



# 1 **Ensemble models from machine learning: an example of wave runup** 2 **and coastal dune erosion**

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## 8 Abstract

9 After decades of study and significant data collection of time-varying swash on sandy beaches, there is  
10 no single deterministic prediction scheme for wave runup that eliminates prediction error — even  
11 bespoke, locally tuned predictors present scatter when compared to observations. Scatter in runup  
12 prediction is meaningful and can be used to create probabilistic predictions of runup for a given wave  
13 climate and beach slope. This contribution demonstrates this using a data-driven Gaussian process  
14 predictor; a probabilistic machine learning technique. The runup predictor is developed using one year  
15 of hourly wave runup data (8328 observations) collected by a fixed LIDAR at Narrabeen Beach,  
16 Sydney, Australia. The Gaussian process predictor accurately predicts hourly wave runup elevation  
17 when tested on unseen data with a root mean-squared-error of 0.18 m and bias of 0.02 m. The  
18 uncertainty estimates output from the probabilistic GP predictor are then used practically in a  
19 deterministic numerical model of coastal dune erosion, which relies on a parameterization of wave  
20 runup, to generate ensemble predictions. When applied to a dataset of dune erosion caused by a storm  
21 event that impacted Narrabeen Beach in 2011, the ensemble approach reproduced ~85% of the observed  
22 variability in dune erosion along the 3.5 km beach and provided clear uncertainty estimates around  
23 these predictions. This work demonstrates how data-driven methods can be used with traditional  
24 deterministic models to develop ensemble predictions that provide more information and greater  
25 forecasting skill when compared to a single model using a deterministic parameterization; an idea that  
26 could be applied more generally to other numerical models of geomorphic systems.



## 27 1 Introduction

28 Wave runup is important for characterizing the vulnerability of beach and dune systems and coastal  
29 infrastructure to wave action. Wave runup is typically defined as the time-varying vertical elevation of  
30 wave action above ocean water levels and is a combination of wave swash and wave setup (Holman,  
31 1986; Stockdon et al., 2006). Most parameterizations of wave runup use deterministic equations that  
32 output a single value for either the maximum runup elevation in a given time period,  $R_{max}$ , or the  
33 elevation exceeded by 2% of runup events in a given time period,  $R_2$ , based on a given set of input  
34 conditions. In the majority of runup formulae, these input conditions are easily obtainable parameters  
35 such as significant wave height, significant wave period, and beach slope (Atkinson et al., 2017;  
36 Holman, 1986; Hunt, 1959; Ruggiero et al., 2001; Stockdon et al., 2006). However, wave dispersion  
37 (Guza and Feddersen, 2012), wave spectrum (Van Oorschot and d'Angremond, 1969), nearshore  
38 morphology (Cohn and Ruggiero, 2016), bore-bore interaction (García-Medina et al., 2017), tidal stage  
39 (Guedes et al., 2013), and a range of other possible processes have been shown to influence swash zone  
40 processes. Since typical wave runup parameterizations do not account for these more complex  
41 processes, there is often significant scatter in runup predictions when compared to observations (e.g.,  
42 Atkinson et al., 2017; Stockdon et al., 2006). Even flexible machine learning approaches based on  
43 extensive runup datasets or consensus-style ‘model of models’ do not resolve prediction scatter in runup  
44 datasets (e.g., Atkinson et al., 2017; Passarella et al., 2018b; Power et al., 2018). This suggests that the  
45 development of a perfect deterministic parameterization of wave runup, especially with only reduced,  
46 easily obtainable inputs (i.e., wave height, wave period, and beach slope), is improbable.

47

48 The resulting inadequacies of a single deterministic parameterization of wave runup can cascade up  
49 through the scales to cause error in any larger model that uses a runup parameterization. It therefore  
50 makes sense to clearly incorporate prediction uncertainty into wave runup predictions. In disciplines  
51 such as hydrology and meteorology, with a more established tradition of forecasting, model uncertainty  
52 is often captured by using ensembles (e.g., Bauer et al., 2015; Cloke and Pappenberger, 2009). The  
53 benefits of ensemble modelling are typically superior skill and the explicit inclusion of uncertainty in  
54 predictions by outputting a range of possible model outcomes. Commonly used methods of generating



55 ensembles include combining different models (Limber et al., 2018) or perturbing model parameters,  
56 initial conditions and/or input data (e.g., via Monte Carlo simulations (e.g., Callaghan et al., 2013)).

57

58 An alternative approach to quantify prediction uncertainty is to incorporate scatter about a mean  
59 prediction into model parameterizations. For example, wave runup predictions at every time step could  
60 be modelled with a deterministic parameterization plus a noise component that captures the scatter  
61 about the deterministic prediction caused by unresolved processes. If parameterizations are stochastic,  
62 or have a stochastic component, repeated model runs (given identical initial and forcing conditions)  
63 produce different model outputs – an ensemble – that represents a range of possible values the process  
64 could take. This is broadly analogous to the method of “stochastic parameterization” used in the  
65 weather forecasting community for sub-grid scale processes and parameterizations (Berner et al., 2017).  
66 In these applications, stochastic parameterization has been shown to produce better predictions than  
67 traditional ensemble methods and is now routinely used by many operational weather forecasting  
68 centers (Berner et al., 2017; Buchanan, 2018).

69

70 Stochastically varying a deterministic wave runup parameterization to form an ensemble still requires  
71 defining the stochastic term — i.e., the stochastic element that should be added to the predicted runup at  
72 each model time step. An alternative to specifying a predefined distribution or a noise term added to a  
73 parameterization is to learn and parameterize the variability in wave runup from observational data  
74 using machine learning techniques. Machine learning has had a wide range of applications in coastal  
75 morphodynamics research (Goldstein et al., 2018) and has shown specific utility in understanding swash  
76 processes (Passarella et al., 2018b; Power et al., 2018) as well as storm driven erosion (Beuzen et al.,  
77 2018; den Heijer et al., 2012; Goldstein and Moore, 2016; Palmsten et al., 2014; Plant and Stockdon,  
78 2012). While many machine learning algorithms and applications are often used to optimize  
79 deterministic predictions, a Gaussian process is a probabilistic machine learning technique that directly  
80 captures model uncertainty from data (Rasmussen and Williams, 2006). Recent work has specifically  
81 used Gaussian processes to understand coastal processes such as large scale coastline erosion (Kupilik  
82 et al., 2018).



83

84 The work presented here is focused on using a Gaussian process to build a data-driven probabilistic  
85 predictor of wave runup that includes estimates of uncertainty. While quantifying uncertainty in runup  
86 predictions from data is useful in itself, the benefit of this methodology is in explicitly including the  
87 uncertainty with the runup predictor in a larger model that uses a runup parametrization, such as a  
88 coastal dune erosion model. Dunes on sandy coastlines provide a natural barrier to storm erosion by  
89 absorbing the impact of incident waves and storm surge and helping to prevent or delay flooding of  
90 coastal hinterland and infrastructure (Mull and Ruggiero, 2014; Sallenger, 2000; Stockdon et al., 2007).  
91 The accurate prediction of coastal dune erosion is therefore critical for characterizing the vulnerability  
92 of dune and beach systems and coastal infrastructure to storm events. A variety of methods are available  
93 for modelling dune erosion including: simple conceptual models relating hydrodynamic forcing,  
94 antecedent morphology and dune response (Sallenger, 2000); empirical dune-impact models that relate  
95 time-dependent dune erosion to the force of wave impact at the dune (Erikson et al., 2007; Larson et al.,  
96 2004; Palmsten and Holman, 2012); data-driven machine learning models (Plant and Stockdon, 2012);  
97 and more complex physics-based models (Roelvink et al., 2009). In this study, we focus on dune-impact  
98 models, which are simple, commonly used models that typically rely on a parameterization of wave  
99 runup to model time-dependent dune erosion. As inadequacies in the runup parameterization can  
100 jeopardize the success of model results (Overbeck et al., 2017; Palmsten and Holman, 2012; Splinter et  
101 al., 2018), it makes sense to use a runup predictor that includes prediction uncertainty.

102

103 The overall aim of this work is to demonstrate how probabilistic data-driven methods can be used with  
104 deterministic models to develop ensemble predictions, an idea that could be applied more generally to  
105 other numerical models of geomorphic systems. **Sect. 2** first describes the Gaussian process model  
106 theory. In **Sect. 3** the Gaussian process runup predictor is developed. In **Sect. 4** an example application  
107 of the Gaussian process predictor of runup inside a morphodynamic model of coastal dune erosion to  
108 build a ‘hybrid’ model (Goldstein and Coco, 2015; Krasnopolsky and Fox-Rabinovitz, 2006) that can  
109 generate ensemble output is presented. A discussion of the results and technique is provided in **Sect. 5**  
110 followed by conclusions in **Sect. 6**. The data and code used to develop the Gaussian Process runup



111 predictor in this manuscript are publicly available at  
112 [https://github.com/TomasBeuzen/BeuzenEtAl\\_GP\\_Paper](https://github.com/TomasBeuzen/BeuzenEtAl_GP_Paper).



## 113 2 Gaussian Processes

### 114 2.1 Gaussian Process Theory

115 Gaussian processes (GPs) are data-driven, non-parametric models. A brief introduction to GPs is given  
116 here; for a more detailed introduction the reader is referred to Rasmussen and Williams (2006). There  
117 are two main approaches to determine a function that best parameterizes a process over an input space:  
118 1) select a class of functions to consider, e.g., polynomial functions, and best fit the functions to the data  
119 (a parametric approach); or, 2) consider all possible functions that could fit the data, and assign higher  
120 weight to functions that are more likely (a non-parametric approach) (Rasmussen and Williams, 2006).  
121 In the first approach it is necessary to decide on a class of functions to fit to the data – if all or parts of  
122 the data are not well modelled by the selected functions, then the predictions may be poor. In the second  
123 approach there is an infinite set of possible functions that could fit a data set (imagine the number of  
124 paths that could be drawn between two points on a graph). A GP addresses the problem of infinite  
125 possible functions by specifying a probability distribution over the space of possible functions that fit a  
126 given dataset. Based on this distribution, the GP quantifies what function most likely fits the underlying  
127 process generating the data and gives confidence intervals for this estimate. Additionally, random  
128 samples can also be drawn from the distribution to provide examples of what different functions that fit  
129 the dataset might look like.

