

Response to Referee #1

Dear Referee #1,

We would like to thank you for the time you dedicated on our manuscript. We are pleased to know that you consider the revised version as greatly improved both in terms of clarity of our approach and the method used. In the revised version of the manuscript, as you kindly pointed out, the use of figures and tables within the section 3 has allowed a better explanation of our approach and a more clear definition of the credible duration.

In addition, thank you to evaluate our manuscript ready for publication following minor edits.

Your suggestions to improve the quality and the form of our study are taken into account in the last version of the manuscript and, in the following, we will give point to point detailed answer to every comments:

1) Title: 'information for regional'

The error has been corrected.

2) Abstract: insert (RFA) after first mentioning Regional Frequency Analysis (L2).

The acronym RFA has been inserted after first mentioning Regional Frequency Analysis.

3) P1 L31, insert (EVT) at this point and use EVT throughout manuscript (e.g., P4 L21)

The acronym EVT has been inserted. On the other hand, the 'Extreme Value Theory' terminology is not used anymore throughout the manuscript.

4) P2 L3: RFA – the full name is required on first use followed by (RFA).

The full name 'Regional Frequency Analysis' and its acronym have been already specified just before in **P2, L1**.

5) P2 L14: 'evaluation of the return periods of extreme events' might read better.

The error has been corrected.

6) P2 L24: 'The proposed FAB method (a declination of RFA using credible duration) allows the estimate of reliable...'

The suggestion has been accepted and performed.

7) P2 L34: 'During extreme storms,....and so this kind of event cannot always be detected.'

The error has been corrected.

8) P3 L3: 'A historical event'

This grammatical error has been corrected throughout the manuscript.

9) P3 L11: 'data points, the duration'

The error has been corrected.

10) P4 L15: Use RFA.

The acronym RFA has been used.

11) P4 L21: Use EVA, the abbreviation should be introduced on P!, see comment above.

The acronym EVA has been inserted here and it is used throughout the manuscript.

12) P4 L24: add POT after Peak-Over-Threshold as POT used in the text later.

The acronym POT has been inserted.

13) P5+: Check the journal style should variables and parameters be presented in italic in the text, e.g., d , n ,...

All variables and parameters in the text are now presented in italic.

14) P7 L11: '(2014b), for the generic RFA approach, which is...'

The suggestion has been accepted and performed as: '(2014b), for the classic RFA approach, which is...'. We would like use the term "classic" rather than "generic". In fact, FAB can be considered as a more generic RFA. RFA is implemented only for systematic data and it is a more specific method. Instead, FAB method can be used also with historical data.

15) P9 sub title: 'Application of the FAB method to a real dataset'

The error has been corrected.

16) P13 L2: 'The application of the FAB method....allows more reliablelevels to be obtained thanks to....'

The suggestion has been accepted and performed.

17) P13 L5: 'reliability of estimations'

The error has been corrected.

18) P13 L8: 'that is based on'

The error has been corrected.

19) P13 L12: 'can also be applied for'

The error has been corrected.

20) P13 L13-14: I'm not sure what is being expanded. Is it the credible duration to a regional scale or is it simply the concept of the credible duration?

'This novelty allows the expansion of the credible duration concept to a regional scale' Or

'This novelty allows the expansion of the credible duration concept at a regional scale'

The concept of credible duration exposed and used in a local scale is expanded **to a** regional scale through the notion of regional credible duration $D_{r,cr}$. The suggestion has been performed in the new version of the manuscript.

21) P13 L14: 'The RAF is another'

The error has been corrected.

22) P13 L22: 'in order to get increasingly more reliable extreme estimations and,'

The suggestion has been accepted and performed.

23) P13 L24: ' estimations by using historic data that is becoming increasingly more available through different investigations of ...'

The suggestion has been accepted and performed.

Response point-by-point to Referee #2

Dear Referee #2,

We would like to thank you again for your interesting first review that has allowed to improve a lot in our manuscript. We are glad to see that now you recommend this study for publication. The exploitation of "historical data" is a prominent way to reduce uncertainty associated with the estimation of return periods. The introduction of the new concept of "credible duration" in EVA allows satisfactorily the use of historical data in RFA.

Overall, we took into account your suggestions to still improve the quality and the clarity and the form of our manuscript.

In the following, we will give to the Referee #2 point to point detailed answer to every specific comments:

- 1) As I wrote in the previous review, the meaning of "historical data" is not immediately clear to readers coming from fields close to the one of the authors. For example, in many contexts the concept of "historical data" overlaps with the concept of hindcast or reanalysis. Though now an acceptable definition of historical data is given at line 25 of page 2, I would suggest the authors to clarify this also earlier, even in the abstract.

Thank you for this suggestion. Indeed, it is really important to clarify to readers the meaning of "historical data" from the beginning of this study. For this reason, we gave the definition of historical data immediately in the abstract as follows:

Abstract, L17-18: *"However, historical data are significant extreme events that are not recorded by tide gauge. They usually look like isolated data and they are different from continuous data from systematic measurements of tide gauges".*

- 2) The credible duration depends on the value of lambda that you use, and its link with the threshold. However, in section 3 it is not completely clear how you estimate lambda for your sample. Do you use only the systematic measurements for that? I would suggest to explain, indicating the formulation that you use here for lambda.

We agree that a direct explanation on how lambda value has been estimated should be provided. This direct explanation has been given in the updated version of our manuscript as follows:

P8, L27-29: *"The λ value is computed from the systematic measurements (Eq.2). Thanks to the hypothesis that the frequency of storms λ has no trend (Eq.3), this λ value is then used to compute the duration of the historical records. Moreover, the hypothesis of Eq.3 leads.....".*

- 3) In the abstract: ". rfa low probability extrapolation" this sentence is unclear, please reformulate it.

This unclear sentence has been reformulated as follows:

Abstract, L15-16: *“The utilisation of historical information would increase the precision and the reliability of regional extreme quantiles’ estimation”.*

4) Line 10 of conclusions: summarize here, what this hypothesis is.

We agree with the suggestion of Referee #2 and the hypothesis has been summarize in the sentence as follows:

P13, L11-12: *“The hypothesis, that the frequency of storms λ remains unchanged in time and on which credible duration concept is based, is satisfactory for skew surge”.*

The use of historical information for regional frequency analysis of extreme skew surge

Roberto Frau^{1,2}, Marc Andreewsky¹, and Pietro Bernardara³

5 ¹EDF R&D Laboratoire National d'Hydraulique et Environnement (LNHE), Chatou, 78401, France

²Université Paris-Est, Saint-Venant Hydraulics Laboratory (ENPC, EDF R&D, CEREMA), Chatou, 78401, France

³CEREA, Joint Laboratory École des Ponts ParisTech – EDF R&D, Université Paris-Est, 77455 Marne la Vallée, France

10 Correspondence to: Roberto Frau (roberto.frau@edf.fr)

Abstract. The design of effective coastal protections requires an adequate estimation of the annual occurrence probability of rare events associated to a return period up to 10^3 years. Regional Frequency Analysis (RFA) has been proved to be an applicant way to estimate extreme events using region into large and spatially distributed datasets. Nowadays, historical data are available providing new insight on past events estimations. **The utilisation of historical information would increase the precision and the reliability of regional extreme quantiles' estimation. However, historical data are significant extreme events that are not recorded by tide gauge. They usually look like isolated data and they are different from continuous data from systematic measurements of tide gauges.** This makes the definition of the duration of our observations period tricky. The duration of the observation period is however crucial for the frequency estimation of the extreme occurrence. For this reason, we introduced here the concept of “credible duration”. The proposed RFA method (hereinafter referenced as FAB, from the name of the authors) allows the use of historical data together with systematic data thanks to the use of the credible duration concept.

Key words: Regional Frequency Analysis; Historical data; Data archaeology; Skew surge; Credible duration; Regional credible duration; Historical duration; FAB method.

25 1 Introduction

Flood risk is a noteworthy threat for anthropic and industrial settlement along the coasts. Coast defence constructions have to withstand extreme events. Therefore, in this context, a statistical estimation of extreme marine variables (such as sea levels or skew surges, the difference between the maximum observed sea level and the maximum predicted astronomical tide level during a tidal cycle (Simon, 2007; Weiss, 2014c)) occurrence probability is necessary. These variables associated to high return periods (up to 1000 years (ASN, 2013)) are estimated by statistical methods linked to the Extreme Value Theory (EVT) (Fréchet, 1928; Gnedenko, 1943; Gumbel, 1958; Coles, 2001).

In the past, time series recorded at a single site have been used to estimate return levels for marine hazards. However, these extreme quantiles are linked to huge estimation uncertainties due to too short local recordings (typically 30 years to 50 years long). A greater amount of data is therefore necessary to get reliable estimations of extreme marine hazards.

