E, Mexia João T., Pereira Luís S. Assessing homogeneous regions relative to drought class
531 transitions using an ANOVA-like inference. Application to Alentejo, Portugal[J]. Stochastic
532 Environmental Research and Risk Assessment, 2013, 27(1): 183-193.
- 533 [32] Muhammad Jehanzaib, Sabab Ali Shah, Jiyoung Yoo, Tae-Woong Kim. Investigating the impacts of
534 climate change and human activities on hydrological drought using non-stationary approaches[J].
535 Journal of Hydrology, 2020, 588: 125052-125052.
- 536 [33] Natsagdorj Enkhjargal, Renchin Tsolmon, Maeyer Philippe De, Darkhijav Bayanjargal. Spatial
537 Distribution of Soil Moisture in Mongolia Using SMAP and MODIS Satellite Data: A Time Series
538 Model (2010–2025)[J]. Remote Sensing, 2021, 13(3): 347-347.
- 539 [34] Palmer, W.C., 1965. Meteorological Drought, Research Paper, 45, U.S. Weather Bureau.
- 540 [35] Ren Weinan, Wang Yixuan, Li Jianzhu, Feng Ping, Smith Ronald J. Drought forecasting in Luanhe
541 River basin involving climatic indices[J]. Theoretical and Applied Climatology, 2017, 130(3-4): 1133-
542 1148.



-
- 543 [36] Shradhanand Shukla and Andrew W. Wood. Use of a standardized runoff index for characterizing
544 hydrologic drought[J]. *Geophysical Research Letters*, 2008, 35(2): 2405-1-2405-7.
- 545 [37] Stojković Milan,Plavšić Jasna,Prohaska Stevan,Pavlović Dragutin,Despotović Jovan. A two-stage time
546 series model for monthly hydrological projections under climate change in the Lim River basin
547 (southeast Europe)[J]. *Hydrological Sciences Journal*,2020,65(3): 387-400.
- 548 [38] Wang Menghao,Jiang Shanhu,Ren Liliang,Xu Chong-Yu,Menzel Lucas,Yuan Fei,Xu Qin,Liu Yi,Yang
549 Xiaoli. Separating the effects of climate change and human activities on drought propagation via a
550 natural and human-impacted catchment comparison method[J]. *Journal of*
551 *Hydrology*,2021,603(PA):126913-123613.
- 552 [39] Wang Yixuan, Duan Limin, Liu Tingxi, Li Jianzhu, Feng Ping. A Non-stationary Standardized
553 Streamflow Index for hydrological drought using climate and human-induced indices as covariates[J].
554 *Science of the Total Environment*,2020,699(C):134278-134278.
- 555 [40] Wang Yixuan, Li Jianzhu , Feng Ping, Hu Rong. Analysis of drought characteristics over Luanhe River
556 basin using the joint deficit index[J]. *Journal of Water and Climate Change*,2016,7(2): 340-352.
- 557 [41] Wang Yixuan, Zhang Ting, Chen Xu, Li Jianzhu, Feng Ping. Spatial and temporal characteristics of
558 droughts in Luanhe River basin, China[J]. *Theoretical and Applied Climatology*,2018,131(3): 1369-
559 1385.
- 560 [42] Wang Youxin,Peng Tao,Lin Qingxia,Singh Vijay P.,Dong Xiaohua,Chen Chen,Liu Ji,Chang
561 Wenjuan,Wang Gaoxu. A New Non-stationary Hydrological Drought Index Encompassing Climate
562 Indices and Modified Reservoir Index as Covariates[J]. *Water Resources Management*,2022,36(7):
563 2433-2454.
- 564 [43] Wu Huizhuo. A Sufficient Condition for Determining the Normal Distribution of Two-dimensional
565 Random Variables[J]. *College Mathematics*,2019,35(06): 63-65. (in Chinese)
- 566 [44] Xu Yang, Zhang Xuan, Wang Xiao, Hao Zengchao, Singh Vijay P, Hao Fanghua. Propagation from
567 meteorological drought to hydrological drought under the impact of human activities: A case study in
568 northern China[J]. *Journal of Hydrology*,2019,579(C): 124147-124147.
- 569 [45] Yixuan Wang,Jianzhu Li,Ping Feng,Fulong Chen. Effects of large-scale climate patterns and human
570 activities on hydrological drought: a case study in the Luanhe River basin, China[J]. *Natural*
571 *Hazards*,2015,76(3): 1687-1710.
- 572 [46] Zhang Ting et al. Drought class transition analysis through different models: a case study in North



-
- 573 China[J]. Water Science & Technology: Water Supply, 2017, 17(1): 138-150.
- 574 [47] Aghelpour Pouya and Varshavian Vahid. Forecasting Different Types of Droughts Simultaneously Using
575 Multivariate Standardized Precipitation Index (MSPi), MLP Neural Network, and Imperialistic
576 Competitive Algorithm (ICA)[J]. COMPLEXITY, 2021, 2021
- 577 [48] Majid Dehghani, Bahram Saghafian and Mansoor Zargar. Probabilistic hydrological drought index
578 forecasting based on meteorological drought index using Archimedean copulas[J]. Hydrology Research,
579 2019, 50(5) : 1230-1250.

580 **Acknowledgements: This work was supported by the State Key Laboratory**
581 **of Hydraulic Engineering Simulation and Safety Program (No. HESS-2206**
582 **and No. HESS-2222).**

583