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2 **A nonstationary analysis for investigating the multiscale**
3 **variability of extreme surges: case of the English Channel coasts**
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14 **Abstract**

15 This research examines the nonstationary dynamics of extreme surges along the English
16 Channel coasts and seeks to make their connection to the climate patterns at different time-
17 scales by the use of a detailed spectral analysis in order to gain insights on the physical
18 mechanisms relating the global atmospheric circulation to the local-scale variability of the
19 monthly extreme surges. This variability highlights different oscillatory components from the
20 interannual (~1.5-years, ~2-4-years, ~5-8-years) to the interdecadal (~12-16-years) scales with
21 mean explained variances of ~ 25 - 32 % and ~ 2 - 4 % of the total variability, respectively.
22 Using the two hypotheses that the physical mechanisms of the atmospheric circulation change
23 according to the timescales and their connection with the local variability improves the
24 prediction of the extremes, we have demonstrated statistically significant relationships of ~1.5-
25 years, ~2-4-years, and ~5-8-years and 12-16-years with the different climate oscillations of
26 Sea-Level Pressure, Zonal Wind, North Atlantic Oscillation and Atlantic Multidecadal
27 Oscillation, respectively.

28 Such physical links have been used to implement the parameters of the time-dependent GEV
29 distribution models. The introduced climate information in the GEV parameters has
30 considerably improved the prediction of the different time-scales of surges with an explained
31 variance higher than 60%. This improvement exhibits their nonlinear relationship with the
32 large-scale atmospheric circulation.

33 **Key-Words:** Coastal extreme surges, multi-timescale variability, climate oscillations,
34 nonstationary GEV models

35 **1. Introduction**

36 Risks assessments has been recognized as an urgent task essential to take effective reduction
37 of disasters and adaptation actions of climate change. The increase in coastal flood risk is
38 generally driven by the extreme surges being the result of episodic water fluctuations due to
39 waves and storm surges. High surges are considered as significant hazards for many low-lying
40 coastal communities (e.g. Hanson et al., 2011; Nicholls et al., 2011) and are expected to be
41 intensified with rising global mean sea level (Menendez and Woodworth, 2010).

42 Being an alarming problem for the coastal vulnerability, extreme events have gained the
43 attention of the scientists who have reported the dynamics (e.g. Haigh et al., 2010; Idier et al.,
44 2012; Masina and Lamberti, 2013; Tomasin and Pirazzoli, 2008; Turki et al., 2020) and the
45 projections (e.g. vousdoukas et al., 2017) of extreme surges considering the stationary and the
46 nonstationary contributions from tides, waves, sea-level-rise components (e.g. Brown et al.,

47 2010; Idier et al., 2017), and large-scale climate oscillations (e.g. Colberg et al., 2019; Turki et
48 al., 2019; 2020).

49 Under the assumption of a stationary surges, the concepts of return level and return period
50 provide critical information for infrastructure design, decision-making, and assessing the
51 impacts of rare weather and climatic events (Rosbjerg and Madsen, 1998). However, the
52 frequency of extremes has been changing and is likely to continue changing in the future (e.g.
53 Milly et al., 2008). Therefore, concepts and models that can account for nonstationary analysis
54 of climatic and hydrologic extremes are needed (e.g. Cooley, 2013; Salas and Obeysekera,
55 2013; Parey et al., 2010).

56 Over the last decade, several studies adopted the nonstationary behaviour of extremes to
57 estimate their evolution and their return-periods from rigorous models of Extreme Value
58 Theory (EVT) by incorporating an information related to climate oscillations.

59 In this way, the recurrence of coastal extreme events over the Northern European continent and
60 the persistence of high energetic conditions around the Atlantic have been associated with the
61 deepening of Icelandic Low and the extension/reinforcement of the Azores High. Those facts
62 can be interpreted, at quasi daily timescale, as the preferred excitation of a given atmospheric
63 regime close to the positive phase of the North Atlantic Oscillation. The recent predominance

64 of this regime can be explained partly by the impact of the North Tropical Atlantic Ocean upon
65 the midlatitude atmosphere and by the increase of greenhouse gas concentration induced by
66 human activities.

67 Menendez and Woodworth (2010) have used a nonstationary extreme values analysis together
68 with the NAO (North Atlantic Oscillation) and Arctic Oscillation (AO) indices for improving
69 the estimation of monthly extreme sea-levels along the European coasts.

70 In the Northern Adriatic region, Masina et al. (2013) investigate changes in extreme sea levels
71 applying a nonstationary approach to the monthly maxima and the climate oscillations of NAO
72 and AO (Arctic Oscillation) indices. They have suggested that the increase in the extreme water
73 levels since the 1990s is related to the changes in the wind regime and the intensification of
74 Bora and Sirocco winds after the second half of the 20th century.

75 Then, Marcos et al. (2015) have investigated the decadal and multidecadal changes in sea level
76 extremes using long tide gauge records distributed worldwide. They have demonstrated that
77 the intensity and the occurrence of the extreme sea levels vary on decadal scales in the most of
78 the sites in relation with a common large-scale forcing. In the same way, the study of extreme
79 sea levels along the coastal zones of the North Atlantic Ocean and the Gulf of Mexico has
80 shown that the mean sea level should be considered as the major driver of extremes (Marcos

81 and Woodworth 2017) since the intensity of extreme episodes increases at centennial time
82 scales, together with multidecadal variability. The extreme sea levels along the United States
83 coastline between 1929 and 2013 have been investigated by Wahls and Chambers (2015;
84 2016). Wahls and Chambers (2015) have identified the relation between the multidecadal
85 variations in extreme sea and the changes in mean sea level. Such relation has been mainly
86 pointed toward some regions where storm surges are primarily driven by extratropical cyclones
87 and should contribute in the variation of relevant return water levels required for coastal design.
88 Such extremes have been then investigated in Wahls and Chambers (2016) works aiming to
89 define their relationship with the large-scale climate variability by the use of simple and
90 multiple linear regression models.

91 In the English Channel, the extreme sea levels have been addressed by several works (e.g.
92 Haigh et al., 2010; Idier et al., 2012; Tomasin and Pirazzoli, 2008; Turki et al., 2015a; Turki et
93 al., 2019) with the aim of investigating their dynamics at different timescales and their
94 connections to the atmospheric circulation patterns.

95 Haigh et al. (2010) investigated the interannual and the interdecadal extreme surges in the
96 English Channel and their strong relationship with the NAO index. Their results showed weak
97 negative correlations throughout the Channel and strong positive correlations at the boundary

98 along the Southern North Sea. Using a numerical approach, Idier et al. (2012) studied the spatial
99 evolution of some historical storms in the Atlantic Sea and their dependence on tides.

100 Recently, Turki et al. (2019) have examined the multiscale variability of the sea-level changes
101 in the Seine bay (NW France) in relation with the global climate oscillations from the SLP
102 composites; they have demonstrated dipolar patterns of high-low pressures suggesting positive
103 and negative anomalies at the interdecadal and the interannual scales respectively.

104 Despite these important advances, no particular studies exist on sea-level dynamics and
105 extreme events linked to the large-scale climate oscillations along the English Channel
106 coastlines. The aforementioned works of Turki et al. (2015a, 2019) have focused on the
107 multiscale sea-level variability along the French coasts related to the NAO and the Sea-Level
108 Pressure (SLP) patterns; however, they have not addressed the regional behaviour of the
109 extreme sea levels in relation with the global climate oscillations.