130

131 A GP is defined as a collection of random variables, any finite set of which has a multivariate Gaussian  
132 distribution. The random variables in a GP represent the value of the underlying function that describes  
133 the data,  $f(x)$ , at location  $x$ . The typical workflow for a GP is to define a prior distribution over the space  
134 of possible functions that fit the data, form a posterior distribution by conditioning the prior on observed  
135 input/output data pairs (“training data”), and to then use this posterior distribution to predict unknown  
136 outputs at other input values (“testing data”). The key to GP modelling is the use of the multivariate  
137 Gaussian distribution, which has simple closed form solutions to the aforementioned conditioning  
138 process, as described below.



139

140 Whereas a univariate Gaussian distribution is defined by a mean and variance (i.e.,  $\mathcal{N}(\mu, \sigma^2)$ ), a GP (a  
141 multivariate Gaussian distribution) is completely defined by a mean function  $m(\mathbf{x})$  and covariance  
142 function  $k(\mathbf{x}, \mathbf{x}')$  (also known as a “kernel”), and is typically denoted:

143

$$144 \quad f(\mathbf{x}) \sim \mathcal{N}(m(\mathbf{x}), k(\mathbf{x}, \mathbf{x}')) \quad (1)$$

145

146 Where  $\mathbf{x}$  is an input vector of dimension  $D$  ( $\mathbf{x} \in \mathbb{R}^D$ ), and  $f$  is the unknown function describing the data.  
147 Note that for the remainder of this paper, a variable denoted in bold text represents a vector. The mean  
148 function,  $m(\mathbf{x})$ , describes the expected mean value of the function describing the data at location  $\mathbf{x}$ ,  
149 while the covariance function encodes the correlation between the function values at locations in  $\mathbf{x}$ .

150

151 These concepts of GP development are further described using a hypothetical dataset of significant  
152 wave height ( $H_s$ ) versus wave runup ( $R_2$ ) shown in **Fig. 1A**. The first step of GP modelling is to  
153 constrain the infinite set of functions that could fit a dataset by defining a prior distribution over the  
154 space of functions. This prior distribution encodes belief about what the underlying function is expected  
155 to look like (e.g., smooth/erratic, cyclic/random, etc.) before constraining the model with any observed  
156 training data. Typically it is assumed that the mean function of the GP prior,  $m(\mathbf{x})$ , is 0 everywhere, to  
157 simplify notation and computation of the model (Rasmussen and Williams, 2006). Note that this does  
158 not limit the GP posterior to be a constant mean process. The covariance function,  $k(\mathbf{x}, \mathbf{x}')$ , ultimately  
159 encodes what the underlying functions look like because it controls how similar the function value at  
160 one input point is to the function value at other input points.

161

162 There are many different types of covariance functions or “kernels”. One of the most common, and the  
163 one used in this study, is the squared exponential covariance function:

164

$$165 \quad k(x_i, x_j) = \sigma_f^2 \exp \left[ - \sum_{d=1}^D \frac{1}{2l_d^2} (x_{d,i} - x_{d,j})^2 \right] \quad (2)$$

166





167 Where  $\sigma_f$  is the signal variance and  $l$  is known as the length-scale, both of which are hyperparameters in  
168 the model that can be estimated from data (discussed further in **Sect. 2.2**). Together the mean function  
169 and covariance function specify a multivariate Gaussian distribution:

170

$$171 \quad f(\mathbf{x}) \sim \mathcal{N}(\mathbf{0}, K) \quad (3)$$

172

173 Where  $f$  is the output of the prior distribution, the mean function is assumed to be  $\mathbf{0}$  and  $K$  is the  
174 covariance matrix made by evaluating the covariance function at arbitrary input points that lie within  
175 the domain being modelled (i.e.,  $K(x, x)_{i,j} = k(x_i, x_j)$ ). Random sample functions can be drawn from this  
176 prior distribution as demonstrated in **Fig. 1B**.

177

178 The goal is to determine which of these functions actually fit the observed data points (training data) in  
179 **Fig. 1A**. This can be achieved by forming a posterior distribution on the function space by conditioning  
180 the prior with the training data. Roughly speaking, this operation is mathematically equivalent to  
181 drawing an infinite number of random functions from the multivariate Gaussian prior (**Eq. (3)**), and  
182 then rejecting those that do not agree with the training data. As mentioned above, the multivariate  
183 Gaussian offers a simple, closed form solution to this conditioning. Assuming that our observed training  
184 data is noiseless (i.e.,  $y$  exactly represents the value of the underlying function  $f$ ) then we can condition  
185 the prior distribution with the training data samples  $(\mathbf{x}, y)$  to define a posterior distribution of the  
186 function value ( $f^*$ ) at arbitrary test inputs  $(\mathbf{x}^*)$ :

187

$$188 \quad f^* | \mathbf{y} \sim \mathcal{N}(K_* K^{-1} \mathbf{y}, K_{**} - K_* K^{-1} K_*^T) \quad (4)$$

189

190 Where  $f^*$  is the output of the posterior distribution at the desired test points  $\mathbf{x}^*$ ,  $\mathbf{y}$  is the training data  
191 outputs at inputs  $\mathbf{x}$ ,  $K_*$  is the covariance matrix made by evaluating the covariance function (**Eq. (2)**)  
192 between the test inputs  $\mathbf{x}^*$  and training inputs  $\mathbf{x}$  (i.e.,  $k(\mathbf{x}^*, \mathbf{x})$ ),  $K$  is the covariance matrix made by  
193 evaluating the covariance function between training data points  $\mathbf{x}$ , and  $K_{**}$  is the covariance matrix  
194 made by evaluating the covariance function between test points  $\mathbf{x}^*$ . Function values can be sampled



195 from the posterior distribution as shown in **Fig. 1C**. These samples represent random realizations of  
196 what the underlying function describing the training data could look like.

197

198 As stated earlier, in **Eq. (4)** and **Fig. 1C** there is an assumption that the training data is noiseless and  
199 represents the exact value of the function at the specific point in input space. In reality, there is error  
200 associated with observations of physical systems, such that:

201

$$202 \mathbf{y} = f(\mathbf{x}) + \varepsilon \quad (5)$$

203

204 Where  $\varepsilon$  is assumed to be independent identically distributed Gaussian noise with variance  $\sigma_n^2$ . This  
205 noise can be incorporated into the GP modelling framework through the use of a white noise kernel that  
206 adds an element of Gaussian white noise into the model:

207

$$208 k(x_i, x_j) = \sigma_n^2 \delta_{ij} \quad (6)$$

209

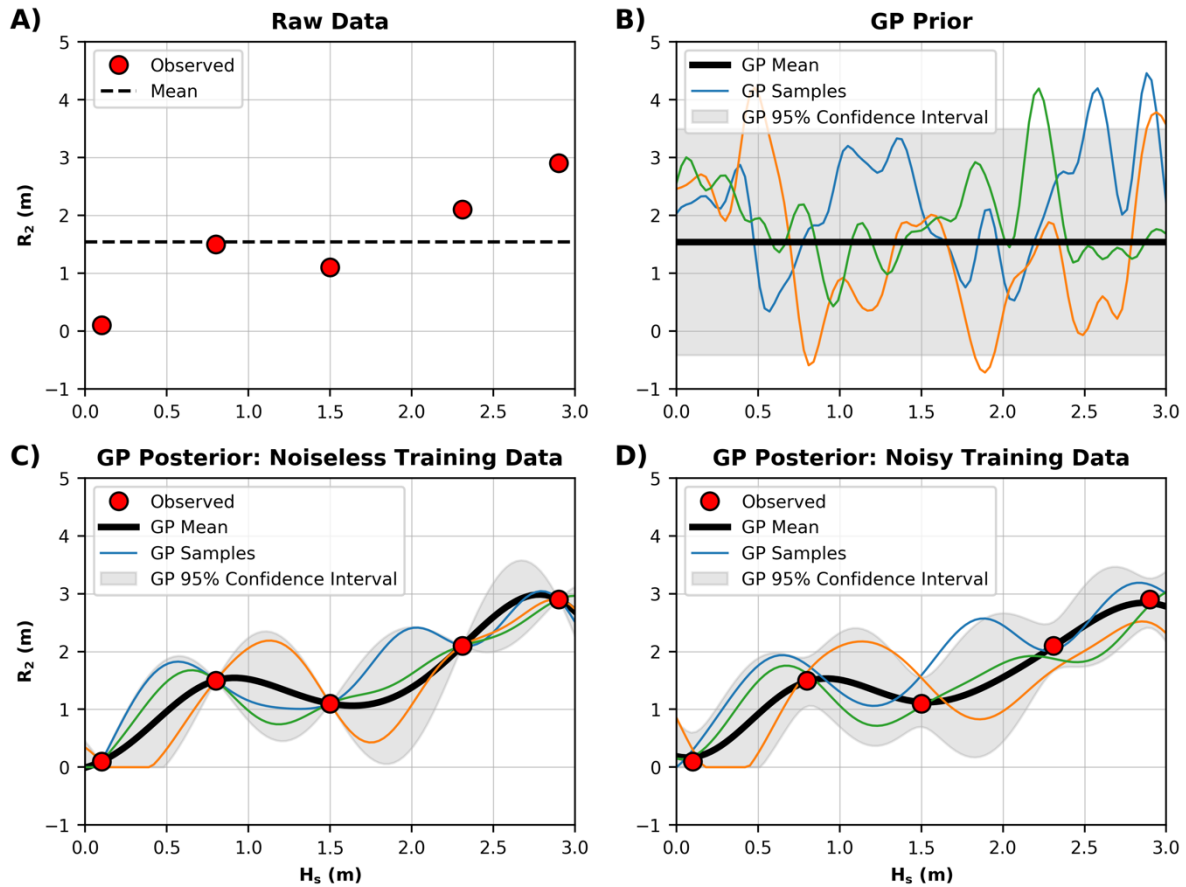
210 Where  $\sigma_n^2$  is the variance of the noise and  $\delta_{ij}$  is a Kronecker delta which is 1 if  $i = j$  and 0 otherwise.  
211 The squared exponential kernel and white noise kernel are closed under addition and product  
212 (Rasmussen and Williams, 2006), such that they can simply be combined to form a custom kernel for  
213 use in the GP:

214

$$215 k(x_i, x_j) = \sigma_f^2 \exp \left\{ - \sum_{d=1}^D \frac{1}{2l_d^2} (x_{d,i} - x_{d,j})^2 \right\} + \sigma_n^2 \delta_{ij} \quad (7)$$

216

217 The combination of kernels to model different signals in a dataset (that vary over different spatial or  
218 temporal timescales) is common in applications of GPs (Rasmussen and Williams, 2006; Reggente et  
219 al., 2014; Roberts et al., 2013). Samples drawn from the resultant “noisy” posterior distribution are  
220 shown in **Fig. 1D** in which the GP can now be seen to not fit the observed training data precisely.



221

222 Fig. 1: A) Five hypothetical random observations of significant wave height ( $H_s$ ) and 2% wave runup elevation ( $R_2$ ). B) The  
 223 Gaussian process (GP) prior distribution. C) The GP posterior distribution, formed by conditioning the prior distribution in (B)  
 224 with the observed data points in (A), assuming the observations are noise-free. D). The GP posterior distribution conditioned on  
 225 the observations with a noise component.

226

## 227 2.2 Gaussian Process Kernel Optimization

228 In Eq. (7) there are three hyperparameters: the signal variance ( $\sigma_f$ ), the length scale ( $l$ ) and the noise  
 229 variance ( $\sigma_n$ ). These hyperparameters are typically unknown but can be estimated and optimized based  
 230 on the particular dataset. Here, this optimization is performed by using the typical methodology of  
 231 maximizing the log-marginal-likelihood of the observed data  $\mathbf{y}$  given the hyperparameters:

232



233  $\log p(y|x, \sigma_f, l, \sigma_n)$  (8)

234

235 The Python toolkit SciKit-Learn (Pedregosa et al., 2011) was used to develop the GP described in this  
236 study.

### 237 **2.3 Training a Gaussian Process Model**

238 It is standard practice in the development of data-driven machine learning models to divide the available  
239 dataset into training, validation and testing subsets. The training data is used to fit model parameters.  
240 The validation data is used to evaluate model performance and the model hyperparameters are usually  
241 varied until performance on the validation data is optimized. Once the model is optimized, the  
242 remaining test dataset is used to objectively evaluate its performance and generalizability. A decision  
243 must be made about how to split a dataset into training, validation and testing subsets. There are many  
244 different approaches to handle this splitting process; for example, random selection, cross-validation,  
245 stratified sampling, or a number of other deterministic sampling techniques (Camus et al., 2011). The  
246 exact technique used to generate the data subsets often depends on the problem at hand. Here, there  
247 were two constraints to be considered; first, the computational expense of GPs scales by  $O(n^3)$   
248 (Rasmussen and Williams, 2006), so it is desirable to keep the training set as small as possible without  
249 deteriorating model performance; and, secondly, machine learning models typically perform poorly  
250 with out-of-sample predictions (i.e., extrapolation), so it is desirable to include in the training set the  
251 data samples that captures the full range of variability in the data. Based on these constraints, we used a  
252 maximum dissimilarity algorithm (MDA) to divide the available data into training, validation and  
253 testing sets.

254

255 The MDA is a deterministic routine that iteratively adds a data point to the training set based on how  
256 dissimilar it is to the data already included in the training set. Camus et al. (2011) provide a  
257 comprehensive introduction to the MDA selection routine and it has been previously used in ML studies  
258 (e.g., Goldstein et al., 2013). Briefly, to initialize the MDA routine, the data point with the maximum  
259 sum of dissimilarity (defined by Euclidean distance) to all other data points is selected as the first data



260 point to be added to the training data set. Additional data points are included in the training set through  
261 an iterative process whereby the next data point added is the one with maximum dissimilarity to those  
262 already in the training set - this process continues until a user-defined training set size is reached. In this  
263 way the MDA routine produces a set of training data that captures the range of variability present in the  
264 full dataset. The data not selected for the training set are equally and randomly split to form the  
265 validation dataset and test dataset.

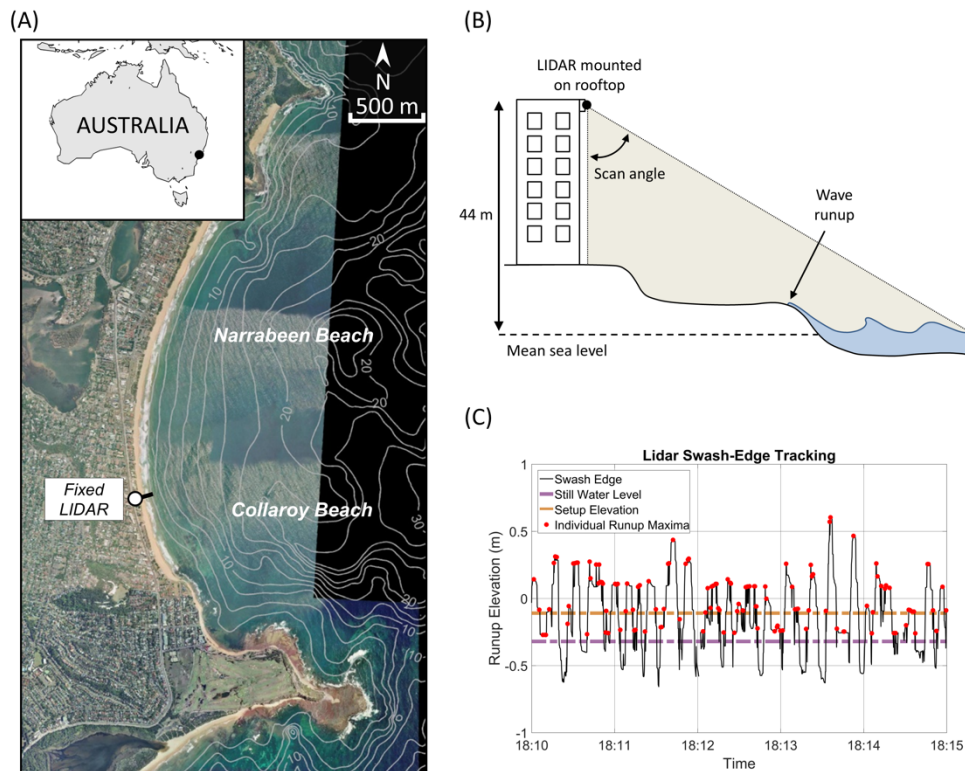


## 266 3 Development of a Gaussian Process Runup Model

### 267 3.1 Runup Data

268 In 2014, an extended-range LIDAR (LIght Detection And Ranging) device (SICK LD-LRS 2110) was  
269 permanently installed on the rooftop of a beachside building (44 m above mean sea level) at Narrabeen-  
270 Collaroy Beach (hereafter referred to simply as Narrabeen) on the south-east coast of Australia (**Fig. 2**).  
271 Since 2014, this LIDAR has continuously scanned a single cross-shore profile transect extending from  
272 the base of the beachside building to a range of 130 m, capturing the surface of the beach profile and  
273 incident wave swash at a frequency of 5 Hz in both daylight and non-daylight hours. Specific details of  
274 the LIDAR setup and functioning can be found in (Phillips et al., 2019).

275  
276



277

278 **Fig. 2:** A) Narrabeen Beach, located on the southeast coast of Australia. B) Conceptual figure of the fixed LIDAR setup. C) A five-  
279 minute extract of runup elevation extracted from the LIDAR data, individual runup maxima are marked with red circles.