For this reason and thanks to the availability of several local datasets, the Regional Frequency Analysis (RFA) is employed to reduce statistical uncertainties allowing the use of bigger extreme samples. The primary purpose of RFA is to cluster different locations in a region and to use all the data together (Cunnane, 1988; Hosking and Wallis, 1997; Van Gelder and Neykov, 1998; Bernardara et al., 2011; Weiss, 2014c). RFA is based on the index flood principle (Dalrymple, 1960) that introduces a local index for each site in order to preserve their peculiarities into a region. A regional distribution can be defined only after checking its statistical homogeneity (Hosking and Wallis, 1997). In particular, the probability distribution of extreme values must be the same everywhere in the region to allow the fitting of a probability distribution with a big data sample. Several applications exist for marine variables and, mostly, for skew surge (Bernardara et al., 2011; Bardet et al., 2011; Weiss et al., 2013).

Specifically, Weiss (2014c) proposed a RFA approach to treat extreme waves and skew surges. According to this approach, homogeneous regions are first built relying on typical storm footprints and their statistical homogeneity is subsequently tested. The addition of physical elements permits to get regions that are both physically and statistically homogeneous. Furthermore, this methodology, and in particular the pooling method, yields an equivalent regional duration in years that enables the evaluation of **the return periods of extreme events**.

As proved in hydrology during the last 30 years, extending datasets with historical data or paleofloods is an alternative way to reduce uncertainties in the estimations of extremes (Hosking and Wallis, 1986; Stedinger and Baker, 1987; Ouarda et al., 1998; Benito et al., 2004; Payrastra et al., 2011) as well as representing better potential outliers. Recent works on local sea levels (Bulteau et al., 2014) and on local skew surges (Hamdi et al., 2015) pointed out that the incorporation of historical data leads to a positive benefit in an extreme frequency analysis.

The combination of these two approaches increases the amount of available data allowing to get better extreme return levels. The use of extraordinary floods events in RFA has already been explored by Nguyen et al. (2014) for peak discharges and, only Hamdi et al. (2016) have explored it for oceanic and meteorological variables. In this study, we investigate the use of historical data in RFA of systematic measurements of skew surge levels. The proposed FAB method (**a declination of RFA using credible duration**) allows to estimate reliable extreme skew surge events taking into account all the available non-systematic records of skew surge levels. Specifically, historical data are all the skew surge events that were not recorded by tide gauge and, consequently, systematic data are all the measurements of skew surge recorded by tide gauge. One obvious and very promising way for collecting information not recorded by a tide gauge is to investigate and reconstruct historical events occurred before the starting date of gauge observations. This activity is also known as data archaeology and has been demonstrated very promising and useful for several study in the literature. Note however, that this definition also include isolated data point reconstructed during gaps of the systematic measurements. This is because statistically speaking they could be treated similarly to the isolated data reconstructed from pre-gaugement period. To find historical events it is thus necessary to go back in time and look for information before the gauging period or during each data gap in a time series data. **During extreme storms**, tide gauge may be affected by a partial or total failure in the sea level measurements **and so this kind of event cannot always be detected**. Thanks to non-systematic

records, we can hold a numerical value to these ungauged extreme events. This allows to use a bigger number of storms in the statistical assessment of the extreme skew surges.

Anyway, Baart et al. (2011) pointed out the hitch in recovering accurate historical information. A historical event is the result of a wide investigation of several qualitative and quantitative sources. Historical data are collected from different sources as ancient newspapers, archives, water marks, palaeographic studies and others. For this reason, each historical data has its own origin and its quality has to be checked. Therefore, a critical analysis of historical data must be performed before applying a statistical approach.

Using historical data with time series data recorded by a tide gauge is not trivial. The extreme statistical analysis requires a knowledge of the time period for which all the events above an extreme threshold are known (Leese, 1973; Payraastre et al., 2011). For the gauged or, better, systematic data, this time is represented by the recording period (i.e. systematic duration). As historical data are isolated data points, the duration of the observation is not defined. What happened during the time period between two or more past extreme events remains unknown. However, the coverage period estimation of historical data is a key step in the extreme value statistical models (Prosdocimi, 2017). Reasonable computation of historical duration is the base of reliable statistical estimation of rare events.

In the past, to overcome this issues, the perception threshold concept has been introduced. The perception threshold represents the minimum value over which all the historical data are reported and documented (Payraastre et al., 2011; Payraastre et al., 2013). The hypothesis is thus set that all the events over the perception threshold must have been recorded and the time period without records corresponds to time period without extreme events. This allows the estimation of the coverage period and the indirect assumption that the collected historical data are exhaustive throughout the coverage period assessed. So far, the perception threshold has allowed to use historical information in an extreme frequency analysis (Payraastre et al., 2011; Payraastre et al., 2013; Bulteau et al., 2014) but the hypothesis for using this approach seems sometimes difficult to prove.

Indeed, the perception threshold is usually defined equal to a historical event value, based on the principles that, if any other historical event would have occurred it would have been recorded (Payraastre et al., 2011; Bulteau et al., 2014). This is a quite strong hypothesis, which is difficult to be verified and validated. Moreover, the duration associated to the perception threshold is usually defined as the duration between the oldest historical data and the start of the systematic series (Payraastre et al., 2011; Hamdi et al., 2015).

The main point of novelty of the paper is to propose a new approach in order to deal with these questions in a more satisfactory way, the equivalent duration of historical data.

Note that the new approach is required for the application of the RFA, at least in the POT approaches (Van Gelder and Neykov, 1998; Bernardara et al., 2011; Bardet et al., 2011; Weiss et al., 2014b). Indeed, the definition of the perception threshold equal to a historical value precludes the possibility to use RFA methodology proposed by Weiss (2014c). Specifically, in Weiss regional methodology, the number of extreme events per year λ must be

equal in each site of a region in order to estimate a regional distribution. This is impossible to achieve if the threshold is linked to the characteristic of the local observations.

In addition, when historical information (i.e. non-systematic data) is collected during a total failure of tide gauge, this influences the frequency of the event. One could only re-estimate λ if the credible duration is re-estimated (for details, see the section 2).

In order to deal with these questions, in this paper, a different approach is proposed to estimate the equivalent duration of historical data. This evaluation is based on the hypothesis that the frequency of the extreme events is not changing significantly in time. In other words, the average number of extreme events should be the same (obviously this has to be carefully verified and demonstrated case by case). In this way, we can assess the coverage period of historical data. Based on the frequency value of systematic storms occurrence, our hypothesis introduces a physical and tangible element in the estimation of the historical duration. For this reason, we state that this novelty allows to overtake the concept of perception threshold providing a credible estimation of the unknown historical duration used to estimate the extreme quantiles.

Thanks to this new concept of duration, in this study, the FAB methodology is developed in order to use together historical data and systematic data in a RFA of extreme events. This methodology, described in details in section 2, is applied to the extreme skew surges given by 71 tide gauges located in the UK's, French's and Spanish's coasts and 14 historical collected skew surges is shown in section 4. In addition, a functional example on the computation of credible duration is illustrated in section 3.

2 FAB Methodology

2.1 The concept of credible duration

In the Extreme Value Analysis (EVA) of natural hazards, the estimations of extreme quantiles are performed applying statistical tools to a dataset of extreme variables. Usually, these datasets of extremes can be generated, from continuous time series data, or taking as extremes the maximum value in several predefined time blocks (Block Maxima approach) or fixing a sampling threshold (Peak-Over-Threshold (POT) approach).

Based on a POT approach, our study examines datasets of skew surges. From now on, we use skew surge as extreme variable but all concepts that we will describe can be extended also to other natural hazards.

In the POT framework, an extreme sampling threshold has to be considered in order to extract the dataset of extreme skew surges from a continuous time series data of systematic skew surges recorded by tide gauge. In this way, the extreme sample is created and the extreme quantiles computation of skew surge can be performed in according to the statistical distribution used. The relation between return periods T and quantiles x_T is expressed as:

$$P(X > x_T) = \frac{1}{\lambda T} \quad (1)$$

where $P(X > x_T)$ is the exceedance probability of x_T or, thoroughly, the probability that in every λT years an event X , strongest than x_T , occur. Thanks to Eq.1, the results can be illustrated through a return level plot.

The variable λ represents the number of skew surge events per year that exceed the sampling threshold. In other words, λ is the frequency of storms occurrence and it can be expressed as the ratio between the number of skew surges n over the sampling threshold and the recording time of systematic skew surge measurements in years d (hereinafter called systematic duration):

$$5 \quad \lambda = \frac{n}{d} \quad (2)$$

The knowledge of n and d is necessary to compute the frequency of extreme skew surges occurrence λ and to estimate extreme quantiles of skew surges. In other words, a duration estimation is needed to transform the probability [no units] in frequency [time⁻¹ dimension].