110 Then, similar approaches have been used by Turki et al. (2020) to quantify the nonstationary
111 behaviour of extreme surges and their relationship with the global atmospheric circulation at
112 different timescales along the English Channel coasts (NW France) between 1964 and 2012.
113 They have reported that the intermonthly and the interannual variability of monthly extrema

114 are statistically modelled by nonstationary GEV distribution using the full information related
115 to the climate teleconnections.

116 In the same context, the present contribution aims to investigate the interannual and the
117 interdecadal dynamics of extreme surges along the English Channel coasts (NW France and
118 SW England) by the use of combining techniques of spectral analyses and probabilistic models.

119 We hypothesize that different large-scale climate variables may be involved in explaining the
120 occurrence of extreme surges, and that this dependence can be a function of each timescale.

121 The rationale behind this hypothesis is based on the following: (1) each timeseries of extreme
122 surges should depend on different timescales; (2) each timescale should be related to a specific
123 large-scale oscillation. Using this hypothesis, the linkages between the local extreme surges
124 and the large-scale climate oscillations are deciphered with the aim to improve the extreme
125 models using the most consistent large-scale oscillations as covariates.

126

127 The overall approach for testing our hypotheses can be described as follows, for a given
128 extreme surge timeseries: i) identify the short to long timescale oscillations characterizing the
129 local variability of the extreme surges; ii) explore the correlation between the local extreme
130 surges and the selected large-scale variable from short to long timescales; iii) select the most

131 appropriate large-scale variable as an explanatory parameter to be used as a covariate in
132 nonstationary GEV models and estimate the extreme surges.

133 The paper is structured as follows. The used hydro-climatic data are presented in section 2,
134 including local extreme surges and large-scale variables. Section 3 explains the methodological
135 approach used. Finally, the sections 4 and 5 report the results related to the multiscale
136 variability of extreme surges along the English Channel and their teleconnections with the
137 large-scale climate oscillations required for their estimation by the use of GEV extreme models.
138 The concluding remarks of these findings are addressed in section 6.

139
140 **2. Database description**

141 The present research focuses on the dynamics of extreme surges along the English Channel
142 coasts (French and the Britannic coasts); It has been conducted in the framework of some
143 French research programs: RICOCHET (ANR program), RAIV COT (Normandy Region
144 program) and the international project COTEST (CNES-TOSA program) related to the future
145 mission Surface Water and Ocean Topography (SWOT).

146 The English Channel (Figure 1) is a shallow sea between Northern France and South England,
147 connecting Atlantic Ocean to North Sea. Melting of retreating glaciers formed a megaflood in
148 the southern North Sea and it geographically separated Britain from Europe and formed English

149 Channel at the last Quaternary Period (Collier et al. 2015). English Channel has a complex sea
150 floor due to its characteristics of formation. It is deep and wide on the western side, narrower
151 and shallower towards Strait of Dover. Largest width of the Channel is around 160 km (Figure
152 1). The average depth of the channel is about 120 m. It gradually narrows eastward to a width
153 of 35 km and depth of around 45 m in the Dover Strait. The east to west extent of the Channel
154 is about 500 km. The overall width of shallow depths is wider in the French side of the channel.
155 The extreme storm surges of this area are mostly occurred by low pressure systems from the
156 Atlantic Ocean, propagating eastwards or storm surges propagating south from the North Sea
157 (Law, 1975). The area is exposed to major storms from Atlantic side of the channel, having a
158 maximum fetch of winds, from west to southeast then to northwest.

159 Three tide gauge sites along the French coasts have been used in the present study: (1) Dunkirk
160 station which is a few kilometres away from Belgian borders, (2) Cherbourg station located on
161 the Cotentin Peninsula and at the opening of the Atlantic Sea, (3) Brest station which is a
162 sheltered bay located at the western extremity of metropolitan France and connected to the
163 Atlantic Ocean.

164 *Figure 1. Geographical location of the study area and the different tide gauges along the*
165 *English Channel coasts: Brest, Cherbourg, Dunkirk (NW France); Dover and Weymouth (SW UK).*

166 Two tide gauge sites along the Britannic coasts have been used: (1) Dover station which is
167 separated from Dunkirk by the North Sea and (2) Weymouth station symmetrically with
168 respect to Cherbourg.

169 The French tide gauges are operated and maintained by the National French Center of
170 Oceanographic Data (SHOM) while the Britannic tide gauges are operated by the British
171 Oceanographic Data Center. All stations are referenced to the hydrographic zero level; they
172 provide time-series of hourly observations measurements until 2018.

173 Available data are summarized as the following: Brest (168 years between 1850 and 2018);
174 Cherbourg and Dunkirk (54 years between 1964 and 2018); Dover (53 years between 1963 and
175 2018); Weymouth (28 years between 1990 and 2018).

176 The hourly measurements suffer from some gaps of daily length distributed along the time-
177 series. These gaps have been processed by the hybrid model for filling gaps developed by Turki
178 et al. (2015b) by using the SLP as covariate in ARMA methods and the memory effects of the
179 previous distribution of surges to estimate the missing values and fill the gaps. This model has
180 been used in the recent works of Turki et al., (2019; 2020).

181 The large-scale atmospheric circulations are represented in this work by four different climate
182 indices which are considered as fundamental drivers in the Atlantic regions (Massei et al.,

183 2017; Turki et al., 2019; 2020): the Atlantic Multidecadal Oscillation (AMO), the North
184 Atlantic Oscillation (NAO), the Zonal Wind (ZW) component extracted at 850hPa, and the
185 Sea-Level Pressure (SLP).
186 Monthly time-series of climate indices have been provided by the NCEP-NCAR Reanalysis
187 fields (<http://www.esrl.noaa.gov/psd/data/gridded/data.ncep.reanalysis.derived.html>) until
188 2017. The different indices have been extracted during the same period of the sea-level
189 observations at the four stations Cherbourg, Dunkirk, Dover and Weymouth. For the longest
190 timeseries of Brest (1850 - 2018), the use of climate indices has been limited according to their
191 initial date availability (AMO: 1880 – 2017; NAO: 1865-2017; SLP: 1948-2017; ZW: 1865-
192 2017).

193 **3. Methodological Approach**

194 ***3.1 Extraction of residual sea level: ‘surges’***

195
196 The total sea-level height, resulting from the astronomical and the meteorological processes,
197 exhibits a temporal non-stationarity which is explained by a combination of the effects of the
198 long-term trends in the mean sea level, the modulation by the deterministic tidal component
199 and the stochastic signal of surges, and the interactions between tides and surges. The
200 occurrence of extreme sea levels is controlled by periods of high astronomically generated
201 tides, in particular at inter-annual scales when two phenomena of precession cause systematic
202 variation of high tides. The modulation of the tides contributes to the enhanced risk of coastal
203 flooding. Therefore, the separation between tidal and non-tidal signals is an important task in

204 any analysis of sea-level time-series. By the hypothesis of independence between the
205 astronomical tides and the stochastic residual of surges, the nonlinear relationship between the
206 tidal modulation and surges is not considered in the present analysis. Using the classical
207 harmonic analysis, the tidal component has been modelled as the sum of a finite set of sinusoids
208 at specific frequencies to determine the determinist phase/ amplitude of each sinusoid and
209 predict the astronomical component of tides. In order to obtain a quantitative assessment of the
210 non-tidal contribution in storminess changes, technical methods based on MATLAB t-tide
211 package have been applied to the seal level measurements, demodulated from long-term
212 components (e.g. mean sea level, vertical local movement), for estimating year-by-year tidal
213 constituents. A year-by-year tidal simulation (Shaw and Tsimplis, 2010) has been applied to
214 the sea-level time-series to determine the amplitude and the phase of tidal modulations using
215 harmonic analysis fitted to 18.61-, 9.305-, 8.85-, and 4.425-year sinusoidal signals (Pugh,
216 1987). The radiational components have been also considered for the extraction of the
217 stochastic component of surges (Williams et al., 2018).