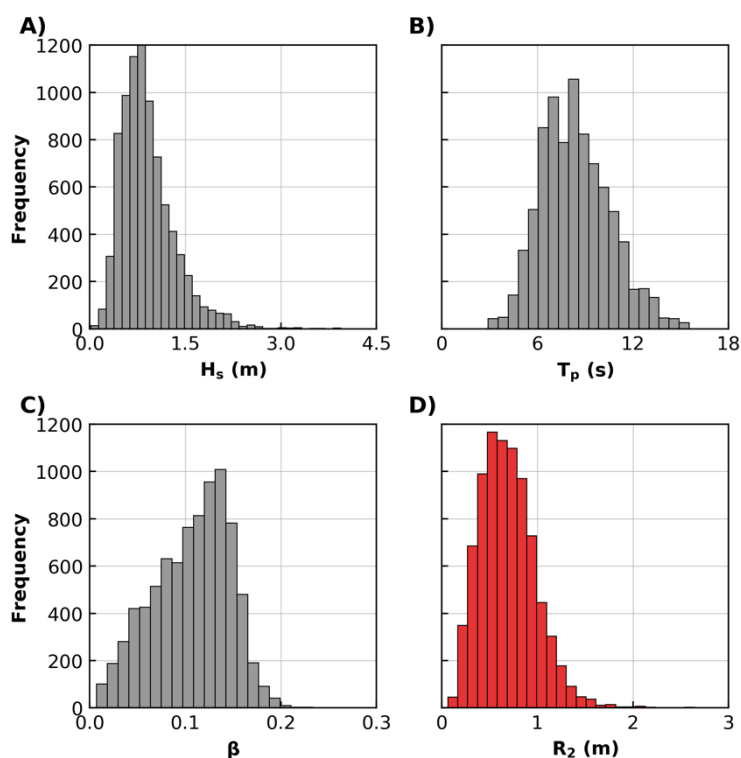
280

281 Narrabeen Beach is a 3.6 km long embayed beach bounded by rocky headlands. It is composed of fine  
282 to medium quartz sand ( $D_{50} \approx 0.3$  mm), with a  $\sim 30\%$  carbonate fraction. Offshore, the coastline has a  
283 steep and narrow (20 – 70 km) continental shelf (Short and Trenaman, 1992). The region is microtidal  
284 and semidiurnal with a mean spring tidal range of 1.6 m and has a moderate to high energy deep water  
285 wave climate characterized by persistent long-period SSE swell waves that is interrupted by storm  
286 events (significant wave height  $> 3$  m) typically 10 – 20 times per year (Short and Trenaman, 1992). In  
287 the present study, approximately one year of the high-resolution wave runup LIDAR dataset available at  
288 Narrabeen is used to develop a data-driven parameterization of the 2% exceedance of wave runup ( $R_2$ ).  
289 Data used to develop this parameterization were at hourly resolution and include:  $R_2$ , the beach slope  
290 ( $\beta$ ), offshore significant wave height ( $H_s$ ), and peak wave period ( $T_p$ ). These data are described below  
291 and have been commonly used to parameterize  $R_2$  in other empirical models of wave runup (e.g.,  
292 Holman, 1986; Hunt, 1959; Stockdon et al., 2006).

293

294 Individual wave runup elevation on the beach profile was extracted on a wave-by-wave basis from the  
295 LIDAR dataset (**Fig. 2C**). Hourly  $R_2$  was calculated as the 2% exceedance value for a given hour of  
296 wave runup observations.  $\beta$  was calculated as the linear (best-fit) slope of the beach profile over which  
297 two standard deviations of wave runup values were observed during the hour. Hourly  $H_s$  and  $T_p$  data  
298 were obtained from the Sydney Wave Rider buoy, situated 11 km offshore of Narrabeen in  $\sim 80$  m  
299 water depth. Narrabeen is an embayed beach, where prominent rocky headlands both attenuate and  
300 refract incident waves. To remove these effects in the wave data and to emulate an open coastline and  
301 generalize the parameterization of  $R_2$  presented in this study, offshore wave data were first transformed  
302 to a nearshore equivalent (10 m water depth) using the SWAN spectral wave model (Booij et al., 1999),  
303 and then reverse shoaled back to deep water wave data. A total of 8328 hourly samples of  $R_2$ ,  $\beta$ ,  $H_s$  and  
304  $T_p$  were extracted to develop a parameterization of  $R_2$  in this study. Histograms of this data are shown in  
305 **Fig. 3**.

306



307

308 Fig. 3: Histograms of the 8328 data samples extracted from the Narrabeen LIDAR: (A) significant wave height ( $H_s$ ); (B) peak wave  
309 period ( $T_p$ ); (C) beach slope ( $\beta$ ); and, (D) 2% wave runup elevation ( $R_2$ ).

### 310 3.2 Training Data for the GP Runup Predictor

311 To determine the optimum training set size, kernel and model hyperparameters, a number of different  
312 user-defined training set sizes were trialed using the MDA selection routine discussed in Sect. 2.3. The  
313 GP was trained using different amounts of data and hyperparameters were optimized on the validation  
314 data set only. It was found that a training set size of only 5% of the available dataset (training dataset =  
315 416 of 8328 available samples, validation dataset = 3956 samples, testing dataset = 3956 samples) was  
316 required to develop an optimum GP model. Training data sizes beyond this value produced negligible  
317 changes in GP performance but considerable increases in computational demand, similar to findings of  
318 previous work (Goldstein and Coco, 2014; Tinoco et al., 2015). Results presented below discuss the  
319 performance of the GP on the testing dataset which was not used in GP training or validation.

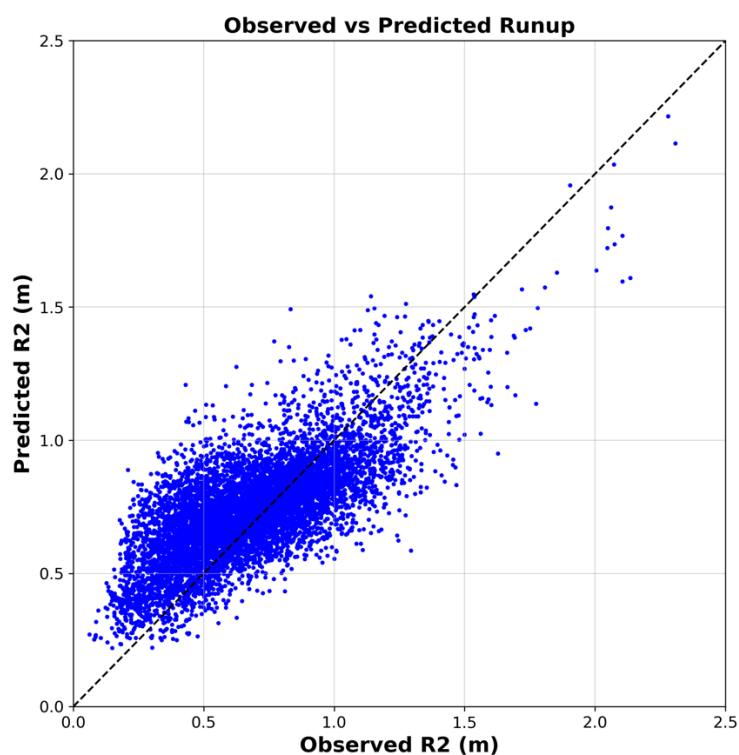




### 320 3.3 Runup Predictor Results

321 Results of the GP  $R_2$  predictor on the 3956 test samples are shown in **Fig. 4**. This figure plots the mean  
322 GP predictions against corresponding observations of  $R_2$ . The mean GP prediction performs well on the  
323 test data, with a root-mean-squared-error (RMSE) of 0.18 m and bias (B) of 0.02 m. For comparison,  
324 the commonly used  $R_2$  parameterization of Stockdon et al. (2006) tested on the same data has a RMSE  
325 of 0.36 m and B of 0.21 m. Despite the relatively accurate performance of the GP on this dataset, there  
326 remains significant scatter in the observed versus predicted  $R_2$  in **Fig. 4**. This is consistent with recent  
327 work by Atkinson et al. (2017) showing that commonly used predictors of  $R_2$  always result in scatter.

328



329

330 **Fig. 4: Observed 2% wave runup ( $R_2$ ) versus the  $R_2$  predicted by the Gaussian process model. Root-mean-squared-error (RMSE)**  
331 **is 0.36 m, bias (B) is 0.02 m and squared correlation ( $r^2$ ) is 0.54.**

332

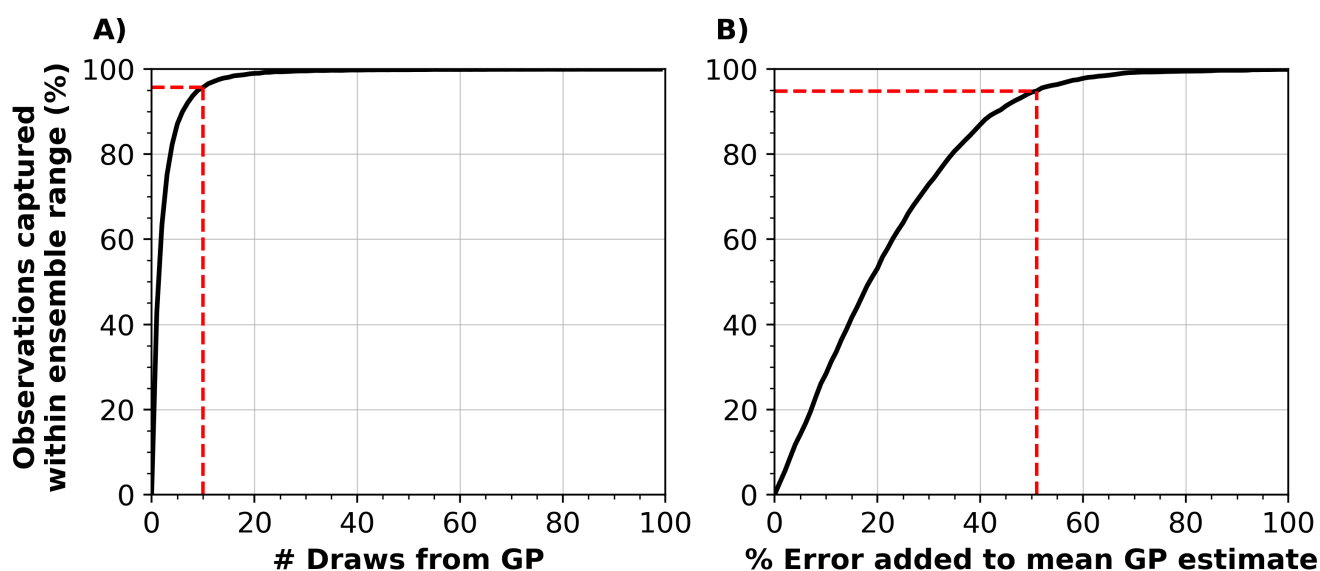
333 Here the scatter (uncertainty) is used to form ensemble predictions. The GP developed here not only  
334 gives a mean prediction as used in **Fig. 4**, but it specifies a multivariate Gaussian distribution from  
335 which different random functions that describe the data can be sampled. Random samples of wave



336 runup from the GP can capture uncertainty around the mean runup prediction (as was demonstrated in  
337 the hypothetical example in **Fig. 1D**). To assess how well the GP model captures uncertainty, random  
338 samples are successively drawn from the GP and the number of  $R_2$  measurements captured with each  
339 new draw are determined. Only 10 random samples drawn from the GP are required to capture 95% of  
340 the scatter in  $R_2$  (**Fig. 5A**). This process of drawing random samples from the GP was repeated 100  
341 times with results showing that the above is true for any 10 random samples, with an average capture  
342 percentage of 95.7% and range of 94.9% to 96.1% for 10 samples across the 100 trials. As a point of  
343 contrast, **Fig. 5B** shows how much arbitrary error would need to be added to the mean  $R_2$  prediction to  
344 capture scatter about the mean to emulate the uncertainty captured by the GP. It can be seen that the  
345 mean  $R_2$  prediction would need to vary by  $\pm 51\%$  to capture 95% of the scatter present in the runup  
346 data. This demonstrates how random models of runup drawn from the GP effectively capture  
347 uncertainty in  $R_2$  predictions. These randomly drawn  $R_2$  models can be used within a larger dune-impact  
348 model to produce an ensemble of dune erosion predictions that includes uncertainty in runup  
349 predictions, as demonstrated in **Sect. 4**.