When historical skew surge are available, the computation of λ is a tricky topic. In fact, although the number of historical data n_{hist} over the sampling threshold is known, we do not have any information on which happens during non-gauged period. It is thus impossible to define a duration over which the n_{hist} data are observed. In other words, fixing whatever sampling threshold, we need this information for defining the frequency of occurrence of the events over the threshold (see the Eq.2). We called the missing information “historical duration” d_{hist} or, in other words, the time period for which historical skew surge above sampling threshold are estimated to be exhaustively recorded at the best of our knowledge. The inclusion of historical data in a systematic dataset of extreme skew surge is shown as an example in Fig.1. It needs the formulation of a credible hypothesis in order to estimate the duration of the whole dataset composed by systematic and historical records.

Taking advantage to the perfect knowledge of systematic skew surges, the basic hypothesis proposed here for the computation of historical duration is that the frequency of storms λ remains unchanged in time. More precisely, λ_{sys} , namely the number of systematic skew surges per year over the sampling threshold, is supposed to be equal to λ_{hist} , namely the number of historical skew surges per year over the same sampling threshold.

$$20 \quad \lambda = \lambda_{\text{sys}} = \lambda_{\text{hist}} \quad (3)$$

Eq.3 is in accordance with the lack of a significant trend on the extreme skew surges frequency during the second part of 19th century, 20th century and the first part of 21st century in the regions considered in this study (Fig.5).

According to current scientific literature, based on past observations, frequency of storms has no defined trend (Hanna et al., 2008; Barring and Fortuniak, 2009; MICORE Project, 2009). Therefore, the hypothesis that storm frequency is constant has been assumed. No significant trends are detected in terms of magnitude or frequency of storms between 1780 and 2005 in the North Atlantic and European regions since the Dalton minimum (Barring and Fortuniak, 2009). Over the past 100 years in the North Atlantic basin, no robust trends in annual numbers of tropical storms, hurricanes and major hurricanes counts have been identified by the last IPCC report of 2013. Hanna et al. (2008), Matulla et al. (2008) and Allan et al. (2009) state that observations have no clear trends over the past century or longer with substantial decadal and longer fluctuations but with the exceptions of some regional and seasonal trends (Wang et al., 2009; Wang et al., 2011). Moreover, MICORE (2009) in the “Review of climate

change impacts on storm occurrence” affirms that results from coastal areas do not detect significant trend in frequency of storms for the existing and available datasets.

Therefore, the hypothesis that storm frequency is constant may be a reasonable working hypothesis for demonstrating the approach. Obviously, it is crucial to test this hypothesis before the application of the approach in the present state to any new dataset. Note that the methodology could be adapted quickly in the future for non-stationary datasets, where the introduced concept can be adjusted. For skew surges, this hypothesis is ensured.

The assumption that the annual frequency of skew surges is constant in Europe during the last two centuries is widely warranted by the several references mentioned above. A solid check must be done for others marine variables to know if they can fit this hypothesis as well.

Anyway, a new duration of observation d_{cr} is defined as “credible” (i.e. based on credible hypothesis) considering a steady λ . This novelty is expressed as:

$$d_{cr} = \frac{n_{tot}}{\lambda} = \frac{n_{syst} + n_{hist}}{\lambda} = \frac{n_{syst}}{\lambda} + \frac{n_{hist}}{\lambda} \quad (4)$$

Eq.4 shows how the number of extreme data is divided in two sides according to the nature of data: n_{syst} is the number of extreme systematic data above the sampling threshold u and n_{hist} is the number of extreme historical data over the same threshold u . Therefore, the credible duration depends also on the sampling threshold value.

Moreover, this splitting provides indirectly the coverage period of historical skew surges in addition to the duration of recorded data. The historical side of the duration d_{cr}^{hist} can be formulated as follows:

$$d_{cr}^{hist} = \frac{n_{hist}}{\lambda} \quad (5)$$

The coverage period on average that every historical skew surge hold is called associated duration $d_{cr,ass}$

$$d_{cr,ass} = \frac{d_{cr}^{hist}}{n_{hist}} = \frac{1}{\lambda} \quad (6)$$

Eq.6 displays how the associated duration can be evaluated just knowing λ . Therefore, every extreme data not originated from a recorded time series has an associated duration. This is also the case for historical events occurred during a time gap in gauged skew surges series.

Anyway, a functional example is illustrated in Section 3 order to grasp all the notions explained up to here and, in particular, their relevance in the sampling threshold choice.

2.2 The use of credible duration in RFA framework

Weiss (2014c) defined a methodology to fit a regional probability distribution for sea extreme variables. This approach enables to put together data from different locations and to form regions. The probability distribution of extreme skew surges must be the same inside a region. Thus, the extreme data of a same region are gathered together by the use of a normalisation parameter that preserves the peculiarities of each site and the regional probability distribution is so defined. The formation of regions based on typical storm footprints (Weiss et al., 2013) is used to create physically and statistically homogeneous regions and it leads to know a master component of this regional process: the regional effective duration D_{eff} .

The regional effective duration D_{eff} represents the effective duration (in years), filtered of any intersite dependence, in which the number of regional data N_r are observed. D_{eff} is defined as the product between the degree of regional dependence φ [1; N] and the mean of local durations

$$5 \quad D_{\text{eff}} = \varphi \times \bar{d} = \frac{\lambda_r}{\lambda} \bar{d} = \frac{N_r}{\lambda} \quad (7)$$

in Eq.7, λ_r is the number of regional storms per year $\lambda_r = N_r/\bar{d}$ and \bar{d} is the mean of local observations' duration of the N regional sites $\bar{d} = \sum_{i=1}^N d_i/N$.

10 The local frequency of storms λ must be the same in all the region. More details about the D_{eff} are clarified in Weiss (2014b), **for the classic RFA approach, which** is set for systematic data.

The knowledge of the regional sample duration, as exposed for a generic frequency analysis in the previously paragraph of this section, is a needed factor. In fact, the evaluation of the regional return period T_r of a particular regional quantile x_{T_r} is expressed as in Eq.8:

$$T_r = \frac{1}{\lambda_r P(X_s > x_{T_r})} \quad (8)$$

15 where $P(X_s > x_{T_r})$ is the probability that a storm s impacts at least one site in the region with a normalized intensity greater than x_{T_r} . Therefore, thanks to the regional distribution F_r , the local return period of the storm s is calculated as:

$$T = \frac{1}{\lambda(1-F_r(x_{T_r}))} \quad (9)$$

20 In addition, Weiss (2014b) use for the computation of the empirical local return periods a Weibull plotting position. This element is correlated with the regional duration D_r :

$$\tilde{T}_{s,loc} = \frac{D_r+1}{n_r+1-rank(s)} \quad (10)$$

Then, the biggest storm s_{max} of the regional sample is linked to a local return period thanks to D_r .

25 However, the need to employ historical data in the extreme analysis has been considered as highly relevant by scientific community to get reliable quantiles and to not miss significant information. In the RFA, the use of historical records is allowed by the new notion of credible duration. A new way to process historical skew surges at regional scale is introduced. The addition of one or more historical skew surges in one or more sites of a region requires the knowledge of the duration of each site. Thanks to the credible duration concept applied previously to a single site, a coverage period for both types of data can be evaluated. It allows to switch from the notion of the regional effective duration $D_{r,\text{eff}}$ to the regional credible duration $D_{r,\text{cr}}$:

$$D_{r,cr} = \varphi \times \bar{d} = \varphi \sum_{i=1}^N \frac{d_{i,cr}}{N} \quad (11)$$

Regional credible duration depends also on the degree of regional dependence φ that increases when the data are independent. Moreover, Eq.11 shows that a local duration surplus brings consequently an excess of the coverage period of regional sample. With some simplifications, the Eq.11 can thus be so reformulated:

$$D_{r,cr} = \frac{\lambda r}{\lambda} \bar{d} = \frac{Nr}{\lambda} \quad (12)$$

The lack of a trend on storms' frequency is a necessary condition to assume a steady λ . If this hypothesis is satisfied, the regional credible duration can be computed.

The regional frequency analysis of skew surges including historical data or, FAB method, can be performed and these two combined approaches ensure better estimations of extreme values.

3 Illustration of credible duration

In order to better illustrate the concept of credible duration and how it varies according to different thresholds, a specific illustration on connection between these variables is described in this section by the example shown in Figure 1.

Fig. 1 shows the functional example of the whole skew surge series recorded at La Rochelle tide gauge (France). Systematic skew surges are represented by the continuous series of data plotted in black from about 1940 to 2017. They are observed during a fixed period equal to the recording time period, in this case, 32.88 years.