218 ***3.2 Wavelet spectral analysis***

219 The Continuous Wavelet Transform CWT is generally used for data analysis in hydrology,
220 geophysics, and environmental sciences (Labat, 2005; Sang, 2013; Torrence and Compo,
221 1998). This technique produces the timescale with the means of the Fourier transform contour
222 diagram on which the time is indicated on the x-axis, the timescale (period,) on the y-axis, and
223 the variance (power) on the z-axis.

224 Then, a wavelet multiresolution analysis has been used to decompose the signal of monthly
225 extreme surges into different internal components corresponding to different timescales. This

226 decomposition consists on applying a series of iterative filtering to the signal by the use of low-
227 pass and high-pass filters able to produce the spectral components describing the total signal.
228 More details are presented in the recent works of Massei et al. (2017) and Turki et al., (2019).
229 In summary, the total signal has been separated into a relatively small number of wavelet
230 components from high to low frequencies that altogether explains the variability of the signal;
231 this will be illustrated later using the hourly measurements and the monthly maxima of surges.
232 The wavelet coherence has been calculated to investigate the relationship between the extreme
233 surges and the climate oscillations by identifying the timescales where the two timeseries co-
234 vary, even if they do not display high power. Here, a significance test has been implemented
235 by the use of a Monte Carlo analysis based on an autocorrelation function of two timeseries
236 (Grinsted et al., 2004).

237 ***3. 3 Nonstationary extreme value model***

238 Finally, and with the aim of addressing the nonstationary behaviour of extreme surges, the
239 monthly maxima of the surges have been calculated and decomposed with the multiresolution
240 analysis. Then, a nonstationary extreme value analysis based on the GEV distribution with
241 time-dependent parameters (Coles, 2001) has been implemented to model the series of the
242 monthly maxima surges. There are several GEV families which depend on the shape parameter,
243 e.g. Weibull ($\varepsilon < 0$), Gumbel ($\varepsilon = 0$), and Fréchet ($\varepsilon > 0$). The three parameters of the GEV (i.e.
244 location μ , scale ψ , shape ε) are estimated by the maximum likelihood function.

245 The nonstationary effect was considered by incorporating the selected climate indices (NAO,
246 AMO, ZW, and SLP) into the parametrization of the GEV models. Akaike Information

247 Criterion (AIC) has been used to select the most appropriate probability function models. The
248 methods of maximum likelihood were used for the estimation of the distribution's parameters.
249 The approach used considers the location (μ), the scale (ψ), and the shape (ε) parameters with
250 relevant covariates, which are described by a selected climate index:

$$251 \quad \mu(t) = \beta_{0,\mu} + \beta_{1,\mu}Y_1 + \dots + \beta_{n,\mu}Y_n \quad (1)$$

$$252 \quad \psi(t) = \beta_{0,\psi} + \beta_{1,\psi}Y_1 + \dots + \beta_{n,\psi}Y_n \quad (2)$$

$$253 \quad \varepsilon(t) = \beta_{0,\varepsilon} + \beta_{1,\varepsilon}Y_1 + \dots + \beta_{n,\varepsilon}Y_n \quad (3)$$

254 Where $\beta_0, \beta_1, \dots, \beta_n$ are the coefficients, and Y_i is the covariate represented by the climate
255 index. For each spectral component, only one climate index can be used to be introduced into
256 the parameters μ, ψ , and ε of the nonstationary GEV model (into one of them, into two of them
257 or into the three parameters). With the aim of optimizing the best use of the most appropriate
258 climate index (detailed in section 3.4) into the different GEV parameters, a series of sensitivity
259 analyses were implemented for each timescale. The AIC measures the goodness of the fitting
260 of the model (Akaike, 1973) to the relation $AIC = -2l + 2K$; where l is the log-likelihood value
261 estimated for the fitted model, and K is the number of the model parameters. Higher ranked
262 models should result from lower AIC scores.

263 ***3. 4 Determination of the most appropriate climate oscillation***
264 ***connected to each timescale extreme surges for GEV models***

265

266 As suggested previously, the main hypothesis presented in this research is that effects of the
267 physical mechanisms on the extreme surges vary according to the timescale and each scale
268 should be related to a given climate oscillation.

269 This hypothesis has been investigated by two approaches:

270 (1) a spectral approach based on the use of wavelet techniques (wavelet multiresolution and
271 wavelet coherence as detailed in section 3.2) for optimizing the physical relationship between
272 the climate index and the extreme surges at each timescale.

273 The Bootstrap is a resampling technique used to estimate the sampling distribution of an
274 estimator of sample statistics by drawing randomly with replacement from a set of data points.

275 Here, a bootstrap approach has been applied to assess the statistical significance of the
276 correlation between the spectral component of the extreme surges and the climate oscillation
277 at each timescale. By resampling the timeseries 10.000 times, the extreme surges have been
278 simulated and compared to the original records; the 95% confidence intervals have been
279 considered to extract the best climate information fitting the extreme surges (Villarini et al.,
280 2009).

281 (2) a Bayesian estimation has been used to make inferences from the Likelihood function. The
282 reason behind the choice of this approach is overcoming the limitation of short time-series with
283 small size, the case of Weymouth station where the measurements covers the period from 1991
284 and 2018. A technique of Markov Chain Monte Carlo (MCMC), implemented in the evbayes
285 package within R software, has been used basing on multiple simulations (the number of
286 simulations is varying as a function of the length of the timeseries).

287 For each spectral component, a sample of 100.000 simulations has been modelled by GEV
288 using a given climate index. The upper and lower quantiles of the posterior probability
289 distribution for the parameters of the MCMC sample are taken. The goodness of fit has been

290 taken as a function of the values of the upper and the lower quantiles; best results have been
291 considered when these values are higher than 92.5% and lower than 5.2%, respectively.

292 Both approaches have been used to select the most appropriate atmospheric physical
293 mechanism to each timescale of extreme surges. Then and at each timescale (i.e. spectral
294 component), the selected mechanism (i.e. the climate oscillation in this case) has been used as
295 covariate for modelling the extreme surge by nonstationary the implemented GEV models.
296 The best use of the covariate into the different GEV parameters (location, scale and shape)
297 have been investigated by means of AIC criterion.

298 Once the best GEV models defined for each time scale, a series of simulations have been
299 carried out to compare modelled and observed surges.

300 This Bayesian inference has been also used to calculate: (1) the return levels of the
301 nonstationary simulated surges which were compared to those of the observed ones; (2) the
302 confidence interval (CI) assessing the goodness of this comparison.

303 Figure 2 summarizes the methodological approach proposed in the present research and the
304 different steps implemented. The statistical methods used to resolve each step are also
305 synthetized in this figure.