350

351



352



353 Fig. 5: A) Percent of observed runup values captured within the range of ensemble predictions made by randomly sampling  
354 different runup values from the Gaussian process. Only 10 randomly drawn models can form an ensemble that captures 95% of  
355 the scatter in observed  $R_2$  values. B) An experiment showing how much arbitrary error would need to be added to the mean GP  
356 runup prediction in order to capture scatter in  $R_2$  observations. The mean GP prediction would have to vary by 51% in order to  
357 capture 95% of scatter in  $R_2$  observations.



## 358 4 Application of a Gaussian Process Runup Predictor in a Coastal Dune Erosion Model

### 359 4.1 Dune Erosion Model

360 We use the dune erosion model of Larson et al. (2004) as an example of how the GP runup predictor  
361 can be used to create an ensemble of dune erosion predictions, and thus provide probabilistic outcomes  
362 with uncertainty bands needed in coastal management. The dune erosion model is subsequently referred  
363 to as LEH04 and is defined as follows:

364

$$365 \quad dV = 4C_s(R_2 - z_b)^2\left(\frac{t}{T}\right) \quad (9)$$

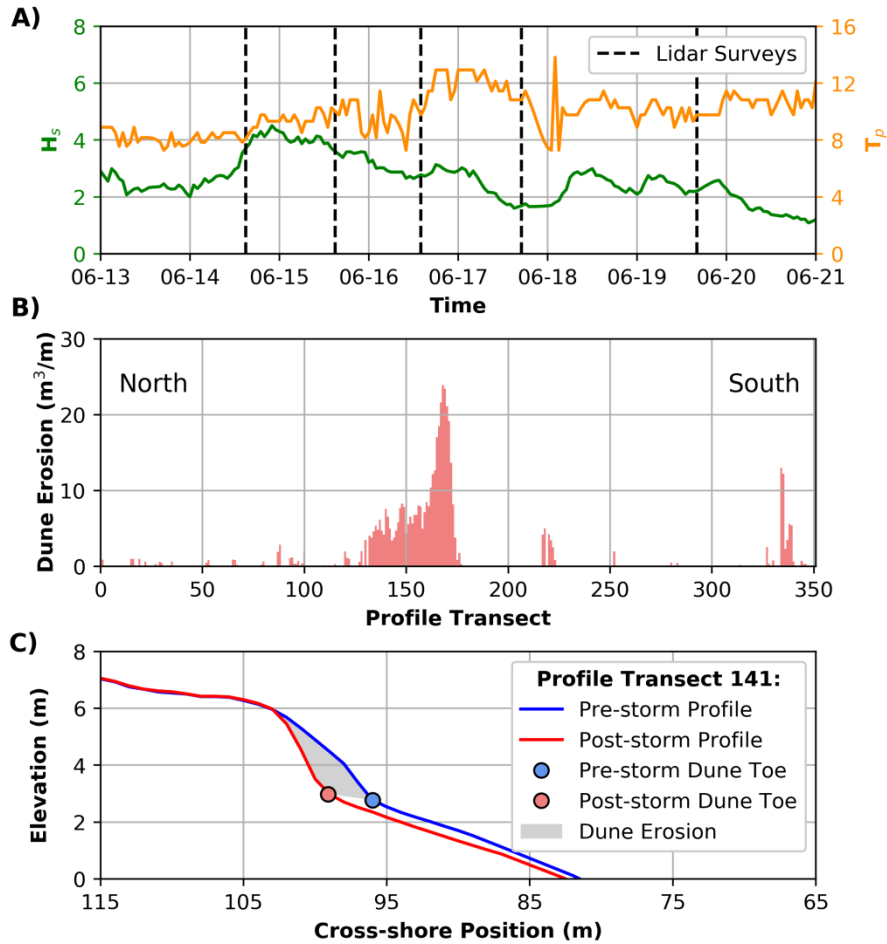
366

367 Where  $dV$  ( $\text{m}^3/\text{m}$ ) is the volumetric dune erosion per unit width alongshore for a given time step  $t$ ,  $z_b$   
368 (m) is the time-varying dune toe elevation,  $T$  (s) is the wave period for that time step,  $R_2$  (m) is the 2%  
369 runup exceedance for that time step, and  $C_s$  is the transport coefficient. Note that the original equation  
370 used a best-fit relationship to define the runup term,  $R$  (see Eq. (36) in Larson et al., 2004) rather than  
371  $R_2$ . Subsequent modifications of the LEH04 model have been made to adjust the collision frequency  
372 (i.e. the  $t/T$  term; e.g., Palmsten and Holman (2012), Splinter and Palmsten (2012)), however we retain  
373 the model presented in **Eq. (9)** for the purpose of providing a simple illustrative example. At each time  
374 step, dune volume is eroded in bulk and the dune toe is adjusted along a predefined slope (defined here  
375 as the linear slope between the pre- and post-storm dune toe) so that erosion causes the dune toe to  
376 increase in elevation and recede landward. Dune erosion and dune toe recession only occurs when wave  
377 runup ( $R_2$ ) exceeds the dune toe (i.e.,  $R_2 - z_b > 0$ ) and cannot progress vertically beyond the maximum  
378 runup elevation. When  $R_2$  does not exceed  $z_b$ ,  $dV = 0$ . The GP  $R_2$  predictor described in **Sect. 3** is used  
379 to stochastically parameterize wave runup in the LEH04 model and form ensembles of dune erosion  
380 predictions. The model is applied to new data not used to train the GP  $R_2$  predictor, using detailed  
381 observations of dune erosion caused by a large coastal storm event at Narrabeen Beach, southeast  
382 Australia in 2011.



## 383 4.2 June 2011 Storm Data

384 In June 2011 a large coastal storm event impacted the southeast coast of Australia. This event resulted  
385 in variable alongshore dune erosion at Narrabeen Beach, which was precisely captured by airborne  
386 LIDAR immediately pre-, during, and post-storm by five surveys conducted approximately 24 hours  
387 apart. Cross-shore profiles were extracted from the Lidar data at 10 m alongshore intervals as described  
388 in detail in Splinter et al. (2018), resulting in 351 individual profiles (**Fig. 6**). The June 2011 storm  
389 lasted 120 hours. Hourly wave data was recorded by the Sydney wave rider buoy located in ~80 m  
390 water depth directly to the southeast of Narrabeen Beach. As with the hourly wave data used to develop  
391 the GP model of  $R_2$  (**Sect. 3.1**), hourly wave data for each of the 351 profiles for the June 2011 storm  
392 was obtained by first transforming offshore wave data to the nearshore equivalent at 10 m water depth  
393 directly offshore of each profile using the SWAN spectral wave model (Booij et al., 1999), and then  
394 reverse shoaling back to equivalent deep water wave data, to account for the effects of wave refraction  
395 and attenuation caused by the distinctly curved Narrabeen embayment. The tidal range during the storm  
396 event was measured in-situ at the Fort Denison Tide Gauge (located within Sydney Harbour  
397 approximately 16 km south of Narrabeen) as 1.58 m (mean spring tidal range at Narrabeen is 1.6 m).  
398 The hydrodynamic time series and airborne LIDAR observations of dune change are used to  
399 demonstrate how the LEH04 model can be used with the GP predictor of runup to generate stochastic  
400 parameterizations and create probabilistic model ensembles (**Eq. (9)**).



401

402 Fig. 6: June 2011 storm data. A) Offshore  $H_s$  and  $T_p$  with vertical dashed lines indicating the time of the LIDAR surveys, B)  
403 Measured (pre vs post storm) dune erosion volumes for the 351 profile transects extracted from LIDAR data, C) Example pre-  
404 (blue) and post-storm (red) profile cross sections showing dune toes (coloured circles) and dune erosion volume (grey shading).

405 For each of the 351 available profiles, the pre-, during and post-storm dune toe positions were defined  
406 as the local maxima of curvature of the beach profile following the method of Stockdon et al. (2007).  
407 Dune erosion at each profile was then defined as the difference in subaerial beach volume landward of  
408 the pre-storm dune toe, as shown in Fig. 6C. Of the 351 profiles, only 117 had storm driven dune  
409 erosion (Fig. 6B). For the example demonstration presented here, only profiles for which the post-storm  
410 dune toe elevation was at the same or higher elevation than the pre-storm dune toe are considered;  
411 which is a basic assumption of the LEH04 model. Of the 117 profiles with storm erosion, 40 profiles



412 met these criteria. For each of these profiles, the linear slope between the pre- and post-storm dune toe  
413 was used to project the dune erosion calculated using the LEH04 model.