The historical skew surges are artificial data generated by ourselves for this descriptive illustration and they are represented as green cross points. As they are not supposed to be obtained by continuous systematic records of skew surge measurements, the duration of their support is unknown.

As expressed above, d_{cr} and in particular d_{cr}^{hist} depends on the sampling threshold value u . Eq.4 and Eq.5 show this dependence through the λ value.

The application of the new proposed concept allows the computation of the historical duration of skew surges that represents a basic way to take into account the historical data on the estimation of extreme values.

Although the period of systematic observation is known and fixed by tide gauge's records, the coverage period of historical observations is associated to the value of sampling threshold u . **The λ value is computed from the systematic measurements (Eq.2). Thanks to the hypothesis that the frequency of storms λ has no trend (Eq.3), this λ value is then used to compute the duration of the historical records.** Moreover, the hypothesis of Eq.3 leads to different values of historical duration for each λ selected value and consequently, for each likely sampling threshold. Four cases of different values of high sampling threshold are examined (Fig.2 and Tab.1). These particular cases are chosen in order to let us know the effects of different sampling thresholds on the computation of the credible duration.

Case 1 uses a sampling threshold u_{C1} corresponding to 0.5 extremes per year ($\lambda=0.5$) and lead to a credible duration of 42.88 years (Tab.1). A higher threshold u_{C2} is used in Fig.2 for the Case 2 ($\lambda=0.23$) and how you can see the credible duration grows until 54.62 years because the number of historical data over the threshold does not change. In Case 3, the threshold (u_{C3}) increases ($\lambda=0.2$) but the credible duration decreases (52.88 years) because the smallest value of historical skew surge does not exceed the threshold. On average, the coverage period of each historical event still rises. Case 4 displays the use of a very high threshold u_{C4} ($\lambda=0.07$) to define the sample of extremes. In this case, only one historical event is bigger than threshold and its coverage period or, associated duration, corresponds to the historical duration. More the historical data value is big and more high is the probability that the historical event covers a longer period.

As stated above, the four cases (resume in Fig.2 and Tab.1) show the behaviour of credible duration for different values of threshold. As threshold value increases, associated duration increases as well. Eq.4 expresses the credible duration as the sum of systematic duration and historical duration. Systematic duration is steady and it is independent from threshold value. For this reason, the variation of credible duration depends only on the variation of historical duration. In particular, when the threshold is extremely high, it is almost sure that no extreme skew surges occurred during the coverage period of historical data. If the threshold is lower, then it is less certain that all extreme skew surges have been observed in the historical part.

However, it is not always true that the credible duration is biggest when threshold grows (Fig.3). In fact, the credible duration and historical duration depends not only of the threshold value but also of the number of historical events exceeding it. If the threshold is extremely high, not all historical data are taken into account in the extremes' sample. On the contrary, the associated duration of each historical data increases when threshold grows for the reasons explained in the previous paragraph (shown by the blue line in Fig.3).

In conclusion, the sampling threshold choice is an important step in the extreme quantiles estimation allowing the computation of the credible duration. Finally, with all of these elements, we can deal with historical data in a regional framework.

4 Application of the FAB method to a real dataset

4.1 Systematic skew surge database and historical data

The systematic skew surge database is originated from time series of sea levels recorded by 71 tide gauges located along French, British and Spanish coasts of the Atlantic Ocean, the English Channel, the North Sea and the Irish Sea (Fig.4).

French data are provided by SHOM (Service Hydrographique et Océanographique de la Marine, France), English data by BODC (British Oceanographic Data Center, UK) and Spanish data by IEO (Instituto Espanol de Oceanografía, Spain). All data, except Spanish data, are updated until January 2017. Every site has different systematic durations depending on the tide gauge records. Starting from 1846, the longest time series is recorded

at Brest for a systematic period of 156.57 years. Sea levels are hourly recorded except for British ports. BODC supplies sea levels every 15 minutes from 1993.

Skew surge data are generated in each site from the difference between the maximum sea levels recorded around the time of the predicted astronomical high tide and the self-same predicted astronomical high tide. Before computing skew surge, sea levels must be corrected by a likely long-term alteration of mean sea levels, or eustatism. Only sea level data with significant linear trends are edited. More details on the computation of skew surges are described by Simon (2007), Bernardara et al. (2011) and Weiss (2014c). In this way, the systematic database of 71 skew surge series is created.

Finding accurate historical data is hard and tricky and, at the moment, only 14 historical skew surge are available.

These historical records are located in 3 of the 71 ports: 9 at La Rochelle (Gouriou, 2012; Garnier and Surville, 2010; Breilh, 2014), 4 at Dunkerque (DREAL, 2009; Parent et al., 2007; Maspataud, 2011) and 1 at Dieppe (provided by the Service hydrographique et océanographique de la Marine SHOM, France). The historical data are introduced in corresponding skew surge series of the concerned sites.

This skew surge database of systematic and historical data is exploited in this study for the estimation of regional return periods.

4.2 Formation of regions

The method to form homogeneous regions proposed by Weiss et al. (2014a) is used to get physically and statistically homogeneous regions for the database shown above. This methodology is based on a storm propagation criterion identifying the most typical storms footprints. The configuration of parameters to detect storms are set using the same criteria of Weiss (2014c). The most typical storm footprints are revealed by Ward's hierarchical classification (Ward, 1963) and correspond to four physical regions. The homogeneity test (Hosking and Wallis, 1997) is applied to check if all regions are also statistically homogeneous. Only one physical region is not statistically homogeneous and, after an inner partition and a further check, five physically and statistically homogeneous regions are obtained.

Moreover, a double threshold approach (Bernardara et al., 2014) is employed to separate the physical part from the statistical part of the used extreme variables and two different thresholds are used: the first to define physical storms (physical threshold) and the second one to define regional samples (statistical sampling threshold or, more simply, sampling threshold).

Fig.5 displays the five regions founded out through the application of Weiss' methodology. Although a different input database is used, the number and the position of regions are similar to Weiss' study.

The regionalisation processing shows that only Region 1 and Region 2 hold historical data. The three sites with historical data belong actually to these two regions: La Rochelle to Region 1 and, Dieppe and Dunkerque to Region 2.

The region with more historical information is the Region 1 (nine historical skew surges) and the results of FAB method proposed are focused on this region.

4.3 Focus on Region 1

4.3.1 Statistical threshold and credible duration

The sample of regional extremes is analysed with statistical tools in order to estimate regional return levels associated to very extreme return periods. A Regional Pooling Method is used to select independent maximum data. A regional sample is formed through normalised local data. Maximum local data are divided by a local index, equal to the local sampling threshold value (Weiss, 2014c). Moreover, this methodology enables us to estimate the regional distribution F_r considered as the distribution of the normalised maximum data above the statistical sampling threshold.

This threshold corresponds to a storm's frequency λ that is assumed equal at all sites of a region. The sampling threshold value selection or, better, λ value selection is a tricky topic (Bernardara et al., 2014). There is not an univocal approach to define the exact sampling threshold. Twelve sensitivity indicators depending on all possible λ values are proposed to assign the best λ value for the regional analysis of extreme skew surge using historical data: the statistical homogeneity test (Hosking and Wallis, 1997), the stationarity test applied to storm intensities of regional sample (Hosking and Wallis, 1993), the chi-squared test for the regional distribution (Cochran, 1952), the statistical test to detect outliers in regional sample (Hubert and Van der Veecken, 2008; Weiss, 2014c), the value of regional credible duration, the number of regional sample's data, the value of the degree of regional dependence φ , the visual look of the regional return plot and the stability of regional credible duration, local return levels, shape parameter and scale parameter of the GPD distribution (Coles, 2001). The λ value selection can not be carried out through one of these methods and only the better combination of all indicators can enable us to appoint the optimal λ . All sensitivity indicators have the same relevance in the proposed sensitivity analysis.

A skew surge's frequency of 0.36 is chosen for Region 1 thanks the sensitivity analysis of the statistical factors mentioned above. Taking $\lambda=0.36$ means to have on average one extreme skew surge over the sampling threshold every 2.78 years at each site. This threshold selection permits the computation of the credible duration for every site belonging to the Region 1. In the Region 1, all ports, except for La Rochelle, have only systematic skew surge and their credible duration, equal to systematic duration, is determined and independent from the sampling threshold value.

On the contrary, the site of La Rochelle has nine historical skew surges and its credible duration depends both on systematic duration and historical duration. The systematic duration does not change and it is equal to the recording period of La Rochelle tide gauge. As shown in the fictive illustration of section 3, the historical duration differs while threshold value changes.

The credible duration at La Rochelle is 57.88 years and the contribution of historical data is almost the 43% of the total duration. Obviously, without these nine historical data, the coverage period of extreme data at La Rochelle would have been less than 25 years, equivalent to the historical duration. All the historical data are over the chosen threshold and, consequently, the associated duration of each historical data is equal to 2.77 years.