306 *Figure 2. A summary of the methodological approach implemented in this research.*

307 **4. Multi-timescale variability of extreme surges**

308 The variability of the monthly extreme surges along the English Channel coasts has been
309 investigated using the continuous wavelet transform (CWT). In the spectrum of Figure 3, the
310 colour scale represents an increasing power (variance) from red to blue and pink. The CWT
311 diagrams highlight the existence of several scales for all sites with different ranges of

312 frequencies: the interannual scales of ~ 1.5-yr, ~ 2-4-yr, ~ 5-8-yr and the interdecadal scale of
313 ~ 12-16-yr.

314 The variability of surges is clearly dominated by the interannual frequencies (~ 1.5-yr, ~ 2-4-
315 yr, ~ 5-8-yr) explaining a mean variance between 32% and 25% of the total energy (Table 1).

316 In Dover and Weymouth, the low frequencies of ~ 2-4-yr are well-structured with a mean
317 explained variance of 9.5% while it is of 7% for ~ 5-8-yr. These percentages decrease slightly
318 for the French sites to 8% and 5%, respectively. At ~ 1.5-yr, the explained variance is higher
319 than 16% and 13% respectively in Britannic and French coasts. The interdecadal frequency of
320 ~ 12-16-yr varies between 2% and 4% from the total signal. This frequency is not observed in
321 the shortest timeseries of Weymouth (Table 1).

322 The interannual variability (time-scales higher than ~1 year) seems to be highly represented in
323 the monthly extrema CWT (Figure 3). It's not the case for the monthly mean surges (Figure
324 4.a) where most power spectrum is concentrated on the annual cycle with an explained variance
325 higher than 50%.

326 The time-dependent PDF of the monthly mean and maximum surges over a period of 10 years,
327 for illustration purpose, is displayed in Figure 4.b. The ~1-yr component of monthly mean
328 surges is largely manifested with a pronounced variation of the Gaussian curves in time; such

329 variations take wavelengths of approximately ~ 2 -yr and ~ 4 -yr. This result exhibits that the
330 interannual frequencies of ~ 2 -yr and ~ 4 -yr are modulated within the annual mode for the
331 mean surges while they are implicitly quantified for the monthly maxima.

332 *Figure 3. CWT of monthly maxima of surges in Brest, Cherbourg, Dunkirk, Dover and*
333 *Weymouth.*

334 *Figure 4. Multiscale variability of the monthly mean and maximum surges in Brest. (a) CWT*
335 *of monthly mean surges; (b) Interannual variability of monthly and extreme surges*

336

337 Results have been explored to investigate the nonstationary dynamics of surges at different
338 timescales. We have applied the wavelet multiresolution decomposition of monthly extrema
339 for each site. The process has resulted in the separation of several components with different
340 time-scales. Only the wavelet components, with have been considered in this work. In this
341 research, we are interested in the time-scales higher than 1 year, i.e. traduced by three
342 interannual scales (~ 1.5 -yr, ~ 2 -4-yr, and ~ 5 -8-yr) and an interdecadal scale of ~ 12 -16-yr.
343 We focused only on the interannual and the interdecadal scales whose fluctuations correspond
344 to the oscillation periods less than half the length of the record and exhibit a high-energy
345 contribution on the variance of the total signal. The lowest frequency, corresponding to ~ 12 -
346 16-yr is easily calculated from the longest record of Brest.

347 Figure 5 shows a series of oscillatory components of surges from interannual to interdecadal
348 scales, not easily quantified by a simple visual inspection of the signal. High similarities
349 between the different sites have been highly observed for the interannual and the interdecadal
350 scales of ~ 5-8-yr and ~12-16-yr while they are less pronounced at the small scales of ~ 1.5-yr.
351 At this timescale, the differences in the extreme surges can be explained by local physical
352 phenomena controlling their dynamics. Such processes are mainly induced by combining the
353 effects of meteorological and oceanographic forces including changes in atmospheric pressures
354 and wind velocities in shallow water areas. Beyond ~ 1.5-yr, the variability of extreme surges
355 at larger scales seems to be quite similar in terms of frequency and amplitude for the five sites.
356 Such large variability reveals the physical effects of a global contribution related to climate
357 oscillations. The extent of the large-scale oscillations is not strictly similar and changes
358 according to the timescale variability since the dynamics of surges is not necessarily related to
359 the same type of atmospheric circulation process. This relationship will be addressed later in
360 the second part of this section.

361 Here, the multiscale variability of extremes has been investigated from the spectral components
362 of surges along the English Channel coasts. This signal has been linearly extracted from the
363 total sea level, provided by tide gauges, by the use of the classical harmonic analysis and thanks

364 to the assumption that the water level is the sum of the mean sea level, tides, and surges. This
365 assumption approximates the quantification of both components in the English Channel where
366 the significant tide-surge interactions (Tomassin & Pirazzoli, 2008) and the effects of the sea-
367 level rise on tides and surges are important (e.g. Idier et al., 2017). Neglecting this nonlinear
368 interaction between the surges, tides, and the sea-level rise suggests some uncertainties in the
369 estimation of the high frequencies of the spectral components between daily and monthly
370 scales, which is not the focus of the present work where the interannual and the interdecadal
371 scales are investigated.

372 Similar interannual timescales have been observed along the French coasts of Dunkirk, Le
373 Havre and Cherbourg in Turki et al., (2020) works where the intermonthly and the interannual
374 variability of 48-year hourly surges has been investigated. They have demonstrated that the
375 timescales smaller than ~ 1.5 -yr are differently manifested between the different sites. These
376 differences have been associated to the local variability of surges induced by combining the
377 effects of meteorological and oceanographic forces including changes in atmospheric pressures
378 and wind velocities in shallow water areas. As demonstrated in Turki et al. (2020) works, the
379 mean explained variance of the interannual fluctuations (~ 1.5 -yr, ~ 2 -4-yr, and ~ 5 -8-yr) is
380 around 25% of the total surges along the French coasts (Table 1). This value is higher than 32%

381 in Weymouth and Dover while the explained variance of the interdecadal scales (~ 12-16-yr)
382 is also more important with 3.5% (compared to 2% for the French coasts).

383 The interdecadal variability (~ 12-16-yr) of extreme surges have been evidenced by Turki et
384 al., (2019) in the Seine bay (NW France). Strong physical relations have been exhibited
385 between the interdecadal time of ~ 12-16-yr and the exceptional stormy events produced with
386 surges higher than 10-year return period level. The connections between the low-frequency
387 components and the historical record of the exceptional events suggested that storms would
388 occur differently according to a series of physical processes oscillating at multi-timescales;
389 these processes control their frequency and their intensity (Turki et al., 2019).

390 *Table 1. The explained variance expressed as percentage of total variance of monthly extreme surges*
391 *The lowest frequency of ~ 12-16-yr is not observed in Weymouth. Figure 4 Wavelet details*
392 *(components) resulting from the multiresolution analysis of surges at the interannual (~ 1.5-*
393 *yr , ~2-4-yr and ~5-8-yr) and interdecadal (~12-16-yr) time scales for all sites (Brest,*
394 *Cherbourg, Dunkirk, Dover and Weymouth).*

395 Accordingly, the multiscale variability of extreme surges exhibits a nonstationary behaviour
396 modulated by a non-linear interaction between the different interannual and the interdecadal
397 timescales. Then, assessing the effect of the nonstationary behaviour at different timescales is
398 important for improving the estimation of extreme values and the projection of storm surges.
399

400 **5. Large-scale climate North-Atlantic oscillations and their link to** 401 **extreme surges in the English Channel**

402

403 In this part, a new hybrid approach combining the spectral analysis and the nonstationary GEV
404 models has been used to investigate the connection between the multi-timescale variability of
405 local surges and the large-scale climate North Atlantic oscillations.