414

415 The LEH04 dune erosion model (**Eq. (9)**) has a single tuneable parameter, the transport coefficient  $C_s$ .  
416 There is ambiguity in the literature regarding the value of  $C_s$ . Larson et al. (2004) developed an  
417 empirical equation to relate  $C_s$  to wave height ( $H_{rms}$ ) and grain size ( $D_{50}$ ) using experimental data.  
418 Values ranged from  $1 \times 10^{-5}$  to  $1 \times 10^{-1}$ , and Larson et al. (2004) used  $1.7 \times 10^{-4}$  based on field data from  
419 Birkemeier et al. (1988). Palmsten and Holman (2012) used LEH04 to model dune erosion observed in  
420 a large wave tank experiment conducted at the O.H. Hinsdale Wave Research Laboratory at Oregon  
421 State University. The model was shown to accurately reproduce dune erosion when applied in hourly  
422 time steps using a  $C_s$  of  $1.34 \times 10^{-3}$ , based on the empirical equation determined by Larson et al. (2004).  
423 Mull and Ruggiero (2014) used values of  $1.7 \times 10^{-4}$  and  $1.34 \times 10^{-3}$  as lower and upper bounds of  $C_s$  to  
424 model dune erosion caused by a large storm event on the Pacific Northwest Coast of the USA and the  
425 laboratory experiment used by Palmsten and Holman (2012). For the dune erosion experiment, the  
426 value of  $1.7 \times 10^{-4}$  was found to predict dune erosion volumes closest to the observed erosion when  
427 applied in a single time step, with an optimum value of  $2.98 \times 10^{-4}$ . Splinter and Palmsten (2012) found  
428 a best fit  $C_s$  of  $4 \times 10^{-5}$  in an application to modelling dune erosion caused by a large storm event that  
429 occurred on the Gold Coast, Australia. Ranasinghe et al. (2012) found a  $C_s$  value of  $1.5 \times 10^{-3}$  in an  
430 application at Narrabeen Beach, Australia. It is noted that  $C_s$  values in these studies are influenced by  
431 the time step used in the model and the exact definition of wave runup,  $R$ , used (Larson et al., 2004;  
432 Mull and Ruggiero, 2014; Palmsten and Holman, 2012; Splinter and Palmsten, 2012). In practice,  $C_s$   
433 could be optimized to fit any particular dataset. However, for predictive applications the optimum  $C_s$   
434 value may not be known in advance, since it is unclear if subsequent storms at a given location will be  
435 well predicted using previously optimized  $C_s$  values. A key goal of this work is to determine if using  
436 stochastic parameterizations to generate ensembles that predict a range of dune erosion (based on  
437 uncertainty in the runup parameterization) can still capture observed dune erosion even if the optimum  
438  $C_s$  value is not known in advance. As such, a  $C_s$  value of  $1.5 \times 10^{-3}$  is used in this example application

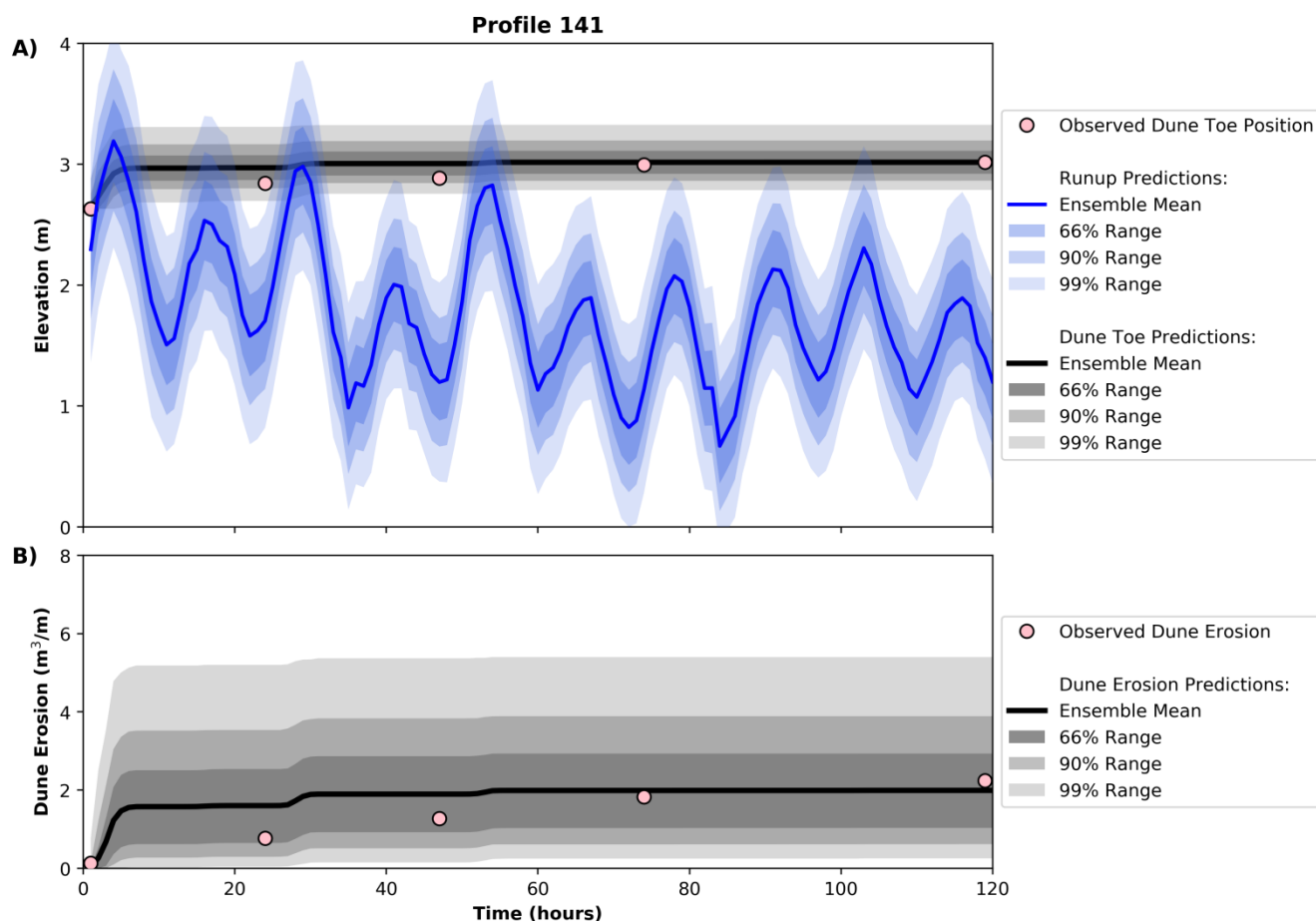


439 based on previous work at Narrabeen Beach by Ranasinghe et al. (2012). Sensitivity of model results to  
440 the choice of  $C_s$  are further discussed in **Sect. 5.2**.

441

442 An example at a single profile (profile 141, located approximately half-way up the Narrabeen  
443 embayment as shown in **Fig. 6B**) of time-varying ensemble dune erosion predictions is provided in **Fig.**  
444 **7**. It was previously shown in **Fig. 5** that only 10 random samples drawn from the GP  $R_2$  predictor were  
445 required to capture 95% of the scatter in the  $R_2$  data used to develop and test the GP. However, it is  
446 trivial to draw many more samples than this from the GP - for example, drawing 10,000 samples takes  
447 less than one second on a standard desktop computer. Therefore, to explore a large range of possible  
448 runup scenarios during the 120-hour storm event, 10,000 different runup time series are drawn from the  
449 GP and used to run LEH04 at hourly intervals, thus producing 10,000 model results of dune erosion.  
450 The effect of using different ensemble sizes is explored in **Sect. 5.2**. **Fig. 7A** shows the time-varying  
451 distribution of the runup models (blue) used to force LEH04 along with the time-varying prediction  
452 distribution of dune toe elevations (grey) throughout the 120-hour storm event. To interpret model  
453 output probabilistically, the mean of the ensemble is plotted, along with intervals capturing 66%, 90%,  
454 and 99% of the ensemble output. These intervals are consistent with those used in IPCC for climate  
455 change predictions (Mastrandrea et al., 2010) and in the context of the model results presented here,  
456 they represent varying levels of confidence in the model output. For example, there is high confidence  
457 that the real dune erosion will fall within the 66% ensemble prediction range. **Fig. 7B** shows the time-  
458 varying predicted distribution of dune erosion volumes from the 10,000 LEH04 runs. It can be seen that  
459 while the mean value of the ensemble predictions deviates slightly from the observed dune erosion, the  
460 observed erosion is still captured well within the 66% envelope of predictions.





461

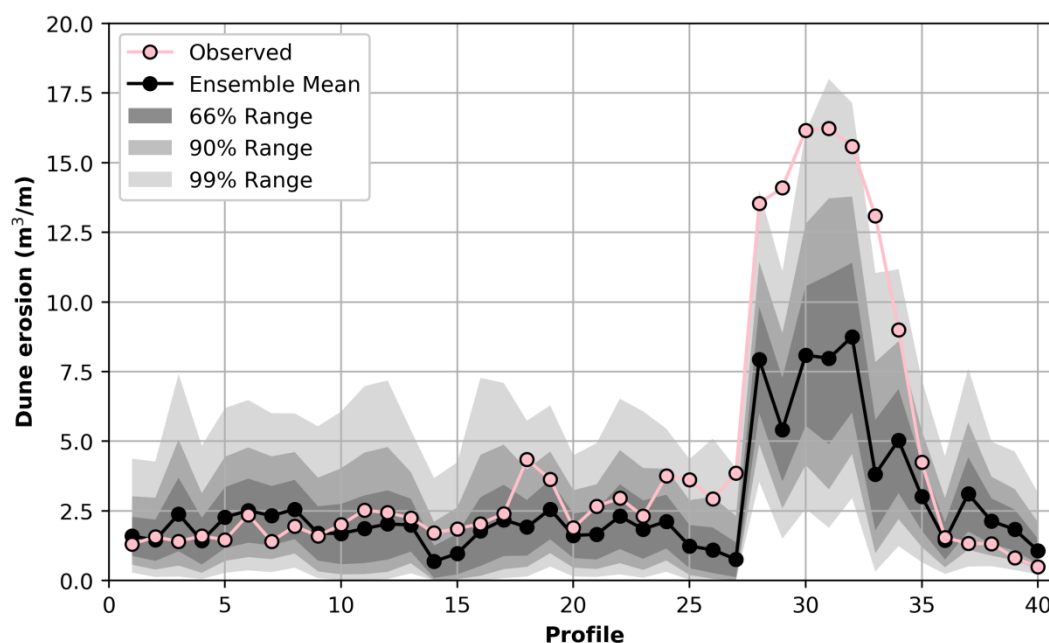
462 **Fig. 7:** Example of LEH04 used with the Gaussian process  $R_2$  predictor to form an ensemble of dune erosion predictions. 10,000  
463 runup models are drawn from the Gaussian process and used to force the LEH04 model. A) Runup (blue) and dune toe (grey)  
464 elevation for the 120-hour storm event. Bold colored line is the mean of the ensemble and shaded areas represent the regions  
465 captured by 66%, 90% and 99% of the ensemble predictions. Pink dots denote the observed dune toe elevation throughout the  
466 storm event. B) The corresponding ensemble of dune erosion predictions.