4.3.2 Regional credible duration

Eq.11 enables us to compute the regional credible duration for Region 1. The introduction of nine historical data in the local time series and, specifically, only eight in the regional sample after the use of Regional Pooling method, grants an earning of 11.11 years of regional duration. As Tab.2 shows, historical data increase the credible durations for sites where there are historical records and, consequently, the regional credible duration increases likewise. In Region 1, only the site of La Rochelle increases its duration. This contribution allows to increase the regional credible duration.

Regional credible duration depends not only on the mean of local durations but also on the degree of regional dependence φ (Eq.11). Obviously, this degree does not change with only nine more added data. In conclusion, the rise of the regional credible duration of Region 1 is produced mostly by the increase of credible duration at La Rochelle.

Tab.2 shows the difference with the FAB method that allows to take into account historical data and a traditional method of RFA (Bernardara et al., 2011; Weiss, 2014c) that does not enable the use of historical data. Fig.6 displays how more λ decreases in value, the more the difference between regional credible duration with the nine historical data (blue line) and the regional effective duration without historical data (green line) for Region 1 increases (red line). This is due to the greater certainty that historical data above the sampling threshold are the biggest during a longer time.

4.3.3 Regional return levels

Regional return levels for Region 1 are obtained through the fit of a Generalized Pareto Distribution (Fig.7). Penalized maximum likelihood estimation (PMLE) (Coles and Dixon, 1999) is used to estimate the regional GPD parameters (γ, k) . As shows Fig.7, the regional distribution is unbounded due to a positive shape parameter k . The return level plot of Region 1 is produced with the 95% confidence intervals generated by a bootstrap of observed storms. A Weibull plotting position is employed for the regional empirical return period. Return levels for each site are achieved through multiplication between regional return levels and local indices, equivalent to the local sampling threshold.

Red dotted points represent the eight historical data taken into account in Region 1. As appears, the benefit of using historical data is not only an increase of the coverage period of extreme data but also able to consider in a frequency analysis non-systematic storms occurred. The historical skew surges used in this study represent very big storms and, for this reason, it is newsworthy to take into account their large magnitude. Indeed, the biggest normalised skew surge considered in Region 1 is 3.01 compared with 2.13 of the regional sample generated without historical data (Tab.2).

5 Conclusions

The application of the FAB method, based on the new concept of credible duration, allows more reliable extreme estimations of return levels to be obtained thanks to the inclusion of historical data to a systematic skew surge database.

5 The introduction of historical data on a frequency analysis of extremes is an innovative way to increase the reliability of estimations. In the past, the estimation of the period covered by the historical data collection has usually limited the use of historical data. The concept of perception threshold was employed so far to use historical data even though its hypothesis are quite strong and difficult to verify. For this reason, the new concept of credible duration has been formulated that is based on physical considerations supported by several scientific studies. This
10 leads to use historical isolated data points on the EVA.

The hypothesis, that the frequency of storms λ remains unchanged in time and on which credible duration concept is based, is satisfactory for skew surge. For other variables, a hypothesis check has to be done. Following the hypothesis check, all the considerations that we raised in this study can also be applied for other natural hazards.

15 When historical data are available, the credible duration depends on sampling threshold. This novelty allows the expansion of the credible duration concept to a regional scale. The RFA is another efficient way to reduce uncertainties on the extreme estimations. These two ways can be used together by the proposed FAB method.

20 The results describe the relevance of using historical records on a regional analysis both in terms of regional duration and of storm magnitude. Sometimes, these data can represent very large events and so we have to consider them during the estimation of extreme variables. Moreover, the existence of historical records enables to raise the coverage period of the extreme sample.

No extreme data must be missed in an extreme frequency analysis. It is very important to use all of the available data in order to get increasingly more reliable extreme estimations and, regional analysis and historical data are two tools that can be merge with satisfactory results.

25 New perspectives are opened to improve extreme quantiles estimations by using historical data that is becoming increasingly more available through different investigations on several qualitative and quantitative sources. The aggregation of all of these data in a historical database is a need to exploit the maximum amount of extremes.

Furthermore, every historical data value is associated to different uncertainties and it is challenging to be able to take into account these uncertainties in the regional analysis.

References

- Allan, R., Tett, S. and Alexander, L.: Fluctuations in autumn-winter severe storms over the British Isles: 1920 to present, *International Journal of Climatology*, 29(3), 357-371, doi:10.1002/joc.1765, 2009.
- ASN: Guide de l'ASN n°13 : Protection des installations nucléaires de base contre les inondations externes, 2013.
- 5 Baart, F., Bakker, M. A. J., van Dongeren, A., den Heijer, C., van Heteren, S., Smit, M. W. J., van Koningsveld, M., and Pool, A.: Using the 18th century storm-surge data from the Dutch Coast to improve the confidence in flood-risk estimates, *Natural Hazards and Earth System Sciences*, 11, 2791-2801, doi:10.5194/nhess-11-2791-2011, 2011.
- 10 Bardet, L., Duluc, C.-M., Rebour, V., and L'Her, J.: Regional frequency analysis of extreme storm surges along the French coast, *Natural Hazards and Earth System Sciences*, 11(6), 1627-1639, doi:10.5194/nhess-11-1627-2011, 2011.
- Barnett, V., and Lewis, T.: *Outliers in Statistical Data*, 3rd Edition. John Wiley, Chichester, 1994.
- Barriendos, M., Coeur, D., Lang, M., Llasat, M.C., Naulet, R., Lemaitre, D., and Barrera, A.: Stationarity analysis of historical flood series in France and Spain (14th–20th centuries), *Natural Hazards and Earth System Sciences*, 15 3(6), 583–592, doi:10.5194/nhess-3-583-2003, 2003.
- Barring, L. and Fortuniak, K.: Multi-indices analysis of southern Scandinavian storminess 1780-2005 and links to interdecadal variations in the NW Europe-North sea region, *International Journal of Climatology*, 29(3), 373-384. doi:10.1002/joc.1842, 2009.
- 20 Benito, G., Lang, M., Barriendos, M., Llasat, M., Frances, F., Ouarda, T., Thorndycraft, V., Enzel, Y., Bardossy, A., Coeur, D., and Bobée, B.: Use of Systematic, Paleoflood and Historical Data for the Improvement of Flood Risk Estimation, *Review of Scientific Methods, Natural Hazards*, 31(3), 623-643, 2004.
- Benson, M. A.: Use of historical data in flood-frequency analysis. *Eos, Transactions American Geophysical Union*, 31(3), 419-424, doi:10.1029/TR031i003p00419, 1950.
- 25 Bernardara, P., Andreewsky, M., and Benoit, M.: Application of regional frequency analysis to the estimation of extreme storm surges, *Journal of Geophysical Research*, 116, C02008, doi:10.1029/2010JC006229, 2011.
- Bernardara, P., Mazas, F., Kergadallan, X., and Hamm, L.: A two-step framework for over-threshold modelling of environmental extremes, *Natural Hazards and Earth System Sciences*, 14, 635-647, doi:10.5194/nhess-14-635-2014, 2014.
- 30 Breilh, J.F., Bertin, X., Chaumillon, E., Giloy, N. and Sauzeau, T.: How frequent is storm-induced flooding in the central part of the Bay of Biscay?, *Global and Planetary Change*, 122, 161-175. doi:10.1016/j.gloplacha.2014.08.013, 2014a.
- Breilh, J.F.: *Les surcotes et les submersions marines dans la partie centrale du Golfe de Gascogne : les enseignements de la tempête Xynthia*, Ph.D. thesis, Sciences de la Terre, Université de La Rochelle, France, available at: <https://tel.archives-ouvertes.fr/tel-01174997/document>, 2014b.