406 As proposed by Turki et al. (2019; 2020), the hypothesis used in the present work is that the
407 multi-timescale variability of the local extreme surges should be strongly related to different
408 climate teleconnections induced by a complex contribution of many physical mechanisms. This
409 non-linear relationship varies according to each timescale which depends on a specific large-
410 scale oscillation of atmospheric circulation.

411 ***5.1 To what extent would large-scale climate oscillations link extreme surges?***

412 The wavelet coherence (WC) diagrams between the monthly maxima of surges and the
413 different climate indices of SLP, ZW, NAO, AMO, introduced previously as the main
414 atmospheric circulation within the English Channel, are illustrated respectively in Figures 6, 7,
415 8 and 9. Results provided by these diagrams highlight:

416 1. The connection between the climate oscillations and the extreme surges is manifested
417 differently as a function of the timescale. From a visual inspection of the different spectra

418 provided WC, the most significant correlations of extreme surges have been identified with
419 SLP, ZW, NAO and AMO respectively at ~ 1.5-yr, ~ 2-4-yr, ~ 5-8-yr and ~ 12-16-yr.

420 2. Each timescale exhibits mainly strong links with its associated climate index (explained
421 variance varying between 55% and 80%) and weak ones with other indices (explained variance
422 varying between 15% and 5%). Table 2 summarizes the contribution of the different climate
423 oscillations in the different interannual and interdecadal timescales of extreme surges. Here,
424 mean values between the different sites are presented.

425 For example, SLP diagrams reveal significant relationships with ~ 1.5-yr surges (well-
426 structured forms with high concentration of pink-blue colours in Figure 6); limited correlations,
427 locally positioned in time, have been observed at ~ 2-4-yr and ~ 5-8-yr scales. ZW shows strong
428 correlations with interannual surges at ~ 2-4-yr (blue to pink colour at this scale; Figure 7) and
429 others correlations at smaller and larger timescales of ~ 1.5-yr and ~ 5-8-yr, respectively.

430 Similarly, NAO presents high links with ~ 5-8-yr surges and small relations with ~ 2-4-yr and
431 ~ 1.5-yr (Figure 8).

432 The ~ 1.5-yr scale highlights strong correlations with SLP with an explained variance of 75%;
433 25% of this scale should be explained by the influence of other climate oscillations (basically

434 ZW and NAO with a mean explained variance of 10% and 6%, respectively; Table 2) and the
435 combining effects of local driven forcing induced by winds and waves.
436 65% of ~ 2-4-yr scale is correlated with ZW while 5% and 12% is explained by the effect of
437 NAO and SLP, respectively. The effects of NAO on the ~ 5-8-yr vary between 55% and 65%;
438 minor influence at this scale has been observed with SLP and ZW explaining a mean variance
439 of 13%. The interannual scales of surges are slightly influenced by AMO oscillations with low
440 values of variance lower than 1% (Table 2).

441 *Table 2. The mean explained variance expressed as percentage of total variance provided by the*
442 *wavelet coherence between the extreme surges and the climate Oscillations (SLP, ZW, NAO, AMO)*
443

444 *Figure 6. Cross-wavelet correlations between monthly extrema of surges and Sea Level*
445 *Pressure (SLP).*

446 *Figure 7. Cross-wavelet correlations between monthly extrema of surges and Zonal Wind*
447 *(ZW).*

448
449 *Figure 8. Cross-wavelet correlations between monthly extrema of surges and North Atlantic*
450 *Oscillation (NAO).*

451 *Figure 9. Cross-wavelet correlations between monthly extrema of surges and Atlantic*
452 *Multidecadal Oscillation (AMO).*

453

454

455

456 At interdecadal scales of ~ 12 - 16 -yr, the extreme surges are mainly controlled by the AMO
457 oscillations with a mean explained variance of 80% while the effects of NAO are limited to
458 10% (Figure 9).

459 Figure 10 displays the spectral components of the four climate oscillations, provided by a multi-
460 resolution analysis, together with the spectral components extracted from the extreme surges
461 (Figure 5) with the aim to quantify the different connections between both variables at the
462 interannual and the interdecadal timescales.

463 For each timescale, a bootstrap approach has been applied to assess the statistical significance
464 of the correlation between the spectral component of the extreme surges and the climate
465 oscillation. By resampling the timeseries 10.000 times, 95% confidence intervals (CI) have
466 been considered to extract the best climate information fitting the extreme surges (Villarini et
467 al., 2009).

468 Results of 95% CI, applied to the maximum values and provided by the bootstrap technique,
469 for the different possible connections between a given scale of extreme surges and a climate
470 oscillation, extracted at the scale timescale of extreme surges, are illustrated in Table 3. For
471 each connection, the lower and the upper bounds of the 95% CI, calculated for the maximum
472 values, are calculated (values in the square brackets in Table 3). The width of CI between both
473 bounds reveals the statistical significance of the simulations and the goodness of the
474 correlations. According to these results, the width of CI is important for the small timescales
475 (high frequencies of ~ 1.5 -yr and ~ 2 - 4 yr) with a mean of 0.05 and a maximum of 0.09. It
476 decreases for the large timescales (low frequencies of ~ 5 - 8 -yr and ~ 12 - 16 -yr) to 0.012 and
477 0.022 of mean and maximum values, respectively.

478 The smallest widths are observed for the connection between ~ 1.5 -yr extreme surges and SLP;
479 ~ 2-4 -yr extreme surges and ZW; ~ 5-8 -yr extreme surges and NAO; ~ 12-16 -yr extreme
480 surges and AMO (grey columns in Table 3). For the large timescales (low frequencies), small
481 widths (lower than the mean value 0.012) are also observed with other climate oscillations,
482 different from the most appropriate one showing the most significant correlation with extreme
483 surges. For these scales, the variability of extreme surges is mainly controlled by a given
484 physical mechanism, related a climate oscillation, and partly affected by the contribution of
485 other oscillations which could interact with the different large driven forces. Such interaction
486 between the different climate oscillations has been also pointed from Table 2.

487 The best correlation of each extreme surge component (i.e. ~ 1.5-yr, ~ 2-4-yr, ~ 5-8-yr and ~
488 12-16-yr) with the most appropriate climate oscillation (i.e. SLP, ZW, NAO and AMO) is
489 illustrated in Figure 10. The interannual and the interdecadal variability of extreme surges and
490 their multiscale connection with the climate oscillations highlight the nonlinear relationship
491 between large- and local- scales.

492 *Table 3 Analysis of the statistical significance of the correlation between the spectral*
493 *component of the extreme surges and the climate oscillation at each timescale for the*
494 *different stations. The 95% Confidence Intervals (applied for the maximum values) from*
495 *Bootstrap technique in Square Brackets. The most significant correlations, shown by the most*
496 *limited (smaller) 95% CI between the lower and the upper bounds, are illustrated by the grey*
497 *columns. The lowest frequency of ~ 12-16-yr is not observed in Weymouth.*

498 *Figure 10 Wavelet details of monthly extreme surges (black lines), at the interannual (~ 1.5-*
499 *yr , ~2-4-yr and ~5-8-yr) and interdecadal (~12-16-yr) time scales for all sites (Brest,*
500 *Cherbourg, Dunkirk, Dover and Weymouth), correlated to the spectral component of climate*
501 *oscillations associated to the different indices SLP, ZW, NAO and AMO (grey line). Only the*
502 *connection maximizing the correlation coefficient between a selected climate index and the*
503 *component of surges (from interannual to the interdecadal timescales) is presented (the*
504 *normalized values have been calculated to superpose both signals).*

505 Therefore, the interannual and the interdecadal extreme surges have proven to be strongly
506 related to different composites of oscillating atmospheric patterns. Such composites seem to be
507 not necessarily similar for the different timescales. The use of a multiresolution approach to
508 investigate the dynamics of the extreme surges into the downscaling studies proves to be useful
509 for assessing the nonstationary dynamics of the local extreme surges and their nonlinear
510 interactions with the large-scale physical mechanisms related to climate oscillations.