467 Pre- and post-storm dune erosion results for the 40 profiles using 10,000 ensemble members and  $C_s$  of  
468  $1.5 \times 10^{-3}$  are shown in **Fig. 8**. The squared-correlation ( $r^2$ ) for the observed and predicted dune erosion  
469 volumes is 0.85. Many of the profiles experienced only minor dune erosion ( $< 2.5 \text{ m}^3/\text{m}$ ) and can be  
470 seen to be well predicted by the mean of the ensemble predictions. In contrast, the ensemble mean can  
471 be seen to under-predict dune erosion at profiles where high erosion volumes were observed. However,  
472 the ensemble range of predictions for these profiles also has a large spread, indicative of high  
473 uncertainty in predictions. It should be noted that the results presented in **Fig. 8** are based on a non-  
474 optimized  $C_s$  value. Increasing  $C_s$  would lead to better mean ensemble predictions of the large dune



475 erosion volumes, but potentially over-prediction of the smaller events. The exact effect of varying  $C_s$  is  
476 quantified in Sect. 5.2. However, regardless of the value of  $C_s$  chosen, an advantage of the GP approach  
477 is that uncertainty in the GP predictions can give an indication of dune erosion, even if the mean dune  
478 erosion prediction deviates from the observation.

479



480

481 Fig. 8: Observed (pink dots) and predicted (black dots) dune erosion volumes for the 40 modelled profiles, using 10,000 runup  
482 models drawn from the Gaussian process and used to force the LEH04 model. Note that the 40 profiles shown are not uniformly  
483 spaced along the 3.5 km Narrabeen embayment. The black dots represent the ensemble mean prediction for each profile, while the  
484 shaded areas represent the regions captured by 66%, 90% and 99% of the ensemble predictions.



## 485 5 Discussion

### 486 5.1 Runup Predictors

487 Studies of commonly used deterministic runup parameterizations such as those proposed by Hunt  
488 (1959), Holman (1986) and Stockdon et al. (2006) amongst others, show that these parametrizations are  
489 not universally applicable and there remains no perfect predictor of wave runup on beaches (Atkinson et  
490 al., 2017; Passarella et al., 2018a; Power et al., 2018). This suggests that the available parametrizations  
491 do not fully capture all the relevant processes controlling wave runup on beaches (Power et al., 2018).  
492 Recent work has used ensemble and data-driven methods to account for unresolved factors and  
493 complexity in runup processes. For example, Atkinson et al. (2017) developed a ‘model-of-models’ by  
494 fitting a least-squares line to the predictions of several runup parameterizations. Power et al. (2018)  
495 used a data-driven, deterministic, Gene-Expression Programming model to predict wave runup against a  
496 large dataset of runup observations. Both of these approaches led to improved predictions, when  
497 compared to conventional runup parameterizations, of wave runup on the datasets tested in these  
498 studies. The work presented in this study used a data-driven Gaussian process (GP) approach to develop  
499 a probabilistic runup predictor. While the mean predictions from the GP predictor developed in this  
500 study using high-resolution LIDAR data of wave runup were accurate (RMSE = 0.18 m) and better than  
501 those provided by the Stockdon et al. (2006) formula tested on the same data (RMSE = 0.36 m), the key  
502 advantage of the GP approach over deterministic approaches is that probabilistic predictions are output  
503 that are specifically derived from data and implicitly account for unresolved processes and uncertainty  
504 in runup predictions. Previous work has similarly used GPs for efficiently and accurately quantifying  
505 uncertainty in other environmental applications (e.g., Holman et al., 2014; Kupilik et al., 2018;  
506 Reggente et al., 2014). While alternative approaches are available for generating probabilistic  
507 predictions, such as Monte Carlo simulations (e.g., Callaghan et al., 2013), the GP approach explicitly  
508 derives uncertainty from data, requires no deterministic equations, and is computationally efficient (i.e.,  
509 as discussed in Sect. 5.2, drawing 10,000 samples of 120-hour runup time series on a standard desktop  
510 computer took less than one second).



## 511 5.2 The Effect of $C_s$ and Ensemble Size on Dune Erosion

512 In **Sect. 4**, the application of the GP runup predictor within the LEH04 model to produce an ensemble  
513 of dune erosion predictions was based on 10,000 ensemble members and a  $C_s$  value of  $1.5 \times 10^{-3}$ . The  
514 sensitivity of results to the number of members in the ensemble and the value of the tunable parameter  
515  $C_s$  in **Eq. (9)** is presented in **Fig. 9**. The mean absolute error (MAE) between the mean ensemble dune  
516 erosion predictions and the observed dune erosion, averaged across all 40 profiles, varies for  $R_2$   
517 ensembles of 5, 10, 20, 100, 1000, and 10,000 members and  $C_s$  values ranging from  $10^{-5}$  to  $10^{-1}$  (**Fig. 9**).  
518 As can be seen in **Fig. 9A** and summarized in **Table 1**, the lowest MAE for the differing ensemble sizes  
519 is similar, ranging from 1.50 to 1.64  $\text{m}^3/\text{m}$ , suggesting that the number of ensemble members does not  
520 have a significant impact on the resultant mean prediction. The lowest MAE for the different ensemble  
521 sizes corresponds to  $C_s$  values between  $2.8 \times 10^{-3}$  (10,000 ensemble members) and  $4.1 \times 10^{-3}$  (5  
522 ensemble members); reasonably consistent with the value of  $1.5 \times 10^{-3}$  previously reported by  
523 Ranasinghe et al. (2012) for Narrabeen Beach and within the range of  $C_s$  values presented in Larson et  
524 al. (2004).

525

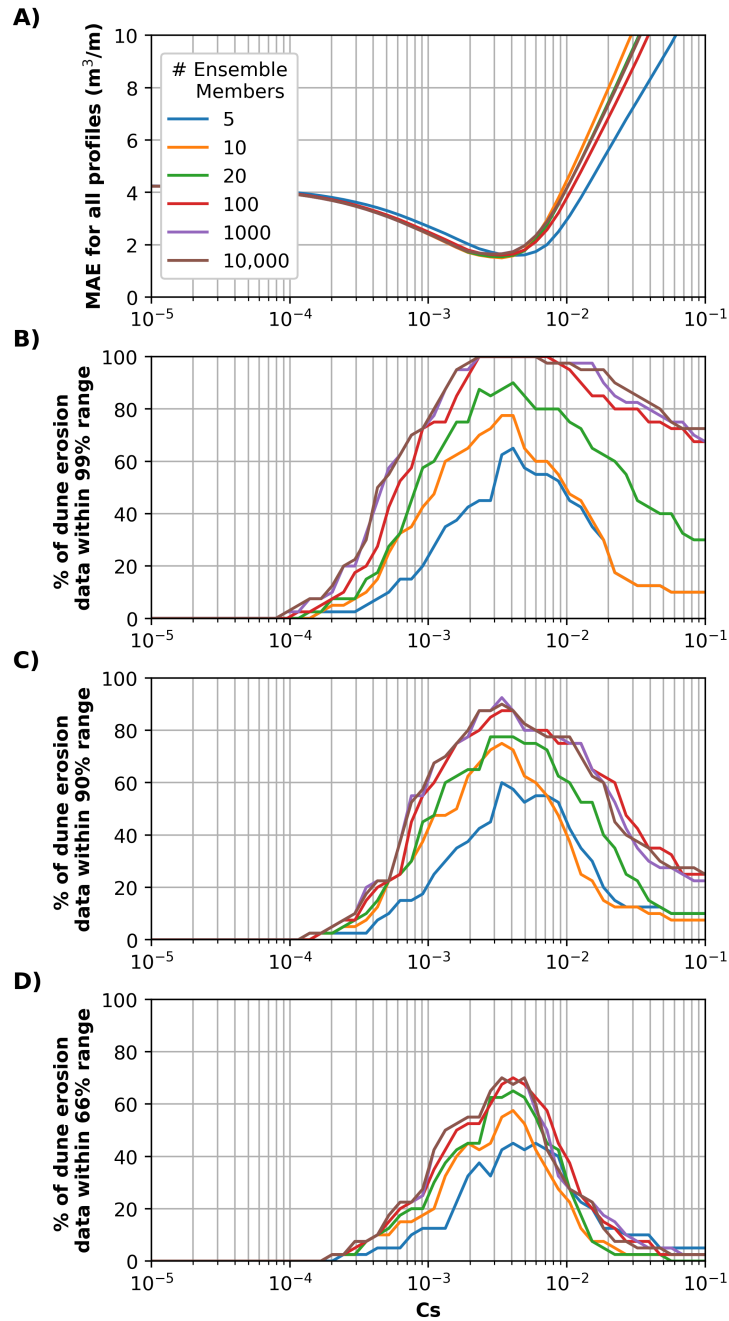
526 The key utility to using a data-driven GP predictor to produce ensembles is that a range of predictions at  
527 every location is provided as opposed to a single erosion volume. The ensemble range provides an  
528 indication of uncertainty in predictions, which can be highly useful for coastal engineers and managers  
529 taking a risk-based approach to coastal hazard management. **Fig. 9B-D** displays the percentage of dune  
530 erosion observations from the 40 profiles captured within ensemble predictions for  $C_s$  values ranging  
531 from  $10^{-5}$  to  $10^{-1}$ . It can be seen that a high proportion of dune erosion observations are captured within  
532 the 66%, 90% and 99% ensemble envelope across several orders of magnitude  $C_s$ . While the main  
533 purpose of using ensemble runup predictions within LEH04 is to incorporate uncertainty in the runup  
534 prediction, this result demonstrates that the ensemble approach is less sensitive to the choice of  $C_s$  than  
535 a deterministic model and so can be useful for forecasting with non-optimized model parameters.