- Bulteau, T., Idier, D., Lambert, J., and Garcin, M.: How historical information can improve extreme coastal water levels probability prediction: application to the Xynthia event at La Rochelle (France), *Natural Hazards and Earth System Sciences*, 15, 1135-1147, doi:10.5194/nhess-15-1135-2015, 2015.
- Cochran, W. G.: The χ^2 Test of Goodness of Fit, *The Annals of Mathematical Statistics*, 23(3), 315-345, 1952.
- 5 Cohn, T. A.: The incorporation of historical information in flood frequency analysis, M. S. thesis, Dep. of Environ. Eng., Cornell Univ., Ithaca, N.Y., U.S., 1984.
- Cohn, T. A., and Stedinger, J. R.: Use of historical information in a maximum-likelihood framework, *Journal of Hydrology*, 96(1-4), 215-223, doi:10.1016/0022-1694(87)90154-5, 1987.
- 10 Coles, S. and Dixon, M.: Likelihood-based inference for extreme value models, *Extremes*, 2(1), 5-23, doi:10.1023/A:1009905222644, 1999.
- Coles, S.: *An Introduction to Statistical Modeling of Extreme Values*, Springer, London, doi:10.1007/978-1-4471-3675-0, 2001.
- Condie, R., and Lee, K. A.: Flood frequency analysis with historic information, *Journal of Hydrology*, 58(1-2), 47-61, doi:10.1016/0022-1694(82)90068-3, 1982.
- 15 Cunnane, C.: Methods and merits of regional floods frequency analysis, *Journal of Hydrology*, 100(1/3), 269-290, doi:10.1016/0022-1694(88)90188-6, 1988.
- Dalrymple, T.: *Flood Frequency Analyses*, *Manual of Hydrology: Part 3, "Flood Flow Techniques"* U.S. Geological Survey Water-Supply Paper, 1543-A, US GPO, 1960.
- DREAL (Ministère de l'Ecologie, de l'Energie, du Développement durable, et de l'Aménagement du territoire),
 20 59/62 SREI/DRNHM: Détermination de l'aléa de submersion marine intégrant les conséquences du changement climatique en région Nord – Pas-de-Calais – Etape 1: compréhension du fonctionnement du littoral, available at: <http://www.hauts-de-france.developpement-durable.gouv.fr/IMG/pdf/Determination-alea-submersion-marine-integrant-consequences-changement-climatique-dreal-npdc.pdf>, 2009.
- Fréchet, M.: Sur la lois de probabilité de l'écart maximum, *Annales de la Société Polonaise de Mathématique*, 6,
 25 93–122, 1928.
- Garnier, E. and Surville, F.: *La tempête Xynthia face à l'histoire, Submersions et tsunamis sur les littoraux français du Moyen Âge à nos jours*, Le croît Vif., Saintes, France, 2010.
- Gnedenko, B.: Sur la distribution limite du terme maximum d'une série aléatoire, *Annals of Mathematics*, 44(3), 423-453, 1943.
- 30 Gouriou, T.: Evolution des composantes du niveau marin à partir d'observations de marégraphe effectuées depuis la fin du 18e siècle en Charente-Maritime, Ph.D. thesis, Océanographie physique, Université de La Rochelle, France, 2012.
- Gumbel, E. J.: *Statistics of Extremes*, Columbia University Press, NewYork, U.S., 1958.

- Hamdi, Y., Bardet, L., Duluc, C.-M., and Rebour, V.: Use of historical information in extreme surge frequency estimation: case of the marine flooding on the La Rochelle site in France, *Natural Hazards and Earth System Sciences*, 15, 1515-1531, doi:10.5194/nhess-15-1515-2015, 2015.
- 5 Hamdi, Y., Bardet, L., Duluc, C.-M., and Rebour, V.: Development of a target-site-based regional frequency model using historical information, *European Geosciences Union General Assembly 2016 (EGU 2016)*, Vienna, Austria, 17-22 April 2016, 18, EGU2016-8765, 2016.
- Hanna, E., Cappelen, J., Allan, R., Jonsson, T., Le Blancq, F., Lillington, T. and Hickey, K.: New Insights into North European and North Atlantic Surface Pressure Variability, Storminess, and Related Climatic Change since 1830, *Journal of Climate*, 21, 6739-6766, doi:10.1175/2008JCLI2296.1, 2008.
- 10 Hosking, J. R. M. and Wallis, J. R.: Paleoflood hydrology and flood frequency analysis, *Water Resources Research*, 22(4), 543-550, doi:10.1029/WR022i004p00543, 1986a.
- Hosking, J. R. M. and Wallis, J. R.: The value of historical data in flood frequency analysis, *Water Resources Research*, 22(11), 1606-1612, doi:10.1029/WR022i011p01606, 1986b.
- Hosking, J. R. M. and Wallis, J. R.: Some statistics useful in regional frequency analysis, *Water Resources*
15 *Research*, 29(2), 271-281. doi:10.1029/92WR01980, 1993.
- Hosking, J. R. M. and Wallis, J. R.: *Regional Frequency Analysis. An approach based on L-moments*, Cambridge University Press, 1997.
- Hubert, M., and Van der Veeken, S.: Outlier detection for skewed data, *Journal of chemometrics* 22(3-4), 235-246, doi:10.1002/cem.1123, 2008.
- 20 IPCC: *Climate change 2013: the physical science basis, Contribution to the IPCC fifth assessment report (AR5), Chapter 2: Observations: Atmosphere and Surface*, 2013.
- Jenkins, G. J., Murphy, J. M., Sexton, D. M. H., Lowe, J. A., Jones, P. and Kilsby, C. G.: *UK Climate Projections: Briefing report*, Met Office Hadley Centre, Exeter, UK, 2009.
- Leese, M. N.: Use of censored data in the estimation of Gumbel distribution parameters for annual maximum flood
25 series, *Water Resources Research*, 9(6), 1534-1542, doi:10.1029/WR009i006p01534, 1973.
- Maspataud, A.: *Impacts des tempêtes sur la morphodynamique du profil côtier en milieu macrotidal*, Ph.D. thesis, Océanographie, Université du Littoral Côte d'Opale, France, available at: <https://tel.archives-ouvertes.fr/tel-00658671/document>, 2011.
- Matulla, C., Schoner, W., Alexandersson, H., von Storch, H. and Wang, X. L.: European storminess: late
30 nineteenth century to present, *Climate Dynamics*, 31(2-3), 125-130, doi:10.1007/s00382-007-0333-y, 2008.
- MICORE Project 202798: *Review of climate change impacts on storm occurrence*, edited by O.Ferreira (UALG), M.Vousdoukas (UALG) and P.Ciavola (UniFe), available at: <http://www.micore.eu/file.php?id=4>, 2009.
- Nguyen, C.C., Gaume, E., Payrastre, O.: Regional flood frequency analyses involving extraordinary flood events at ungauged sites: further developments and validations. *Journal of Hydrology*, 508, 385-396,
35 doi:10.1016/j.jhydrol.2013.09.058, 2014.

- Ouarda, T. B. M. J., Rasmussen, P. F., Bobée, B., and Bernier, J.: Utilisation de l'information historique en analyse hydrologique fréquentielle, *Revue des sciences de l'eau*, 11, 41-49, doi:10.7202/705328ar, 1998.
- Parent, P., Butin, T., Vanhee, S. and Busz, N.: Un territoire soumis au risque de submersion marine. Les tempêtes de 1949 et 1953 à Dunkerque, Institution départementale des Wateringues, available at :
5 <http://www.ingeo.fr/fichier/file/Wateringues/Affiche4-Dunkerque.pdf>, 2007.
- Payraastre, O., Gaume, E., and Andrieu, H.: Usefulness of historical information for flood frequency analyses: developments based on a case study, *Water Resources Research*, 47, W08511, doi:10.1029/2010WR009812, 2011.
- Payraastre, O., Gaume, E., and Andrieu, H.: Historical information and flood frequency analyses: which optimal features for historical floods inventories?, *La Houille Blanche - Revue internationale de l'eau*, 3, 5-11,
10 doi:10.1051/lhb/2013019, 2013.
- Picklands, J.: Statistical inference using extreme order statistics, *The Annals of Statistics*, 3(1), 119-131, 1975.
- Pouvreau, N.: Trois cents ans de mesures marégraphiques en France : outils, méthodes et tendances des composantes du niveau de la mer au port de Brest, Ph.D. thesis, Climatologie, Université de La Rochelle, France, available at: <https://tel.archives-ouvertes.fr/tel-00353660/document>, 2008.
- 15 Prosdocimi, I.: German tanks and historical records: the estimation of the time coverage of ungauged extreme events, *Stochastic Environmental Research and Risk Assessment*, 1-16, doi:10.1007/s00477-017-1418-8, 2017.
- Reis, D. S. and Stedinger, J. R.: Bayesian MCMC flood frequency analysis with historical information, *Journal of Hydrology*, 313(1-2), 97-116, doi:10.1016/j.jhydrol.2005.02.028, 2005.
- Sabourin, A., Renard, B.: Combining regional estimation and historical floods: A multivariate semiparametric peaks-over-threshold model with censored data, *Water Resources Research*, 51(12), doi:10.1002/2015WR017320,
20 2015.
- SHOM (Service Hydrographique et Océanographique de la Marine): Références Altimétriques Maritimes - Edition 2016, 2016.
- Simon, B.: La marée océanique côtière, Editions de l'Institut Océanographique, 2007.
- 25 Stedinger, J. R. and Cohn, T. A.: Flood frequency analysis with historical and paleoflood information, *Water Resources Research*, 22(5), 785-793, doi:10.1029/WR022i005p00785, 1986.
- Stedinger, J. R. and Baker, V. R.: Surface water hydrology: historical and paleoflood information, *Reviews of Geophysics*, 25(2), 119-124, doi:10.1029/RG025i002p00119, 1987.
- Van Gelder, P. H. A. J. M.: A new statistical model for extreme water levels along the Dutch coast, edited by:
30 Tickle, K. S., Goulter, I. C., Xu, C. C., Wasimi, S. A., Bouchart, F., *Stochastic Hydraulics '96*, Proceedings of the 7th IAHR International Symposium, Mackay, Queensland, Australia, 243-249, 1996.
- Van Gelder, P. H. A. J. M., and Neykov, N. M.: Regional frequency analysis of extreme water level along the Dutch coast using L moments: A preliminary study, *Stochastic models of hydrological processes and their applications to problems of environmental preservation*, NATO Advanced Research Workshop, 23-27 November
35 1998, Moscow, Russia, 14-20, 1998.