511 Investigating the complex relationships between the climate oscillations and the multi-
512 timescale surges has exhibited a multimodel climate ensemble that should be used to better
513 understand this complexity.

514 The interannual connections between the local hydrodynamics and the climate variability have
515 been investigated in numerous previous works focused on the atmospheric circulation with
516 different related mechanisms (e.g., Feliks et al., 2011; Lopez-Parages et al. 2012; Zampieri et
517 al., 2017). As demonstrated by the recent works of Turki et al. (2019, 2020), the effects of SLP
518 oscillations on the ~ 1.5 -yr variability of extreme surges are described by dipolar patterns of
519 high-low pressures with a series of anomalies which are probably induced by some physical
520 mechanisms linked to the North-Atlantic and ocean/atmospheric circulation oscillating at the
521 same timescale.

522 The SLP fields combined with the baroclinic instability of wind stress have been related to the
523 Gulf Stream path as given by NCEP reanalysis (Frankignoul et al., 2011); the dominant signal
524 is a northward (southward) displacement of the Gulf Stream when the NAO reaches positive
525 (negative) extrema. Daily mean SLP fields have been used by Zampieri et al. (2017) to analyse
526 the influence of the Atlantic sea temperature variability on the day-by-day sequence of large-
527 scale atmospheric circulation patterns over the Euro-Atlantic region. They have associated the
528 significant changes in certain weather regime frequencies to the phase shifts of the AMO. For
529 hydrological applications, several works have investigated the multiscale relationships between
530 the local hydrological changes and the climate variability. Lavers et al. (2010) associated the
531 7.2-yr timescales to SLP patterns which are not exactly reminiscent of the NAO and define
532 centers of action which are shifted to the North.

533 Regarding the ZW (u850), results have shown its correlation with the interannual scales of ~2-
534 4-yr extreme surges as suggested also in the recent findings of Turki et al. (2020). Its influence
535 has been proven also at smaller (~1.5-yr) and larger scales (5-8-yr). Additionally, to extreme
536 surges, the interaction between the ZW and the temperature at different timescales has been
537 highlighted in some previous researches (e.g., Andrade et al., 2012; Seager et al. 2003;
538 Woodworth et al., 2007). Along UK and Northern English Channel coasts, Changes in trends

539 of extreme waters and storm surges have been explained by variations of energy pressure and
540 ZW variability additional to thermosteric fluctuations linked to NAO (Woodworth et al., 2007).
541 Andrade et al. (2012) have used the component of ZW at 850 hPa to investigate the positive
542 and the negative phases of the extreme temperatures in Europe and their occurrence in relation
543 with the large-scale atmospheric circulation. They suggested that both phases are commonly
544 connected to strong large-scale changes in zonal and meridional transports of heat and
545 moisture, resulting in changes in the temperature patterns over western and central Europe
546 (Corte-Real et al., 1995; Trigo et al., 2002). The physical connections between ZW and the
547 extreme events from 11 Global Climate Model runs have been demonstrated by the studies
548 from Mizuta (2012) and Zappa et al (2013); they have suggested the complex relationship
549 between the climate oscillation and the jet stream activity. They have found a slight increase
550 in the frequency and strength of the storms over the central Europe and decreases in the number
551 of the storms over the Norwegian and Mediterranean seas.

552 The NAO is considered as an influencing climate driver for the large-scale atmospheric
553 circulation, as suggested by other researches (e.g. Marcos et al., 2009; Philips et al., 2013).
554 The existence of long-term oscillations originating from large-scale climate variability and thus
555 controlling the interannual extreme surges has been highlighted from investigating the low

556 frequencies of the sea levels along the English Channel. This is in agreements with the results
557 recently demonstrated by Turki et al. (2020) and the present finding exhibiting the strong links
558 between NAO oscillations and the ~ 5-8-yr extreme surges along the English Channel coasts.
559 The physical mechanisms related to the effects of the continuous changes in NAO patterns on
560 the sea-level variability have been addressed in several studies (e.g., Marcos et al., 2005;
561 Tsimplis et al., 1994). At the interannual scales, the key role of NAO on the sea-level variability
562 has been explained by some previous works: Philips et al. (2013) investigated the influence of
563 the NAO on the mean and the maximum extreme sea levels in the Bristol Channel/Severn
564 Estuary. They have demonstrated that when high NAO winters increase in the positive phase,
565 wind speeds also escalate while increasing the negative NAO winters results in low wind
566 speeds. Then, the correlation between the low/high extreme surges and the NAO in the Atlantic
567 has demonstrated a proportionality between NAO values and the augmentation in the winter
568 storms. Feliks et al. (2011) defined significant oscillatory timescales of ~ 2.8-yr, ~ 4.2-yr, and
569 ~ 5.8-yr from both observed NAO index and NAO atmospheric marine boundary layer
570 simulations forced with SST; they have suggested that the atmospheric oscillatory modes
571 should be induced by the Gulf Stream oceanic front.

572 Strong correlations between the monthly extreme surges and the AMO oscillations have been
573 identified at the timescale of 12-16-yr (Figure 8 and Figure 9; in particular for Brest). Since the
574 period of 1990's, the AMO and the extreme surges oscillate in opposition of phase. This shift
575 should be explained by a substantial change in European climate manifested by cold wet and
576 hot dry summers in the northern and the southern Europe, respectively; as discussed by Sutton
577 and Dong (2012). They have demonstrated that the patterns, identified from the European
578 climate change around 1990's, are synchronised with changes related to the North Atlantic
579 Ocean.

580 Other weak links with the AMO have been identified at the interannual timescales of ~5-8-yr.
581 along the studied sites. In agreement of previous works (e.g., Enfield et al., 2001; Zampieri et
582 al., 2013; 2017), the effects of the AMO oscillations are mainly manifested at the interannual
583 timescales to control the variability of hydrological (e.g. rainfall) and oceanographic (e.g.
584 surges) variables. Generally, the climate oscillations of AMO are associated to the SST
585 variability with a time cyclicity of about 65-70-years (e.g. Delworth and Mann, 2000; Enfield
586 et al., 2001). During the warming periods of the 1990's, the AMO shifts from the negative to
587 the positive phases in the Northern Hemisphere corresponding to cold and warm periods (e.g.
588 Gastineau et al., 2012; Zhang et al., 2013). This shift can be responsible on changes in the
589 hydrodynamic conditions (e.g. Zampieri et al., 2013).

590 The influence of the AMO oceanic low frequencies in the modulation of the mechanisms of
591 the atmospheric teleconnections at the interannual timescales has been investigated in many
592 previous works (e.g Enfield et al. 2001). At decadal timescales, the existing relationships
593 between the winter NAO and the AMO variability is more complex (e.g. Peings and
594 Magnusdottir, 2014).