536

537 Results in **Fig. 9** and **Table 1** demonstrate that there is relatively little difference in model performance  
538 when more than 10 to 100 ensemble members are used; consistent with results presented previously in



539 **Fig. 5** that showed that only 10 random samples drawn from the GP  $R_2$  predictor were required to  
540 capture 95% of the scatter in the  $R_2$  data used to develop and test the GP. This suggests that the GP  
541 approach efficiently captures scatter (uncertainty) in runup predictions and subsequently, dune erosion  
542 predictions, requiring on the order of 10 samples; significantly less than the  $10^3 - 10^6$  runs typically  
543 used in Monte Carlo simulations to develop probabilistic predictions (e.g., Callaghan et al., 2008; Li et  
544 al., 2013; Ranasinghe et al., 2012). Nevertheless, it is noted that drawing a large number of samples  
545 from the GP predictor is trivial, with 10,000 samples taking less than one second on a standard desktop  
546 computer.



547

548 **Fig. 9: Results of the stochastic parameterization methodology for  $R_2$  ensembles of 5, 10, 20, 100, 1000, and 10,000 members and  $C_s$**   
 549 **values ranging from  $10^{-5}$  to  $10^{-1}$ . A) The mean absolute error (MAE) between the median ensemble dune erosion predictions and**  
 550 **the observed dune erosion averaged across all 40 profiles. B), C) and D) show the percentage of dune erosion observations that fall**  
 551 **within the 99%, 90% and 66% ensemble prediction ranges respectively.**



552 **Table 1: Quantitative summary of Fig. 9, showing the optimum  $C_s$  value for differing ensemble sizes, along with the associated**  
 553 **mean-absolute-error (MAE) and percent of the 40 dune erosion observations captured by the 66%, 90% and 99% ensemble**  
 554 **prediction range.**

Ensemble Members	Optimum $C_s$	MAE ( $m^3/m$ )	$r^2$	Percent Captured in 66% Ensemble Range (%)	Percent Captured in 90% Ensemble Range (%)	Percent Captured in 99% Ensemble Range (%)
5	$4.1 \times 10^{-3}$	1.59	0.86	45	57	65
10	$3.4 \times 10^{-3}$	1.50	0.87	55	75	78
20	$3.4 \times 10^{-3}$	1.54	0.86	62	78	88
100	$3.3 \times 10^{-3}$	1.61	0.86	68	88	100
1000	$2.8 \times 10^{-3}$	1.64	0.86	65	88	100
10,000	$2.8 \times 10^{-3}$	1.64	0.86	65	88	100

555

### 556 5.3 Including Uncertainty in Dune Erosion Models

557 Uncertainty in wave runup predictions within dune-impact models can result in significantly varied  
 558 predictions of dune erosion. For example, the model of Larson et al. (2004) used in this study only  
 559 predicts dune erosion if runup elevation exceeds the dune toe elevation and predicts a non-linear  
 560 relationship between runup that exceeds the dune toe and resultant dune erosion. Hence, if wave runup  
 561 predictions are biased too low then no dune erosion will be predicted, and if wave runup is predicted too  
 562 high then dune erosion may be significantly over predicted. Ensemble modelling has become standard  
 563 practice in many areas of weather and climate modelling (Bauer et al., 2015), hydrological modelling  
 564 (Cloke and Pappenberger, 2009), and more recently has been applied to coastal problems such as the  
 565 prediction of cliff retreat (Limber et al., 2018) as a method of handling prediction uncertainty. While  
 566 using a single deterministic model is computationally simple and provides one solution for a given set  
 567 of input conditions, model ensembles provide a range of predictions that can better capture the variety  
 568 of mechanisms and stochasticity within a coastal system. The result is typically improved skill over  
 569 deterministic models (Atkinson et al., 2017; Limber et al., 2018) and a natural method of providing  
 570 uncertainty with predictions.

571



572 As a quantitative comparison, Splinter et al. (2018) applied a modified version of the LEH04 model to  
573 the same June 2011 storm dataset used in the work presented here with a modified expression for the  
574 collision frequency (i.e. the  $t/T$  term in **Eq. (9)**) based on work by Palmsten and Holman (2012). The  
575 parameterization of Stockdon et al. (2006) was used to estimate  $R_2$  in the model. The model was forced  
576 hourly over the course of the storm, updating the dune toe, recession slope, and profiles based on each  
577 daily LIDAR survey. Based on only the 40 profiles used in the present study, results from Splinter et al.  
578 (2018) showed that the deterministic LEH04 approach reproduced 68% ( $r^2 = 0.68$ ) of the observed  
579 variability in dune erosion. As shown in **Table 1**, the simple LEH04 model (**Eq. (9)**) applied here using  
580 the GP runup predictor to generate ensemble prediction reproduced ~85% (based on the ensemble  
581 mean) of the observed variability in dune erosion for the 40 profiles. While there are some discrepancies  
582 in the two modelling approaches, the ensemble approach clearly has an appreciable increase in skill  
583 over the deterministic approach; attributed here to using a runup predictor trained on local runup data,  
584 and the ensemble modelling approach. However, a major advantage of the ensemble approach over the  
585 deterministic approach is the provision of prediction uncertainty (e.g., **Fig. 8**). While the mean ensemble  
586 prediction is not 100% accurate, **Table 1** shows that using just 100 samples can capture all the observed  
587 variability in dune erosion within the ensemble output.

588

589 The GP approach is a novel approach to building model ensembles to capture uncertainty. Previous  
590 work modelling beach and dune erosion has successfully used Monte Carlo methods, which randomly  
591 vary model inputs within many thousands of model iterations, to produce ensembles and probabilistic  
592 erosion predictions (e.g., Callaghan et al., 2008; Li et al., 2013; Ranasinghe et al., 2012). As discussed  
593 earlier in **Sect. 5.2**, advantages of the GP approach over approaches like Monte Carlo include the  
594 explicit quantification of uncertainty directly from data, no deterministic equations are required, and the  
595 approach is computationally efficient; here, drawing 10,000 samples of 120-hour runup time series from  
596 the GP took less than one second on a standard desktop computer.





## 597 6 Conclusion

598 For coastal managers, the accurate prediction of wave runup as well as dune erosion is critical for  
599 characterizing the vulnerability of coastlines to wave-induced flooding, erosion of dune systems, and  
600 wave impacts on adjacent coastal infrastructure. While many formulations for wave runup have been  
601 proposed over the years, none have proven to accurately predict runup over a wide range of conditions  
602 and sites of interest. In this contribution, a Gaussian process (GP) was used with over 8000 high-  
603 resolution LIDAR-derived wave runup observations were used to develop a probabilistically  
604 parametrization of wave runup that quantify uncertainty in runup predictions. The mean GP prediction  
605 performed well on unseen data, with a RMSE of 0.18 m, a significant improvement over the commonly  
606 used  $R_2$  parameterization of Stockdon et al. (2006) (RMSE of 0.36 m) used on the same data. Further,  
607 only 10 randomly drawn models from the probabilistic GP distribution were needed to form an  
608 ensemble that captured 95% of the scatter in the test data.

609  
610 Coastal dune-impact models offer a method of predicting dune erosion deterministically. As an example  
611 application of how the GP runup predictor can be used in geomorphic systems, the uncertainty in the  
612 runup parameterization was propagated through a deterministic dune erosion model to generate  
613 ensemble model predictions and provide prediction uncertainty. The hybrid dune erosion model  
614 performed well on the test data, with a squared-correlation ( $r^2$ ) between the observed and predicted dune  
615 erosion volumes of 0.85. Importantly, the probabilistic output provided uncertainty bands of the  
616 expected erosion volumes which is a key advantage over deterministic approaches. Compared to  
617 traditional methods of producing probabilistic predictions such as Monte Carlo, the GP approach has the  
618 advantage of learning uncertainty directly from observed data, it requires no deterministic equations,  
619 and is computationally efficient; for the GP developed here, drawing 10,000 samples of 120-hour runup  
620 time series on a standard desktop computer took less than one second.

621  
622 This work is an example of how a machine learning model such as a GP can profitably be integrated  
623 into coastal morphodynamic models (Goldstein and Coco, 2015) to provide probabilistic predictions for  
624 nonlinear, multidimensional processes and drive ensemble forecasts. Approaches combining machine



625 learning methods with traditional coastal science and management models present a promising area for  
626 furthering coastal morphodynamic research. Future work is focused on using more and varied datasets  
627 to further train the GP developed here and to integrate it into a real-time coastal erosion forecasting  
628 system.



629 **Code and Data Availability**

630 The data and code used to develop the Gaussian Process runup predictor in this manuscript are publicly  
631 available at [https://github.com/TomasBeuzen/BeuzenEtAl\\_GP\\_Paper](https://github.com/TomasBeuzen/BeuzenEtAl_GP_Paper).



632 **Author Contributions**

633 The order of the authors' names reflects the size of their contribution to the writing of this manuscript.



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