Wang, X. L., Zwiers F. W., Swail V. R., and Feng Y.: Trends and variability of storminess in the Northeast Atlantic region, 1874-2007. *Climate Dynamics*, 33(7-8), 1179-1195, doi:10.1007/s00382-008-0504-5, 2009.

Wang, X. L., Wan, H., Zwiers, F. W., Swail, V. R., Compo, G. P., Allan, R. J., Vose, R. S., Jourdain, S., and Yin, X.: Trends and low-frequency variability of storminess over western Europe, 1878–2007, *Climate Dynamics*, 5 37(11-12), 2355–2371, doi:10.1007/s00382-011-1107-0, 2011.

Ward, J.H.: Hierarchical grouping to optimize an objective function, *Journal of the American Statistical Association*, 58(301), 236-244, 1963.

Weiss, J., Bernardara, P., and Benoit, M.: A method to identify and form homogeneous regions for regional frequency analysis of extreme skew storm surges, *Proceedings of the 1st International Short Conference on Advances in Extreme Value Analysis and Application to Natural Hazards*, 18-20 September 2013, Siegen, 10 Germany, 2013.

Weiss, J., Bernardara, P., and Benoit, M.: Formation of homogeneous regions for regional frequency analysis of extreme significant wave heights, *Journal of Geophysical Research: Oceans*, 119(5), 2906-2922, doi:10.1002/2013JC009668, 2014a.

15 Weiss, J., Bernardara, P., and Benoit, M.: Modeling intersite dependence for regional frequency analysis of extreme marine events, *Water Resources Research*, 50, 5926-5940, doi:10.1002/2014WR015391, 2014b.

Weiss, J.: Analyse régionale des aléas maritimes extrêmes, Ph.D. thesis, Mécanique des fluides, Université Paris-Est, France, available at: <https://pastel.archives-ouvertes.fr/tel-01127291/document>, 2014c.

20

25

30

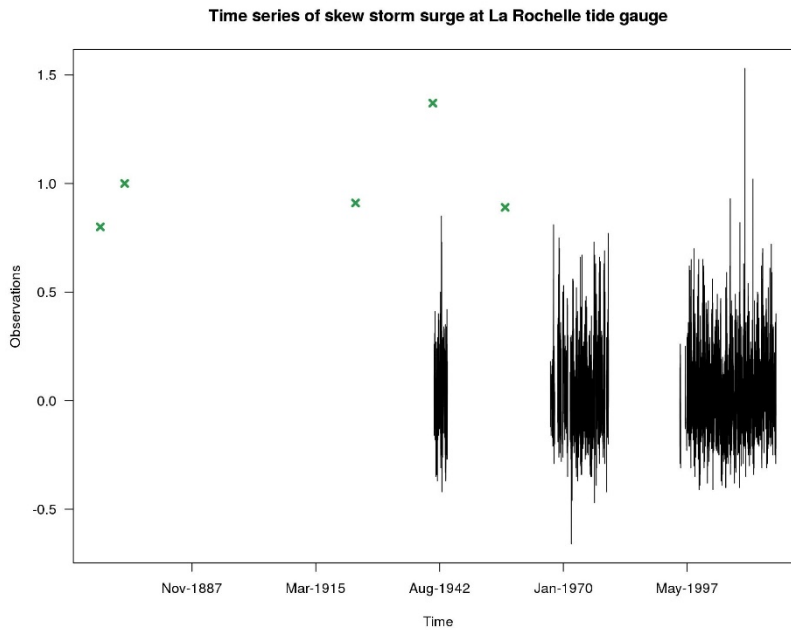
Table 1. Summary table of several variables involved in the computation of credible duration for 4 different cases of study.

	Case 1	Case 2	Case 3	Case 4
Lambda λ	0.5	0.23	0.2	0.07
Sampling threshold (m)	0.7	0.77	0.81	1.02
Systematic duration (years)	32.88	32.88	32.88	32.88
Historical credible duration (years)	10	21.74	20	14.29
Number of historical data OT	5	5	4	1
Associated duration (years)	2	4.35	5	14.29
Credible duration (years)	42.88	54.62	52.88	47.17

5

Table 2. Summary table of the results obtained with the application of FAB method for Region 1.

	RFA without historical data	RFA with 9 historical data (FAB, 2017)
Lambda λ	0.36	0.36
Local duration on average (years)	41.4	42.36
Regional credible duration (years)	558.33	569.44
Intensity of biggest normalized surge	2.13 (SD)	3.01 (HD)



10

Fig. 1. Illustrative example of skew surge dataset at La Rochelle tide gauge. In black the systematic skew surge and in green the historical isolated data point (non-systematic data).

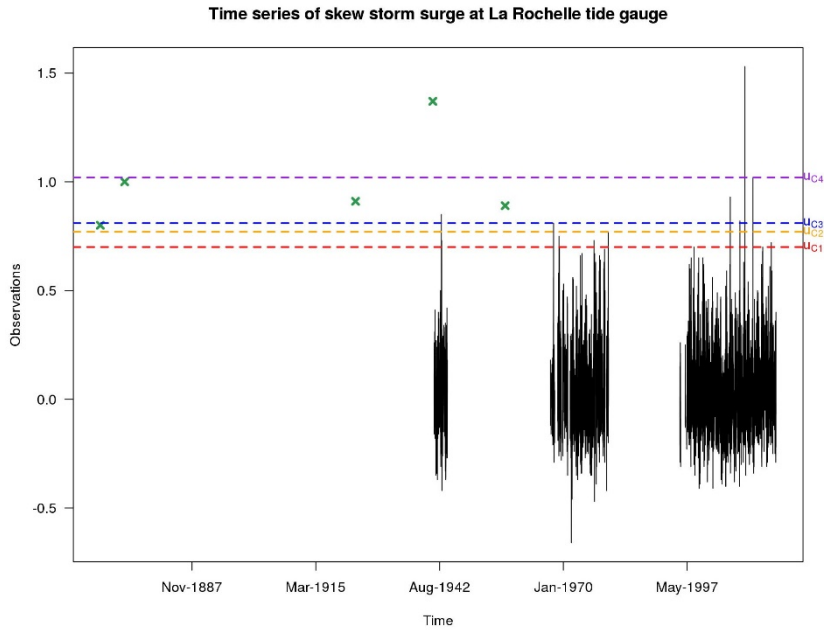
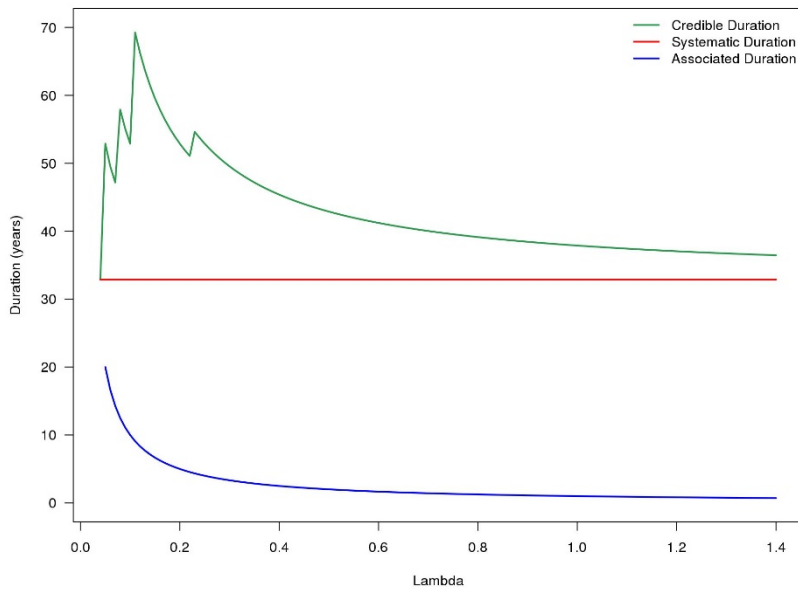


Fig. 2 Skew surge dataset of Section 3 (composed by systematic and historical data) used in the computation of credible duration with 4 different thresholds u_{c1} , u_{c2} , u_{c3} , u_{c4} . The four sampling thresholds u_c are used linked to a rising value of λ ($\lambda_{c1}=0.5$, $\lambda_{c2}=0.23$, $\lambda_{c3}=0.2$, $\lambda_{c4}=0.07$).