595 The effects of the AMO-driven climate variability on the seasonal weather patterns have been
596 investigated by Zampieri et al. (2017) in Europe and the Mediterranean. They have
597 demonstrated significant changes in the frequencies of weather regimes involved by the AMO
598 shifts which are in phase with seasonal surface pressure and temperature anomalies. Such
599 regimes, produced in Spring and Summer periods, are differently manifested in Europe with
600 anomalous cold conditions over Western Europe (Cassou et al., 2005; Zampieri et al., 2017).

601 In summary, four atmospheric oscillations have proven to be significantly linked to the
602 interannual and interdecadal variability of extreme surges. This physical link varies according
603 to the timescale exhibiting a nonlinear interaction of the same oscillations with other scales.

604 Such nonlinear behavior depends on the dynamics of the different sequences of the atmospheric
605 and water vapour transport patterns during the month prior to the sea-level observations (e.g.
606 Lavers et al., 2015). As suggested by Turki et al. (2020), the atmospheric circulation acts as a
607 regulator controlling the multiscale variability of extreme surges with a nonlinear connection
608 between the large-scale atmospheric circulation and the local scale hydrodynamics. This multi-
609 timescale dependence between the local extreme dynamics and the internal modes of climate
610 oscillations is still under debate. Understanding these physical links, even their complexities,
611 are useful to improve the estimation of the extreme values in coastal environments; which is
612 the objective of the next part.

613 ***5.2 Nonstationary modelling of extreme surges***

614 In this part, stationary and nonstationary extreme value analyses based on GEV distribution
615 with time-dependent parameters (Coles, 2001) have been implemented to model separately the
616 different spectral components of extreme surges. Four GEV stationary (GEV0) and
617 nonstationary (GEV1, GEV2 and GEV3) models have been applied to each timescale and each
618 site. The GEV distribution uses the maximum likelihood method by parametrizing the location,
619 scale, and shape of the model. We have used the ‘trust region reflective algorithm’ for
620 maximizing the log-likelihood function (Coleman and Li, 1996).

621 The connections between the climate oscillations and the monthly maxima at the different
622 timescales (Figure 9), preliminary presented in section 5.1, have been explored for the
623 implementation of the nonstationary GEV models. Indeed, multiple simulations of Markov
624 Chain Monte Carlo (MCMC) techniques based on Bayesian approaches have been employed
625 for extreme surge components (i.e. ~ 1.5-yr, ~ 2-4-yr, ~ 5-8-yr and ~ 12-16-yr provided by the
626 multiresolution wavelet decomposition) to identify the best covariates of climate oscillation for
627 parametrizing the nonstationary GEV models. The most of simulations has mainly supported
628 the results outlined in the previous section: the ~ 1.5-yr of SLP, ~ 2-4-yr of ZW, ~ 5-8-yr of
629 NAO and ~ 12-16-yr of AMO oscillations are considered as the best covariates for modelling
630 respectively the ~ 1.5-yr, ~ 2-4-yr, ~ 5-8-yr and ~ 12-16-yr of monthly extreme surges.

631 Once the climate covariate has been selected for each timescale, three nonstationary models
632 have been used by introducing the climate information as a covariate into: (1) the location
633 parameter (GEV1); (2) both location and scale parameters (GEV2); (3) all location, scale and
634 shape parameters (GEV3). The structure of the most appropriate nonstationary GEV
635 distribution has been selected by choosing the most adequate parametrization that minimizes
636 the Akaike information criterion (Akaike, 1974). The goodness of fit for each model has been
637 checked through the visual inspection of the quantile-quantile (Q-Q) plots (Figure 11); these
638 plots compare the empirical quantiles against the quantiles of the fitted model. Any substantial
639 departure from the diagonal indicates inadequacy of the GEV model.

640 At the interannual scales and for all sites, results provided by the nonstationary GEV1-3 reveal
641 a better performance (the lowest values of AIC) of extreme estimation compared to the
642 stationary models of GEV0 and give the most appropriate distributions by the use of the climate
643 large-scale covariates for specific oscillating components of extreme surges. Nevertheless, this
644 improvement from the stationary to the nonstationary models has not been clearly observed for
645 the interdecadal scales where the extreme estimation, provided by the different GEV models,
646 is very similar (Table 3). The lowest values of AIC have been shown by GEV3 for ~1.5-yr,
647 GEV2 for ~2-4-yr and GEV1 for ~5-8-yr (Table 4). The Q-Q plots for the all timescales of all

648 timescales of the monthly maxima in Brest are illustrated in Figure 11; they confirm the
649 suitability of the selected models.

650 *Figure 11 (a). The quantile plot between observed and modelled extreme surges by the use of the best*
651 *GEV models, at different time scales, case of Brest. (b). The Return level of extreme surges estimated*
652 *for Brest using the best GEV models. The 95% confidence interval is presented with the dashed black*
653 *line. (c) Fifty-year return level of monthly values using the original data (grey circles) and the best*
654 *nonstationary GEV model at Brest (solid black line). The lower and the upper limits of the 95%*
655 *confidence interval calculated using the Bayesian method (dashed black line). The associated*
656 *confidence area is plotted with grey shaded area.*
657

658 Accordingly, the nonstationary GEV models have exhibited high improvements at the
659 interannual scales where the AIC scores have significantly decreased by introducing the
660 climate information into the parametrization of the model. Such consideration varies as a
661 function of the spectral components, it concerns all parameters for the smallest scale of ~1.5-
662 yr, both location and scale parameters for ~2-4-yr and only the location parameter for largest
663 scale of ~5-8-yr.

664 Then, the large-scale oscillations introduced for the implementation of GEV parameters depend
665 on the time scale for all sites exhibiting a high nonstationary behaviour of the small interannual
666 scales (~1.5-yr) which decreases at the large interannual scales (~5-8-yr) and get non-
667 significant at the interdecadal scales (~ 12-16-yr).

668 The use of the time-varying GEV parameters at the interannual scales (~ 1.5 -yr and ~ 2 -4-yr)
669 exhibits the relationship between the mode and the standard deviation of the GEV distributions
670 associated with the location and the scale parameters, respectively.

671 The different implications of both parameters for estimating the interannual extreme surges
672 reveal cyclic variations and timescale modulations related to the large-scale climate
673 oscillations. As documented in the previous works (e.g., Menendez et al., 2009; Masina and
674 Lamberti., 2013), the location and the scale parameters used for improving the nonstationary
675 estimation of the extreme water levels highlight a series of annual and semi-annual evolutions.
676 They have reported that the seasonal cycles of the location parameter are related to two maxima
677 of water levels, in early March and September produced during equinoctial spring tides, while
678 the seasonal cycles of the scale parameter are associated to an increase of storms during wintry
679 episodes. Here, we focus on the stochastic signal of surges at scales larger than one year. The
680 SLP and the ZW frequencies, introduced in the location and the scale parameters of
681 nonstationary GEV models, determine an enhancement in the prediction of the interannual
682 scales. The shape parameter, implied for the estimation of the ~ 1.5 -yr extreme surges, derives
683 from its determination of the upper tail distribution behaviour. The time-varying shape
684 parameter uses the ~ 1.5 -yr SLP exhibiting altering negative and positive oscillations.