5

Fig. 3. Fictive example of skew surge dataset of systematic and historical data - Credible duration (green line), systematic duration (red line) and associated duration (blue line) for several threshold values.

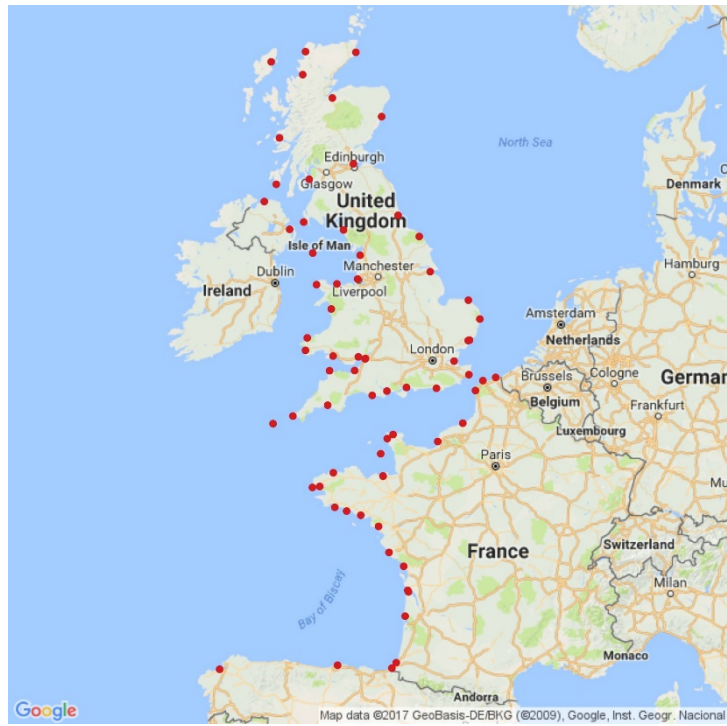


Fig. 4. Location of the 71 tide gauges used in this application.

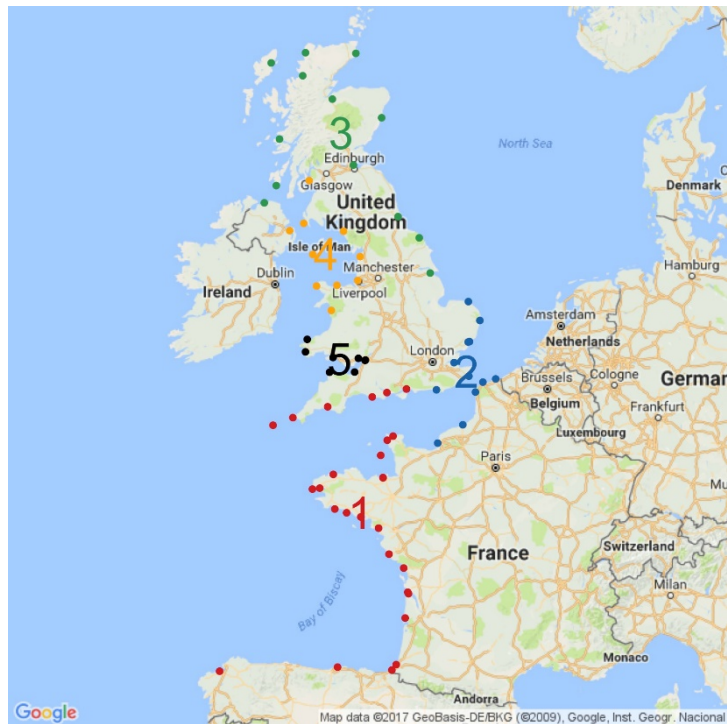


Fig. 5. The five physically and statistically homogeneous regions for the 71 tide gauges.

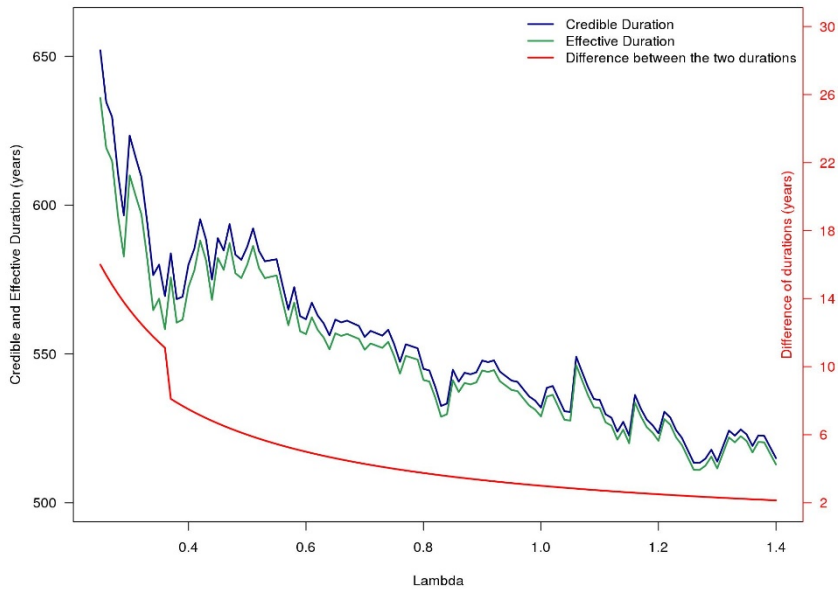
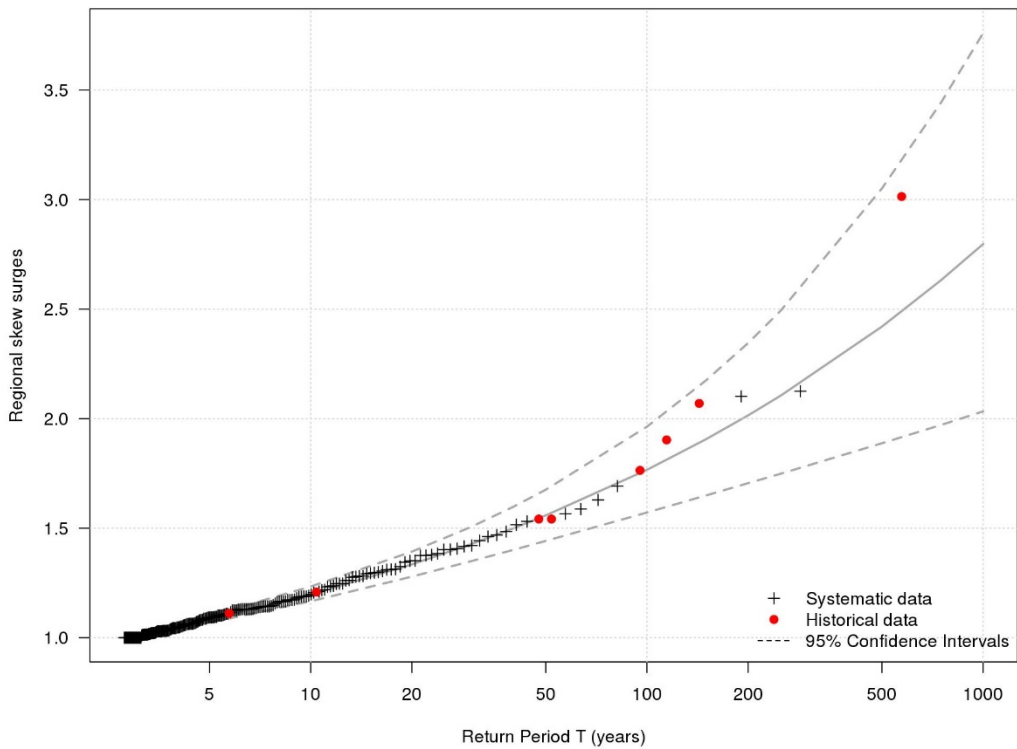


Fig. 6. Variation of regional credible duration (taking into account the nine historical data) and regional effective durations (only with systematic data) and their difference (in red) for Region 1.

Regional Return Level plot - Region 1



5 Fig. 7. Return level plot for the regional distribution of Region 1 with the confidence intervals of 95% estimated by bootstrap. Red points represent the regional historical data and black crosses the regional systematic data.