685 Despite its critical significance, the shape GEV parameter has revealed its relationships with
686 basin attributes in hydrological applications and regional flood frequency analysis (e.g., Tyralis
687 et al., 2019). The dependence of the shape parameter on the climate oscillations has been
688 demonstrated in several extreme frameworks related to hydrological and oceanographic
689 applications (e.g., Menendez et al., 2009; Masina et al., 2013; Turki et al., 2020). Regarding
690 the stationarity of the surge timescale, the ~12-16-yr window sliding matches have been
691 quantified in the previous part exhibiting a substantial cyclic variability consequence of an
692 altering periods of positive and negative correlations. The modelling of the interdecadal
693 extreme surges involves a stationary behavior of the ~12-16-yr.

694 The stationary behavior of low frequencies has been outlined by Zampieri et al. (2017). They
695 have demonstrated a stationary trend of the SST anomalies associated with the AMO over the
696 Euro-Atlantic region. According to their works, the low-frequency variability of the European
697 Climate is influenced by the AMO shift induced by the phase opposition between the negative
698 NAO distribution and the Atlantic patterns.

699 Here, the effects of AMO on ~12-16-yr of extreme surges have been largely observed in Figure
700 9 for the longer timeseries Brest where the lower frequencies could be easily identified.

701 At this timescale, the AIC values given by the different GEV models are pretty close and the
702 difference between the distributions are not statistically significant. The stationary behavior of
703 ~12-16-yr surges should be more investigated from additional applications in light of the
704 available sea level measurements covering a long period of time, a relevant parameter to
705 characterize the uncertainties in extreme value statistical modeling of flood hazards.

706 The return levels of the multiscale extreme surges, provided by the best GEV models (Table
707 3), have been simulated. The example of Brest is illustrated in Figure 10.b for the interannual
708 (nonstationary GEV models) and the interdecadal (GEV stationary model) scales. The 95%

709 confidence interval is also plotted in this graph through a dashed black line. Accordingly, the
710 use of GEV distribution with time-dependent parameters for each timescale should improve
711 the evaluation of the return values and reduce the uncertainty of the quantile estimates.

712 Similar works have been carried out by Wahls and Chambers (2016) to investigate the
713 multidecadal variations in extreme sea levels with the large-scale climate variability. By the
714 use of climate indices on nearby atmospheric/oceanic variables (winds, pressure, sea surface
715 temperature) as covariates in a quasi-nonstationary extreme value analysis, the range of change
716 in the 100-year return water levels has been significantly reduced over time, turning a
717 nonstationary process into a stationary one.

718
719 As suggested by Wong (2018), including a wider range of physical process information and
720 considering nonstationary behaviour can better enable modelling efforts to inform coastal risk
721 management. In his work, he has developed a new approach to integrate stationary and
722 nonstationary statistical models and demonstrated that the choice of covariate timeseries should
723 affect the projected flood hazards. By developing a nonstationary storm surge statistical model
724 with the use of multiple covariate timeseries (global mean temperature, sea level, the North
725 Atlantic Oscillation index and time) in Norfolk and Virginia, he has shown that a storm surge
726 model raises the projected 100-year storm surge return level by up to 23 cm relative to a
727 stationary model or one that employs a single covariate timeseries.

728 This study has expanded the previous works of Turki et al. (2019; 2020) upon a new approach
729 combining spectral and probabilistic methods to integrate multiple streams of information
730 related to climate teleconnections. Indeed, each timescale has been simulated separately with
731 the nonstationary GEV models and expressed as a function of the most suitable climate index
732 improving its fitting. The estimation of the total signal of surges should be determined by
733 combining the developed nonstationary GEV models used for the different timescales.

734 These results should support the hypothesis introduced at the beginning of the present work
735 suggesting that: (i) the extreme surges should depend on different timescales; (ii) each
736 timescale should be related to a specific large-scale oscillation.

737 The finding is in agreement with the previous works of Lee et al. (2017) and Wang et al. (2018)
738 highlighting the importance of a careful consideration when complex physical mechanisms of
739 different climate indices are included into model structures for estimating extreme surges.
740 Indeed, this work provides a guidance on incorporating nonstationary processes of large-scale
741 oscillations to different spectral components informed by the wavelet techniques, the Bayesian
742 approaches and the GEV model probabilities.

743 The primary contribution of the present research is to present a new approach for: (1)
744 investigating the multi-timescale variability of the nonstationary extreme surges; (2)
745 identifying their multi-connection with climate oscillations according to the timescale and (3)
746 resolve in part the problems of uncertainty of most appropriate climate to use as covariate for
747 GEV models at each timescale. However, additional models (e.g. significance tests and
748 sensitivity analyses and modelling uncertainties) and application sites (e.g. Mediterranean and
749 pacific ones controlled by other climate oscillations) are required to expand the developed
750 approach.

751 Also, generating a final robust stochastic model useful for projecting storm surge return levels
752 and assessing the flood risk management requires further efforts to build on the potentially
753 advantageous approach presented here by integrating the GEV models associated with the
754 different timescales through the use of mathematical methods.

755

756

6. Summary and Concluding remarks

757 The dynamics of extreme surges together with the large-scale climate oscillations have been
758 investigated by the use of hybrid methodological approach combining spectral analyses and
759 nonstationary GEV models. Results have demonstrated that the interannual variability of
760 extreme surges (~ 1.5-yr, ~2-4-yr and 5-8-yr) is around 25% for the French coasts and higher
761 than 32% for the Britannic coasts; the interdecadal variability (~12-16-yr) varies between 2%
762 and 4%. The fluctuations of extreme surges at ~1.5-yr are differently manifested between the
763 different sites of the English Channel exhibiting a local variability of surges induced by the
764 effects of meteorological and oceanographic forces including changes in atmospheric pressures
765 and wind velocities in shallow water areas. Similar fluctuations have been observed at larger
766 scales of the interannual and the interdecadal variability. Changes in extreme surges (~1.5-yr,
767 ~2-4-yr, ~5-8-yr and ~12-16-yr) have been proven to be significantly linked to atmospheric
768 oscillations (SLP, ZW, NAO and AMO, respectively) according to the timescale with a
769 nonlinear interaction between different oscillations at the same scale. This exhibits the complex
770 physical mechanisms of the global atmospheric circulation acting as a regulator and controlling
771 the local variability of extreme surges at different timescales. The connections between the
772 multiscale extreme surges and the internal modes of climate oscillations have been explored to
773 improve the estimation of extreme values by the use of nonstationary GEV models. The

774 simulated extreme surges have highlighted that introducing the climate oscillations for the
775 implementation of GEV parameters depends on the timescale for all sites; a high nonstationary
776 behaviour of the small interannual scales (~1.5-yr) decreases at the larger scales (~5-8-yr) and
777 seems to be non-significant at the interdecadal scales (~ 12-16-yr).

778 The conclusion of this research suggests that the physical mechanisms driven by the
779 atmospheric circulation, including the Gulf Stream gradients, play a key role in coastal extreme
780 surges. Establishing a strong connection of the large-scale climate oscillations with extreme
781 surges and flooding risks improve the estimation of the return levels.

782 This finding presents a handful of a new approach potentially useful as a first step forward for
783 (1) understanding the physical relation of downscaling from the global climate patterns to the
784 local extreme surges; (2) inferring the future projections of sea level change and extreme
785 events. Future work should build on this new approach to (1) improve the stochastic modelling
786 of the multi-timescale extreme surges; (2) incorporate others climate mechanisms known to be
787 important at local and regional scales for specific applications; (3) generate a robust tool for
788 the storm surge projections and the flood risk assessments based on the different timescale
789 models connected to their specific climate drivers.

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